

## Sparse Bayesian Inference with Regularized Gaussian Distributions

@ AIP 2023, MS12 1: Fast optimization-based methods for inverse problems

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Joint work with Martin S. Andersen and Yiqiu Dong



Computational **U**ncertainty **Q**uantification for **I**nverse problems

DTU Compute

Department of Applied Mathematics and Computer Science

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## Sparsity in linear least squares estimation

Parameters:  $x \in \mathbb{R}^n$ , linear forward operator:  $A \in \mathbb{R}^{m \times n}$ , and data:  $b = Ax + e \in \mathbb{R}^m$ .

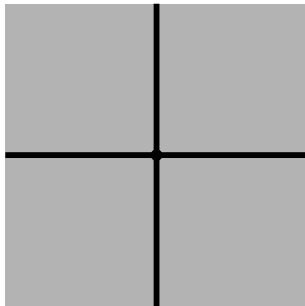
Regularized linear least squares

$$\min_{z \in \mathbb{R}^n} \left\{ \frac{\lambda}{2} \|Az - b\|_2^2 + f(z) \right\}$$

can give sparse solutions,  
i.e., many zeros in  $x$  and/or  $Lx$ .

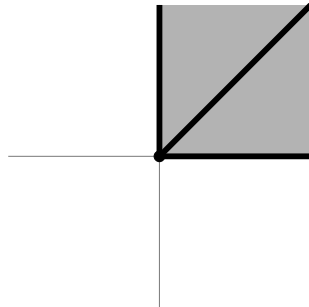
$$f(x) = \|x\|_1$$

LASSO



$$f(x) = \|Lx\|_1 + \chi_{\mathbb{R}_{\geq 0}^n}(x)$$

Constrained generalized  
LASSO



## Sparsity with continuous distributions

Given a continuous posterior distribution,  
e.g.,

$$\pi(x | b) \propto \exp\left(-\frac{\lambda}{2}\|Ax - b\|_2^2 - f(x)\right),$$

then, for "approximate" sparsity,

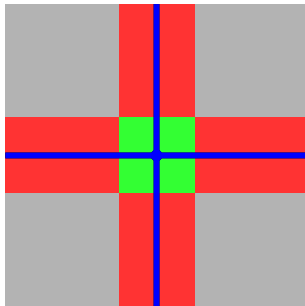
$$\mathbb{P}(\blacksquare) > 0, \quad \mathbb{P}(\blacksquare) > 0,$$

but for "true" sparsity,

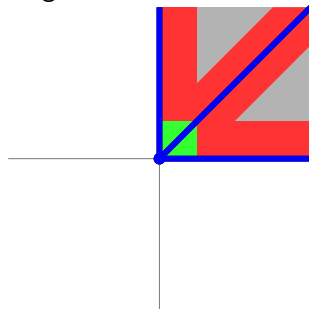
$$\mathbb{P}(\blacksquare) = 0.$$

**Does the difference matter?**

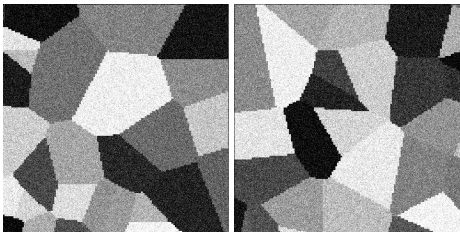
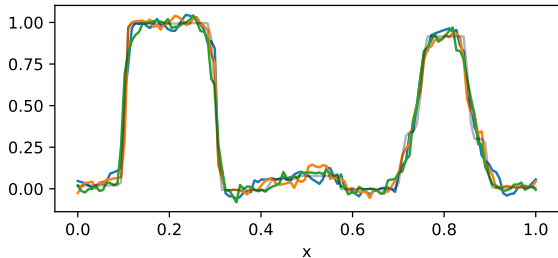
$f(x) = \|x\|_1$   
Bayesian LASSO



$f(x) = \|Lx\|_1 + \chi_{\mathbb{R}_{\geq 0}^n}(x)$   
Bayesian constrained  
generalized LASSO

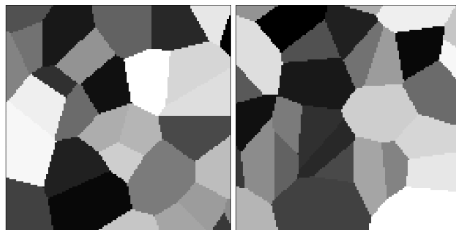
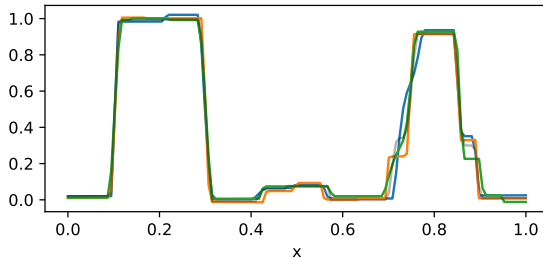


### Samples we get from a continuous distribution



exaggerated

### Sparse samples we (might) want



## Sparsity with distributions of varying dimensions

### Varying dimension model:

Partition the domain in different models

$$F_i,$$

define model priors

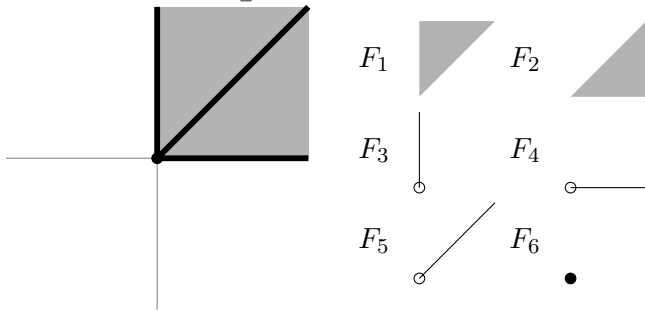
$$\mathbb{P}(F_i) \quad \text{and}$$

define conditional densities

$$\pi(x | F_i),$$

of dimension  $\dim(F_i)$ .

$$f(x) = \|Lx\|_1 + \chi_{\mathbb{R}_{\geq 0}^n}(x)$$



## Sampling with sparsity

### Sampling from varying dimension models

Reversible-Jump Markov Chain Monte Carlo (RJMCMC), e.g., STMALA, which can be difficult to work with in high-dimensional problems.

### Our method: regularized Gaussian distribution

$$x | b := \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 + f(z) \right\}, \quad \text{with } \hat{b} \sim \mathcal{N}(b, \Sigma).$$

### Trade-off:

Disadvantage: Implicit assumptions on the prior behind  $x | b$ .

Advantage: Use tools from optimization theory to analyze and sample.

## Regularized Gaussian distribution

**(Linear) Randomize-Then-Optimize (RTO)/Perturbation Optimization (PO):**

If  $\pi(x | b) \propto \exp\left(-\frac{1}{2}\|Ax - b\|_{\Sigma^{-1}}^2\right)$ , i.e., Gaussian posterior, then

$$x | b = \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 \right\}, \quad \text{with } \hat{b} \sim \mathcal{N}(b, \Sigma).$$

**Randomization/Perturbation:** sample from  $\hat{b}$ ,

**Optimization:** solve the optimization problem with the  $\hat{b}$  sample.

**Regularized Gaussian Distribution:**

$$x | b := \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 + f(z) \right\}, \quad \text{with } \hat{b} \sim \mathcal{N}(b, \Sigma).$$

**Is  $x | b$  is a well-defined probability distribution? How does  $x | b$  look like?**

## Rank issue

$$x | b := \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 + f(z) \right\}, \quad \text{with } \hat{b} \sim \mathcal{N}(b, \Sigma).$$

If  $\text{rank}(A) < n$ , i.e., there is not enough (artificial) data, then **the probability distribution is not always well-defined\***.

For the remainder of this talk, assume  $\text{rank}(A) = n$ .



## Regularization is post-processing

### Formulation 1: Regularized Gaussian distribution

$$x | b := \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 + f(z) \right\}, \quad \text{with } \hat{b} \sim \mathcal{N}(b, \Sigma).$$

### Formulation 2: Proximal post-processed posterior

$$x^* := \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 \right\}, \quad \text{i.e., } \pi(x^*) \propto \exp \left( -\frac{1}{2} \|Ax^* - b\|_{\Sigma^{-1}}^2 \right),$$

then

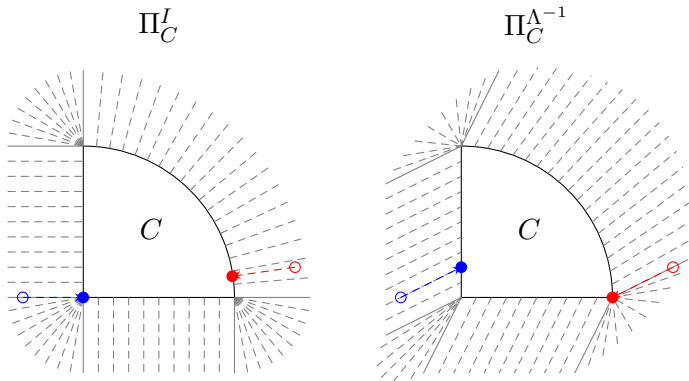
$$x | b = \text{prox}_f^{\Lambda^{-1}}(x^*) := \operatorname{argmin}_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|z - x^*\|_{\Lambda^{-1}}^2 + f(z) \right\}, \quad \text{with } \Lambda^{-1} = \operatorname{Cov}(x^*)^{-1}.$$

**Why not use a cheaper post-processor?**

## Example - Constrained

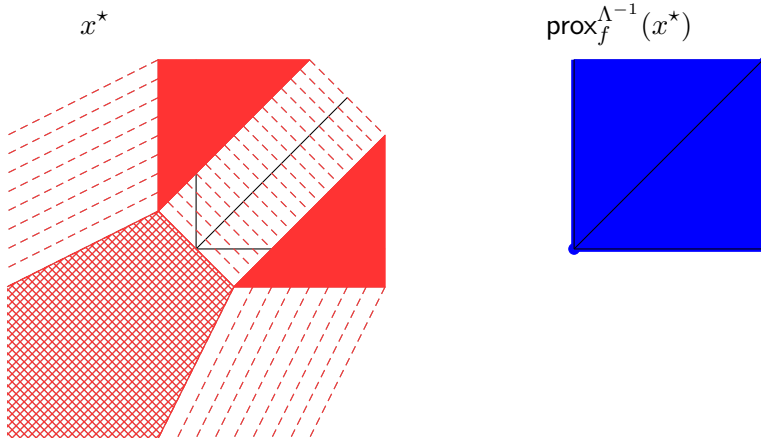
If  $f(x) = \chi_C(x)$ , then  $\text{prox}_f^{\Lambda^{-1}}(x^*) = \Pi_C^{\Lambda^{-1}}(x^*)$  is an oblique projection.

$$\Pi_C^I(x^*) := \underset{x \in C}{\operatorname{argmin}} \|x - x^*\|_2^2 \quad \text{or} \quad \Pi_C^{\Lambda^{-1}}(x^*) := \underset{x \in C}{\operatorname{argmin}} \|x - x^*\|_{\Lambda^{-1}}^2$$



## Example - Nonnegative Total Variation

$$f(x) = |x_2 - x_1| + \chi_{\mathbb{R}_{\geq 0}^2}(x)$$



$$\mathbb{P}(x^* \in \blacksquare) > 0 \quad \implies \quad \mathbb{P}(\text{prox}_f^{\Lambda^{-1}}(x^*) \in \blacksquare) > 0$$

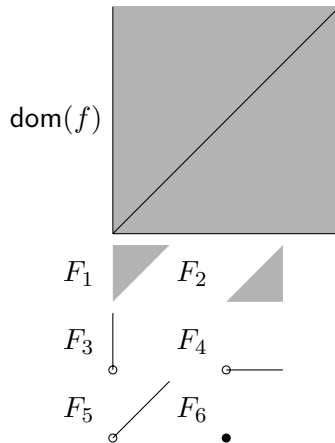
## Low-dimensional subspaces

$\text{prox}_f^{\Lambda^{-1}}(x^*)$  has positive probability on low-dimension subspaces representing sparsity.

Similar results hold if  $f$  is:

- convex piecewise linear/polyhedral
- curved constraints, e.g., ball
- group lasso, e.g.,  $\sum_{i=1}^k \|D_i x\|_2$

$$f(x) = |x_2 - x_1| + \chi_{\mathbb{R}_{\geq 0}^2}(x)$$



## A Bayesian look

For  $f$  convex piecewise linear:

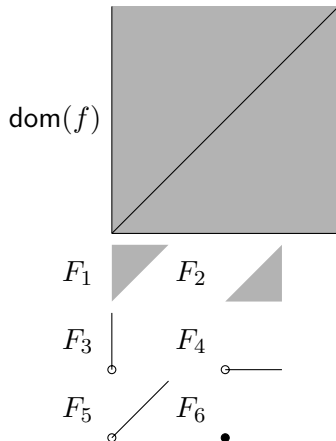
Probability distribution satisfies

$$\pi(x | b, F_j) \propto \exp\left(-\frac{1}{2}\|Ax - b\|_{\Sigma^{-1}}^2 - f(x)\right).$$

Conditional prior:

$$\pi(x | F_i) \propto \frac{\pi(x | b, F_i)}{\pi(b | x, F_i)} \propto \exp(-f(x)).$$

$$f(x) = |x_2 - x_1| + \chi_{\mathbb{R}_{\geq 0}^2}(x) = \|Lx\|_1 + \chi_{\mathbb{R}_{\geq 0}^n}(x)$$



## Hierarchical model

Assume  $f$  is positive homogeneous and convex piecewise linear, e.g,  $f(x) = \|Lx\|_1 + \chi_{\mathbb{R}_{\geq 0}^n}$ .

$$x | b, \lambda, \gamma := \operatorname{argmin}_{x \in \mathbb{R}^n} \left\{ \frac{\lambda}{2} \|Ax - \hat{b}\|_2^2 + \gamma f(x) \right\}, \quad \text{with } \hat{b} \sim \mathcal{N}(b, \lambda^{-1}I) \quad (1)$$

Add hyperpriors:  $\lambda \sim \Gamma(\alpha_\lambda, \beta_\lambda)$  and  $\gamma \sim \Gamma(\alpha_\gamma, \beta_\gamma)$ .

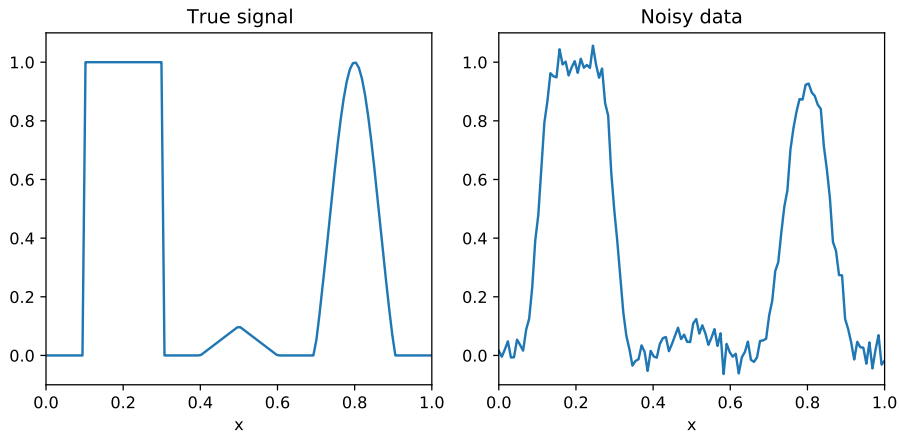
## Hierarchical Gibbs Sampler

Repeat:

- ❶  $\lambda$ :  $\lambda_k \sim \Gamma(m/2 + \alpha_\lambda, \frac{1}{2} \|Ax^{k-1} - b\|_2^2 + \beta_\lambda)$ ,
- ❷  $\gamma$ :  $\gamma_k \sim \Gamma(\dim(F(x^{k-1})) + \alpha_\gamma, f(x^{k-1}) + \beta_\gamma)$ ,  $\dim(F(x^{k-1}))$  is the level of sparsity
- ❸  $x$ : Sample  $x|b, \lambda^k, \gamma^k$  through (1).

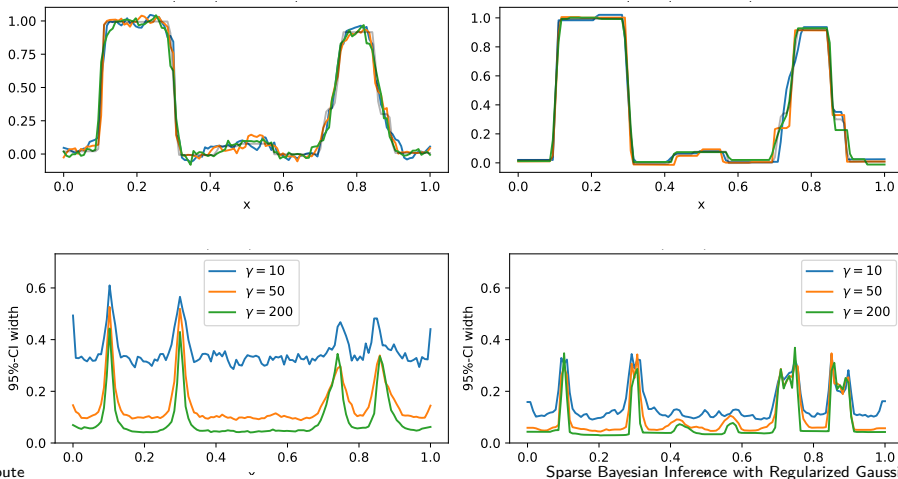
Sampling from (1) can only be done efficiently **approximately**, e.g., using ADMM.

## Numerical example: 1D deblurring with TV



## Numerical example: deblurring with TV (no Gibbs)

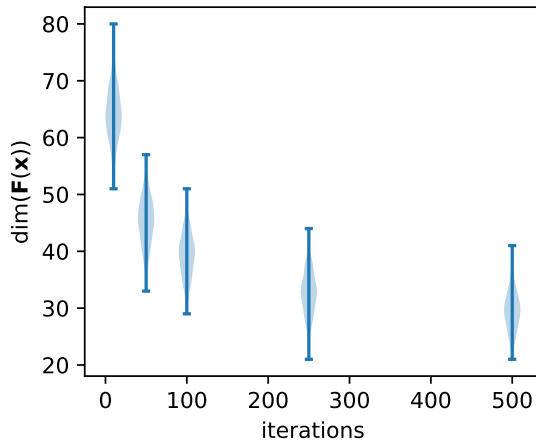
$$\pi(x | b) \propto \exp\left(-\frac{\lambda}{2}\|Ax - b\|_2^2 - \gamma\|Lx\|_1\right) \text{ versus } x | b = \operatorname{argmin}_{x \in \mathbb{R}^n} \left\{ \frac{\lambda}{2}\|Ax - \hat{b}\|_2^2 + \gamma f(x) \right\}$$





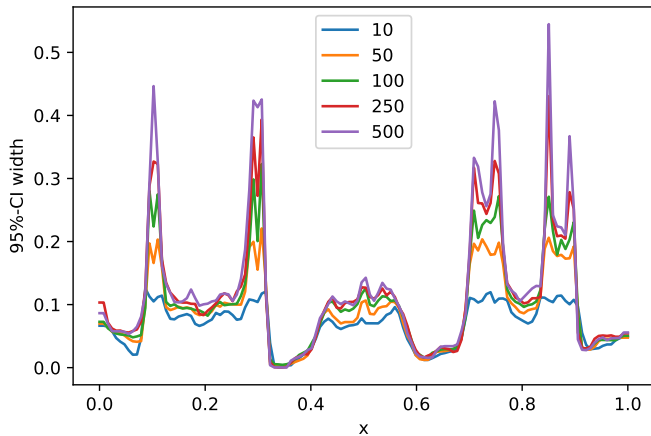
## Numerical example: Gibbs denoising with TV

Computational cost (iterations) *versus* accuracy (sparsity)



## Numerical example: Gibbs denoising

Computational cost (iterations) *versus* accuracy (componentwise credibility interval width)



To be continued

## Overview

Sample efficiently from an implicit varying dimension model:

Regularized Gaussian distribution:

$$x | b := \min_{z \in \mathbb{R}^n} \left\{ \frac{1}{2} \|Az - \hat{b}\|_{\Sigma^{-1}}^2 + f(z) \right\},$$

with  $\hat{b} \sim \mathcal{N}(b, \Sigma)$ .

Assuming  $\text{rank}(A) = n$ , then

- Positive probability on subspaces
- Conditional prior/conditionally Bayesian
- Gibbs sampler for a hierarchical model

### Well-behaved subdifferential

$$\gamma \sum_i \|D_i x - d_i\|_p$$

$$\chi_{D^n \cap \mathbb{R}_{\geq 0}^n}(x)$$

### Piecewise linear

$$\chi_{[0,1]^n}(x)$$

$$\gamma \|Lx - c\|_1$$

$$\max_i \{a_i^T x + b_i\}$$

### Positive homogeneous

$$\chi_{\mathbb{R}_{\geq 0}^n}(x)$$

$$\gamma \|Lx\|_1$$

$$\max_i \{a_i^T x\}$$

Preprint on ArXiv: *Sparse Bayesian inference with regularized Gaussian distributions*

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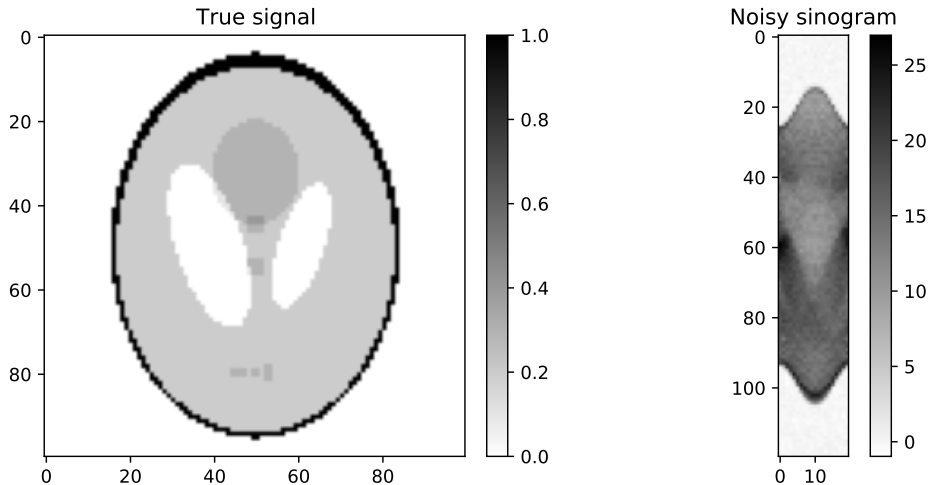
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## Numerical example: NNTV regularized CT ( $\text{rank}(A) < n$ )



# Numerical example: NNTV regularized CT ( $\text{rank}(A) < n$ )

