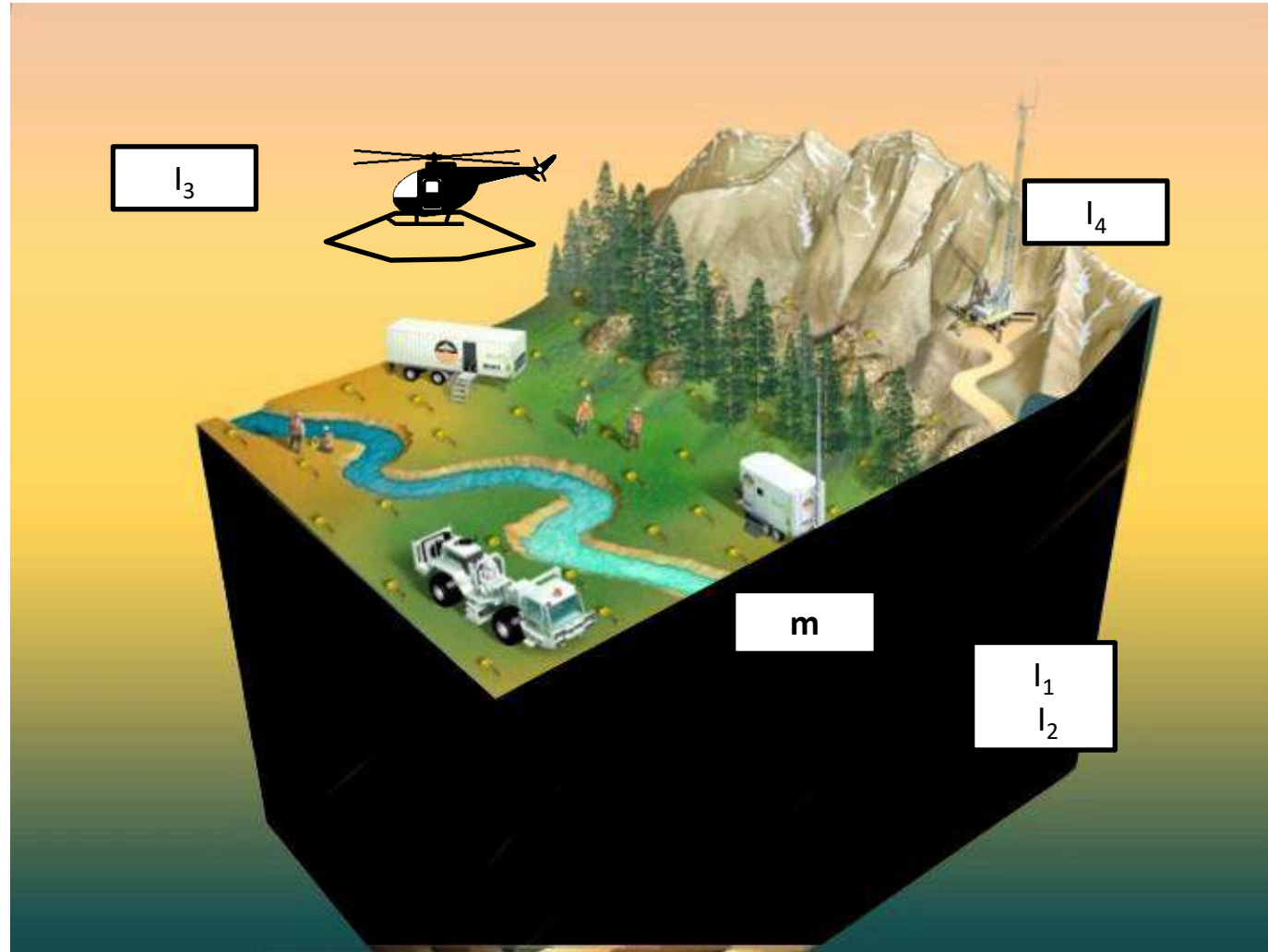
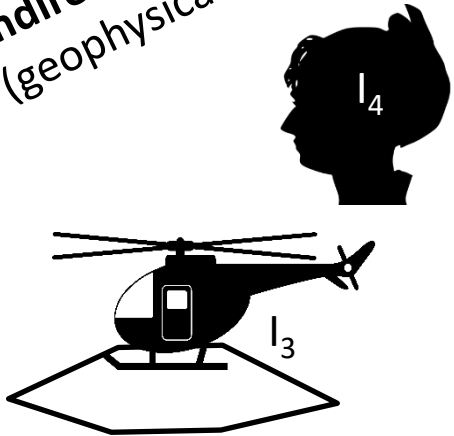


PROBABILISTIC INTEGRATION OF (GEO)-INFORMATION

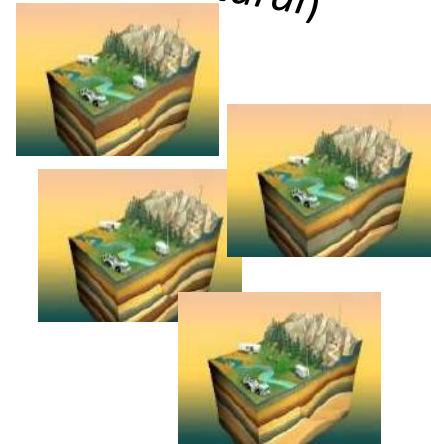
THOMAS MEJER HANSEN

INTEGRATION OF GEO-INFORMATION

I_3, I_4 :
Indirect information
(geophysical data)



I_1 :
Direct information
(prior, structural)



OUTLINE

Probabilistic integration of GEO-information

- Algorithms
- Prior Information
- Modeling error

Examples (Clay?, UXO, Pollution)

Airborne electromagnetic data

- Extended Metropolis algorithm
- Extended rejection sampling
- Machine learning

Probabilistic PET image analysis

GEOLOGY + GEOPHYSICS = ?



GEOLOGY + GEOPHYSICS = ?

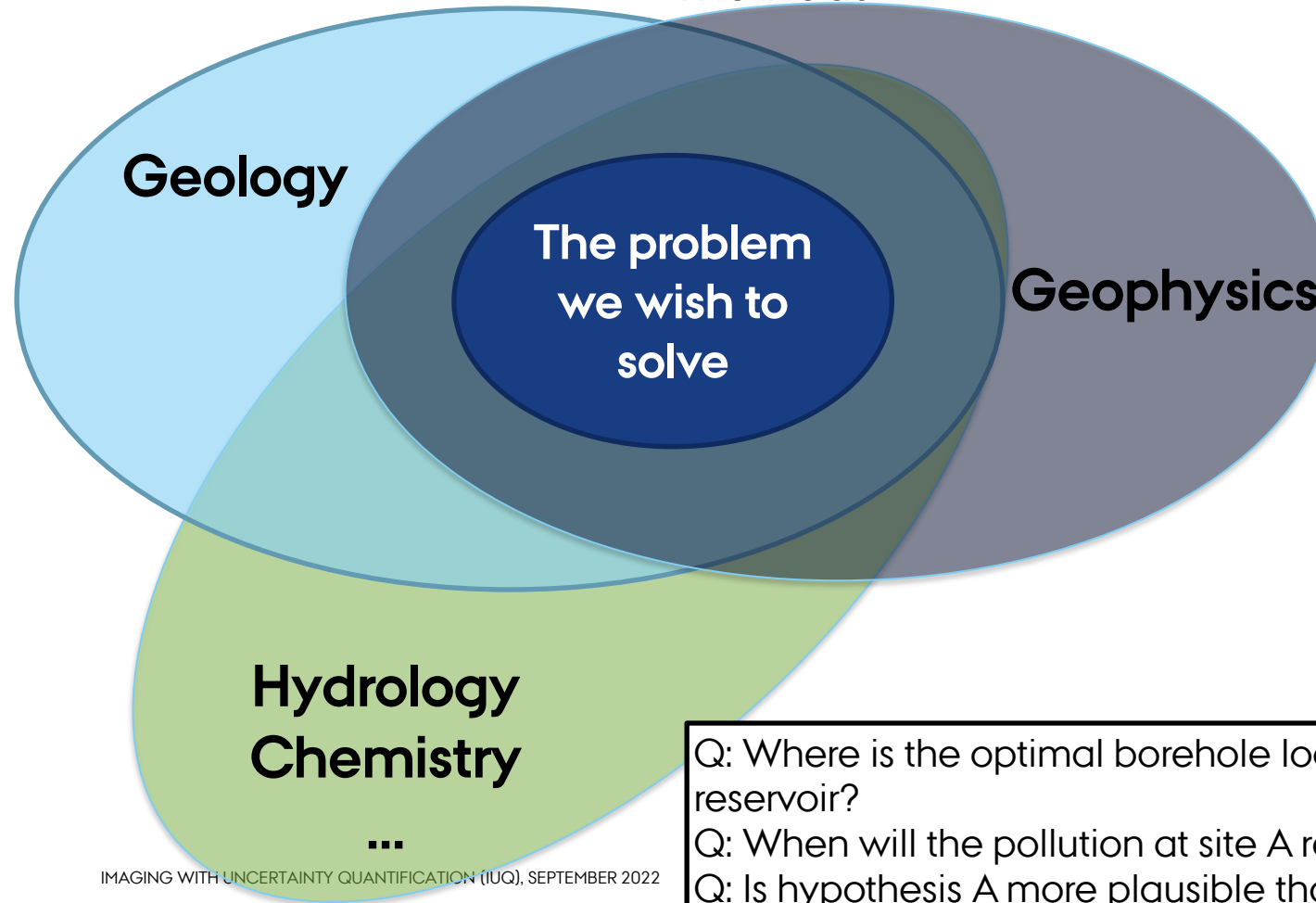
A photograph of a desert canyon with layered red rock formations under a blue sky with clouds. The rock layers are wavy and show various shades of red, orange, and tan. The text "GEOLOGY + GEOPHYSICS = ?" is overlaid in white, bold, sans-serif font across the upper middle of the image.

GEOLOGY + GEOPHYSICS = ?

GEOLOGY VS. GEOPHYSICS

— “**Geology** is the study of the Earth”

“**Geophysics** is the study of the Earth by quantitative physical methods.”

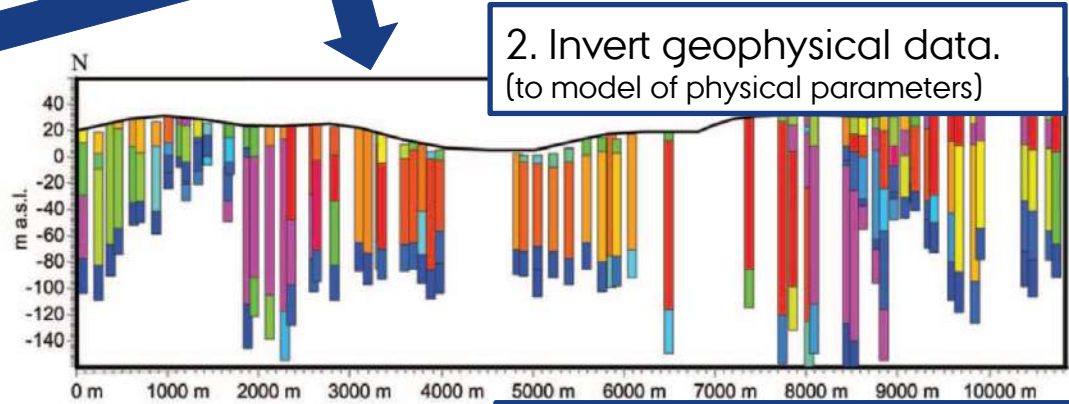
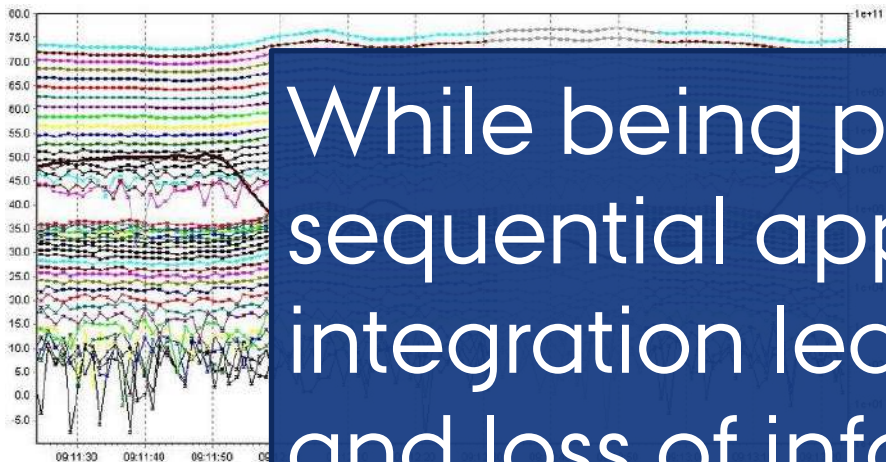


Q: Where is the optimal borehole location to target a water reservoir?
Q: When will the pollution at site A reach point B?
Q: Is hypothesis A more plausible than hypothesis B?

Sequential Information Integration in GeoScience

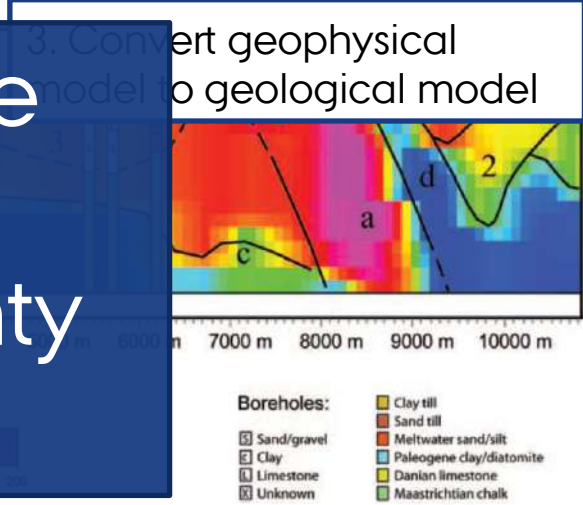
[Jørgensen et al., 2005]

1. Obtain geophysical data



2. Invert geophysical data.
(to model of physical parameters)

While being practical and useful, the sequential approach to information integration leads to loss of uncertainty and loss of information



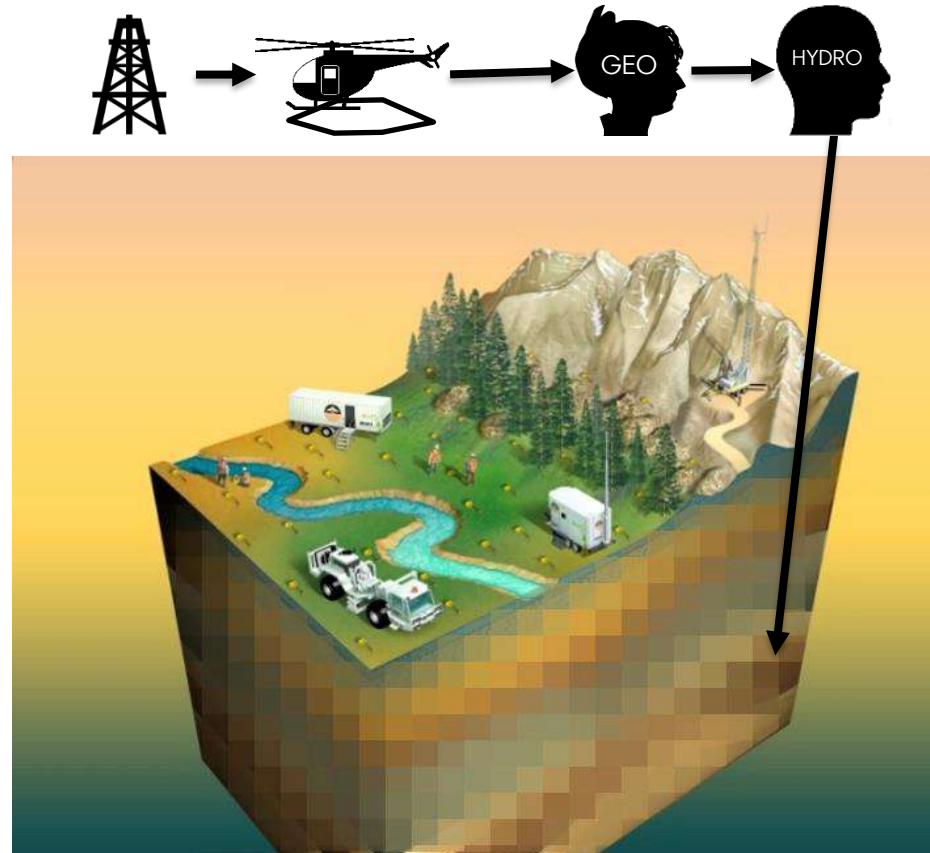
3. Convert geophysical model to geological model

4. Convert geological model to hydrological model

5. Make decisions

Sequential workflow

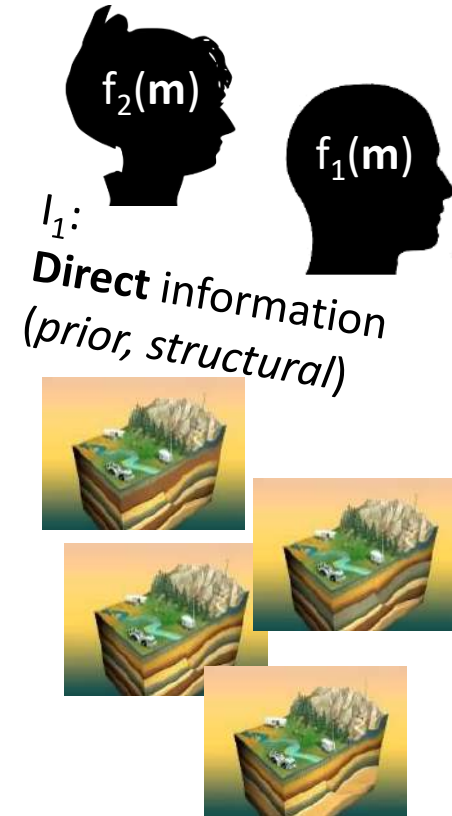
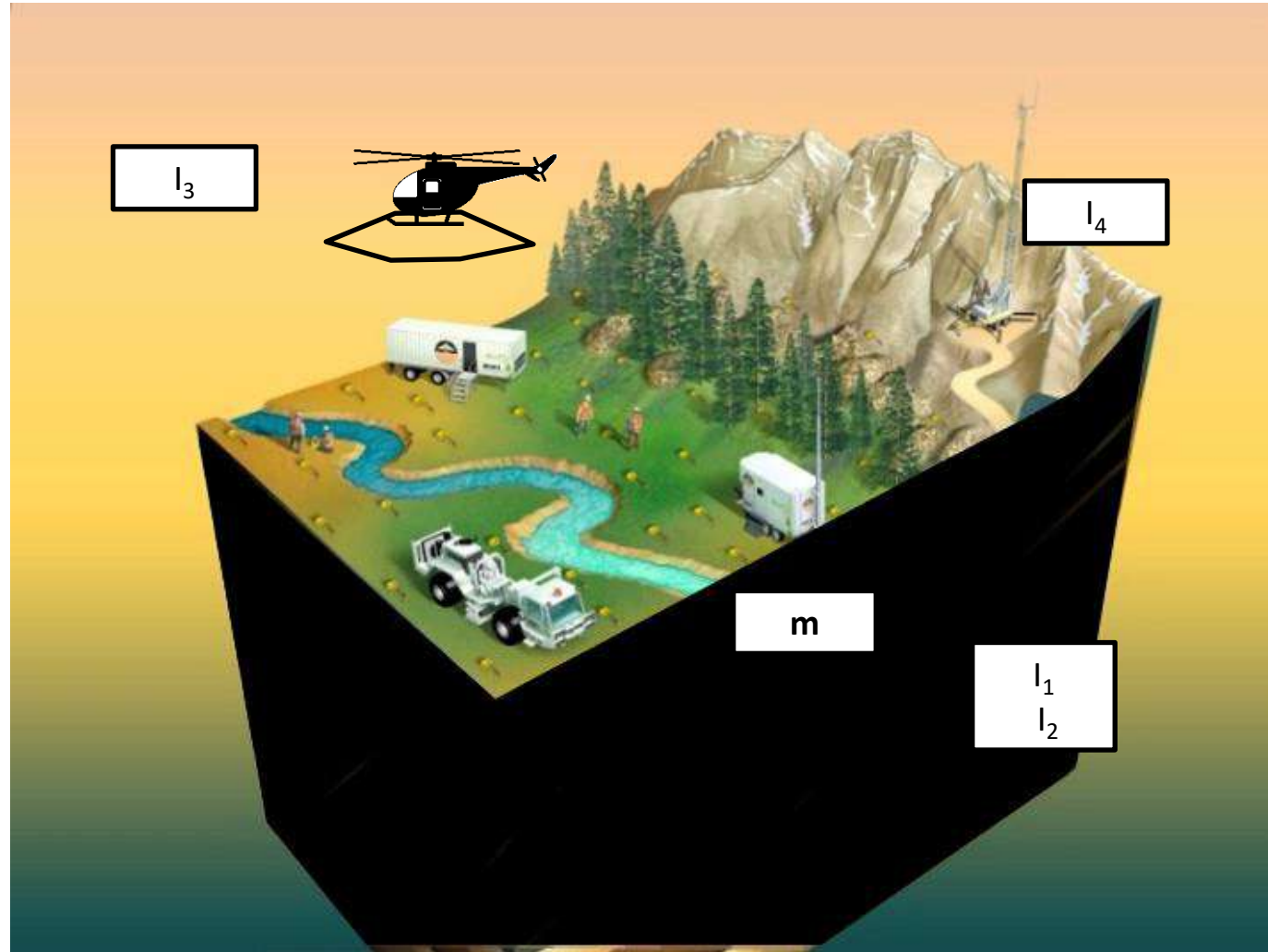
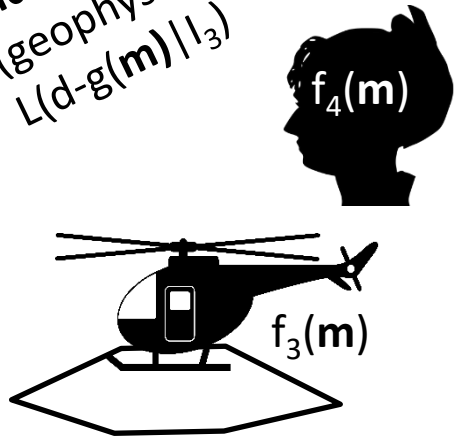
One optimal (smooth) model.
Mathematical regularization.
Consistency ?



PROBABILISTIC INTEGRATION OF GEO-INFORMATION

$$f_1(m) = f(m \mid l_1, l_2, l_3, l_4 \dots) = k f_1(m) f_2(m) f_3(m) f_4(m) \dots$$

l_3, l_4 :
Indirect information
(geophysical data)
 $L(d-g(m) \mid l_3)$



INFORMATION



→ $I_1: f_1(m)$



→ $I_2: f_2(m)$



→ $I_3: f_3(m)$



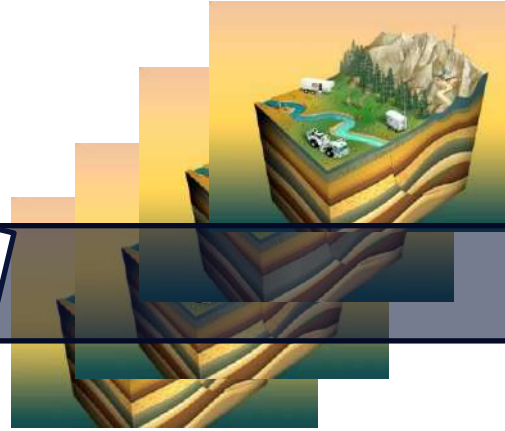
→ $I_4: f_4(m)$



→ $I_5: f_5(m)$

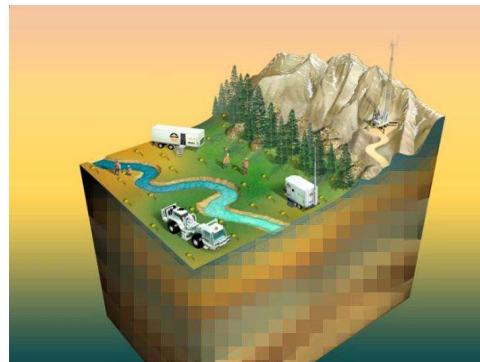
INVERSION / INTEGRATION
ALGORITHM

Data Integration: $f_1(m)$



Calibration

Deterministic IP, m^{optimal}



INFORMED DECISIONS

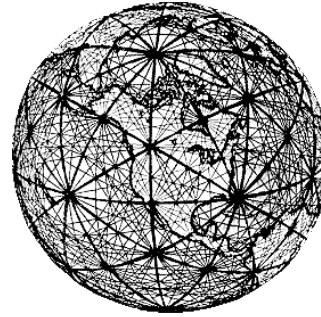
What if ... ?

Is the groundwater safe to using pesticides at the surface?

Will the pollution reach XXX?

CONJUNCTION OF INFORMATION

$$\mathbf{m} = [m_1, m_2, \dots]$$



[parameterized Earth]

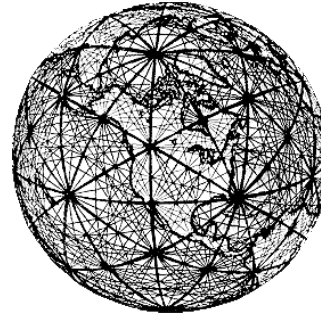
$$f_1(\mathbf{m}) \wedge f_2(\mathbf{m}) \wedge f_3(\mathbf{m}) \wedge f_4(\mathbf{m}) \wedge \dots \propto \prod_{i=1}^{N^i} f_i(\mathbf{m})$$

$f_i(\mathbf{m})$ must be obtained independently from $f_j(\mathbf{m})$, for all sets of $[i,j]$.

Tarantola and Valette, 1982: Inverse Problems = Quest for information

CONJUNCTION OF INFORMATION

$$\mathbf{m} = [m_1, m_2, \dots]$$



[parameterized Earth]

$$\sigma_M(\mathbf{m}) = k \rho_M(\mathbf{m}) L(\mathbf{m})$$

$$L(\mathbf{m}) = \int_D d\mathbf{d} \frac{\rho_D(\mathbf{d})\theta(\mathbf{d}|\mathbf{m})}{\mu_D(\mathbf{d})}$$
$$\approx \rho_D(\mathbf{g}(\mathbf{m}))$$

Tarantola and Valette, 1982: Inverse Problems = Quest for information

THE EXTENDED METROPOLIS ALGORITHM

The goal: Sample from $\sigma(\mathbf{m}) = k \rho(\mathbf{m}) L(\mathbf{m})$

0. Propose a starting from $\rho(\mathbf{m}) \rightarrow \mathbf{m}_{\text{cur}}$
 - 1. Propose a model from $\rho(\mathbf{m}) \rightarrow \mathbf{m}_{\text{pro}}$ *in the vicinity of \mathbf{m}_{cur}*
 - 2. Accept the move from \mathbf{m}_{cur} to \mathbf{m}_{pro} with probability
 - $P_{\text{acc}} = L(\mathbf{d}_{\text{obs}} - \mathbf{g}(\mathbf{m}_{\text{pro}})) / L(\mathbf{d}_{\text{obs}} - \mathbf{g}(\mathbf{m}_{\text{cur}}))$
 - 3. Store the current model, \mathbf{m}_{cur} .
4. Go to 1 (until enough realizations have been generated).

To apply the extended Metropolis algorithm, one must be able to

- 1) Sample the prior $\rho(\mathbf{m})$, through a random walk.
- 2) Evaluate the likelihood $L(\mathbf{m})$ for any model, (solve the forward problem, evaluate the noise)

THE EXTENDED REJECTION SAMPLER

The goal: Sample from $\sigma(\mathbf{m}) = k \rho(\mathbf{m}) L(\mathbf{m})$

1. Propose a model from $\rho(\mathbf{m}) \rightarrow \mathbf{m}^*$
2. Accept \mathbf{m}^* as a realization from the posterior with probability

$$P_{\text{acc}} = L(\mathbf{d}_{\text{obs}} - g(\mathbf{m}^*)) / \max(L)$$

To apply the extended rejection sampler one must be able to

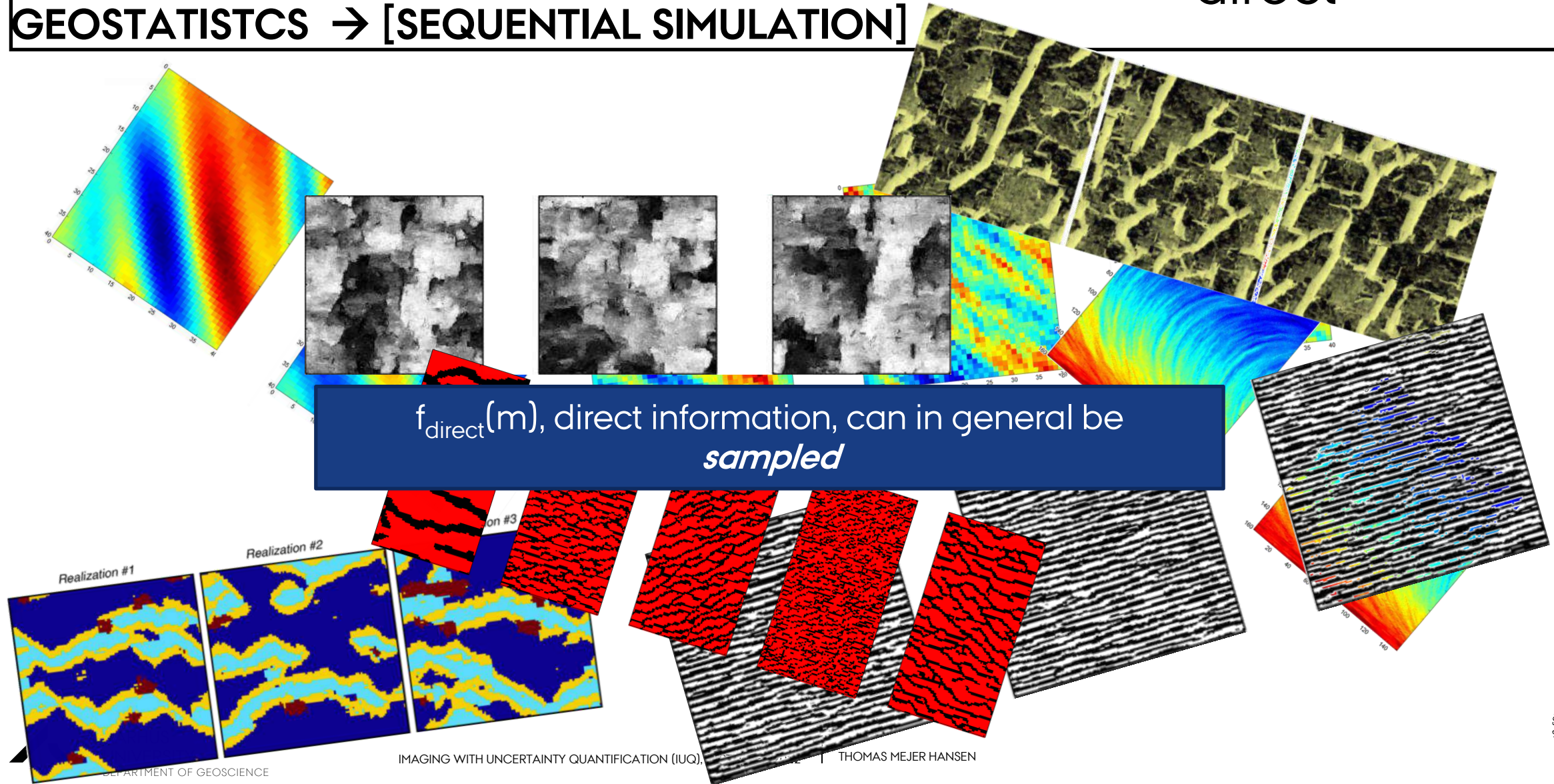
- 1) Sample the prior $\sigma(\mathbf{m})$
- 2) Evaluate the likelihood $L(\mathbf{m})$ for any model, (solve the forward problem, evaluate the noise)

Lookup table: $[\mathbf{M}^*, \mathbf{D}^*] \rightarrow$ Compute once, apply for any \mathbf{d}_{obs}

DIRECT (PRIOR) INFORMATION, $f_{\text{direct}}(\mathbf{m})$

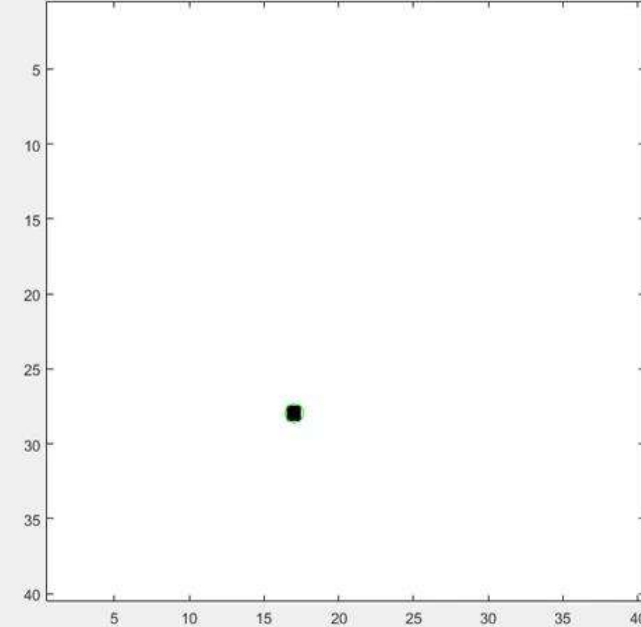
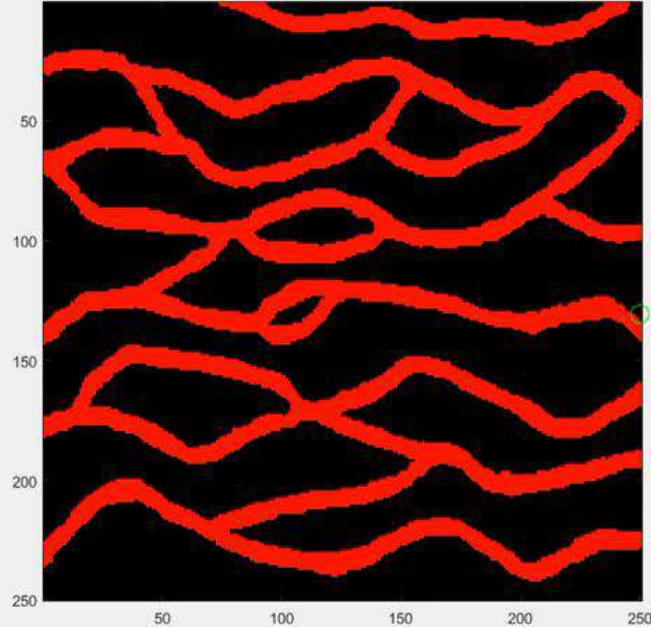
GEOSTATISTICS → [SEQUENTIAL SIMULATION]

$f_{\text{direct}}(\mathbf{m})$, direct information, can in general be
sampled



Sequential Simulation

$$f(\mathbf{m}) = f(m_1) f(m_2 | m_1^*) f(m_3 | m_1^*, m_2^*) \dots$$

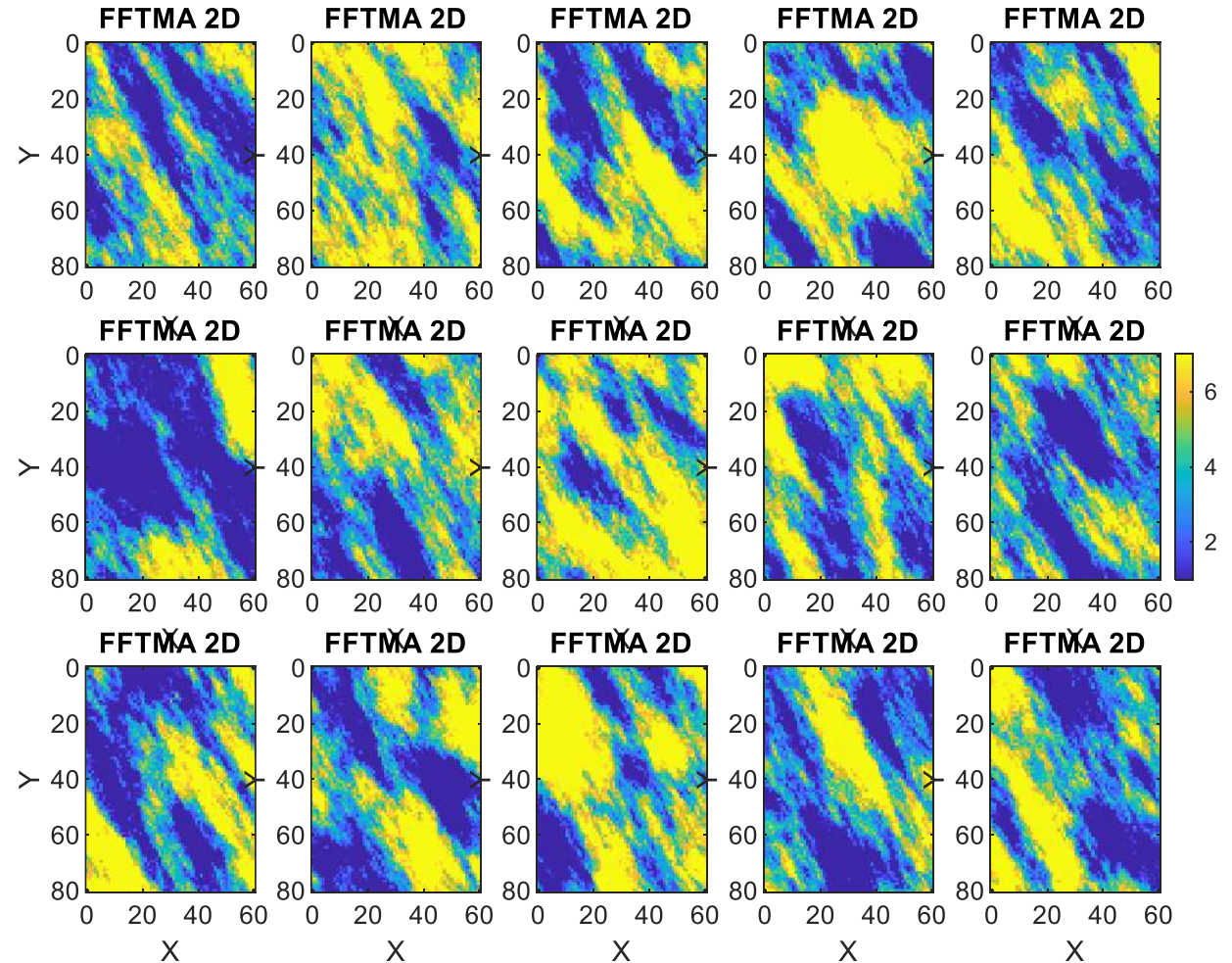


2D Gaussian prior

```
ip=1;  
prior{ip}.type='FFTMA';  
prior{ip}.name='FFTMA 2D';  
prior{ip}.x=0:1:60;  
prior{ip}.y=0:1:80;  
prior{ip}.m0 = 4;  
prior{ip}.Va='10 Sph(10,90,0.3)';  
sippi_plot_prior_sample(prior);
```

From:

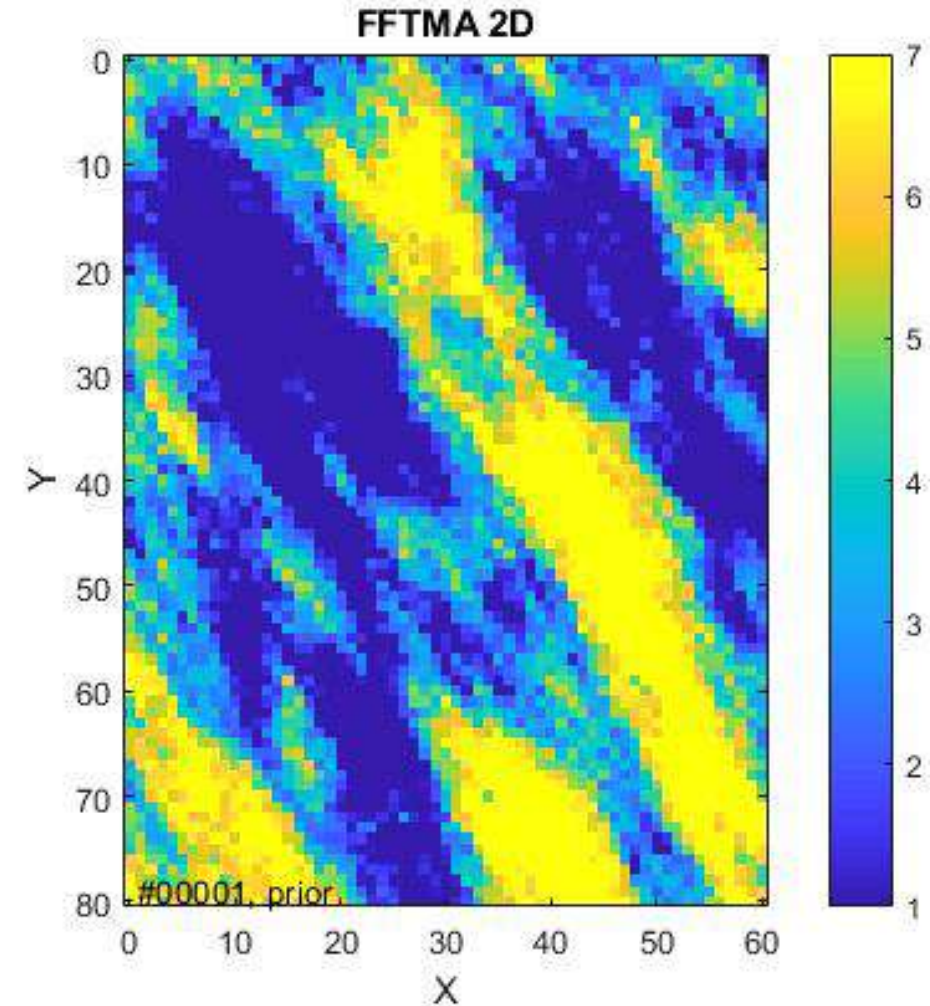
<https://github.com/cultpenguin/sippi>



2D Gaussian prior – random walk

```
ip=1;
prior{ip}.type='FFTMA';
prior{ip}.name='FFTMA 2D';
prior{ip}.x=0:1:80;
prior{ip}.y=0:1:60;
prior{ip}.m0 = 4;
prior{ip}.Va='10 Sph(10,90,0.3)';

prior{ip}.seq_gibbs.step = .02;
sippi_plot_prior_movie(prior,200);
```



2D Gaussian prior – random walk – uncertain covariance model parameters

```
clear prior
ip=1;
prior{ip}.type='FFTMA';
prior{ip}.name='FFTMA 2D';
prior{ip}.x=0:1:60;
prior{ip}.y=0:1:80;
prior{ip}.m0 = 4;
prior{ip}.Va='10 Sph(50,30,0.3)';
prior{ip}.cax=[-3 3]+4;

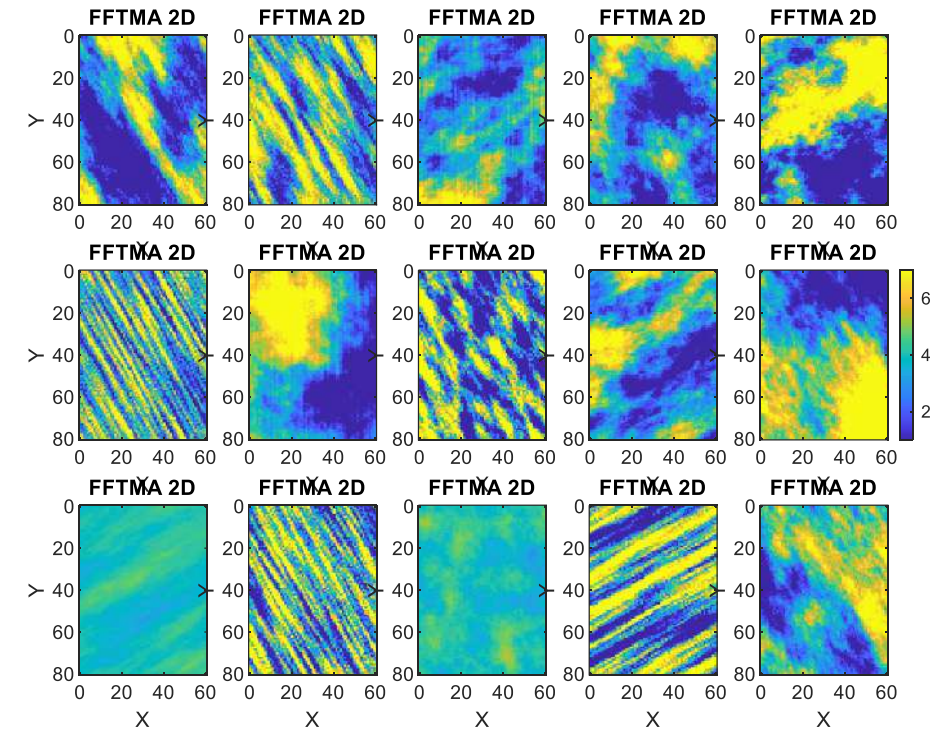
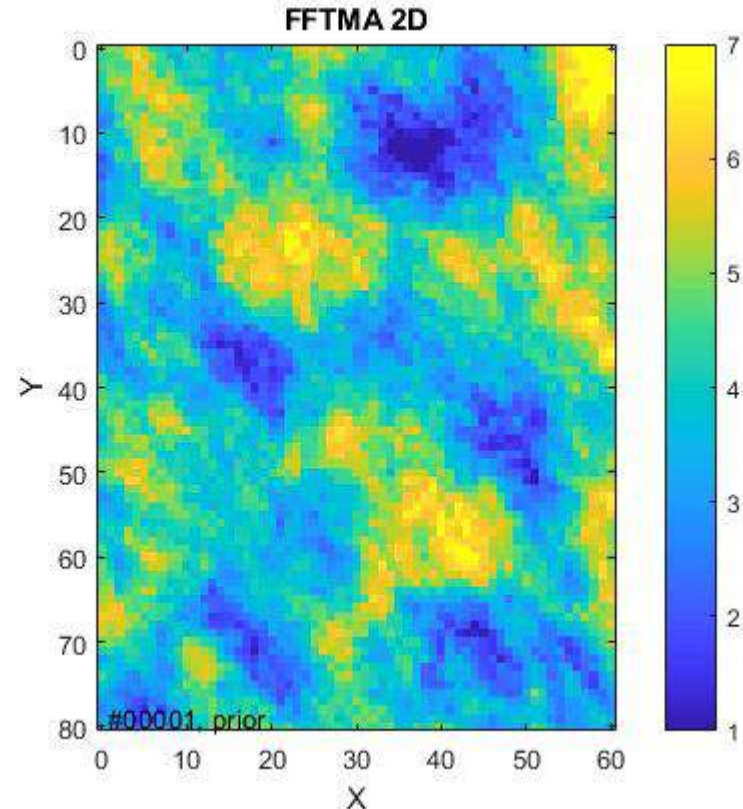
ip=2;
prior{ip}.name='range_1';
prior{ip}.type='uniform';
prior{ip}.min=1;
prior{ip}.max=80;
prior{ip}.prior_master=1;

ip=3;
prior{ip}.name='range_2';
prior{ip}.type='uniform';
prior{ip}.min=1;
prior{ip}.max=80;
prior{ip}.prior_master=1;

ip=4;
prior{ip}.name='angle_2';
prior{ip}.type='uniform';
prior{ip}.min=-30;
prior{ip}.max=30;
prior{ip}.prior_master=1;

ip=5;
prior{ip}.name='sill';
prior{ip}.type='uniform';
prior{ip}.min=0;
prior{ip}.max=10;
prior{ip}.prior_master=1;

sippi_plot_prior_movie(prior,200);
```

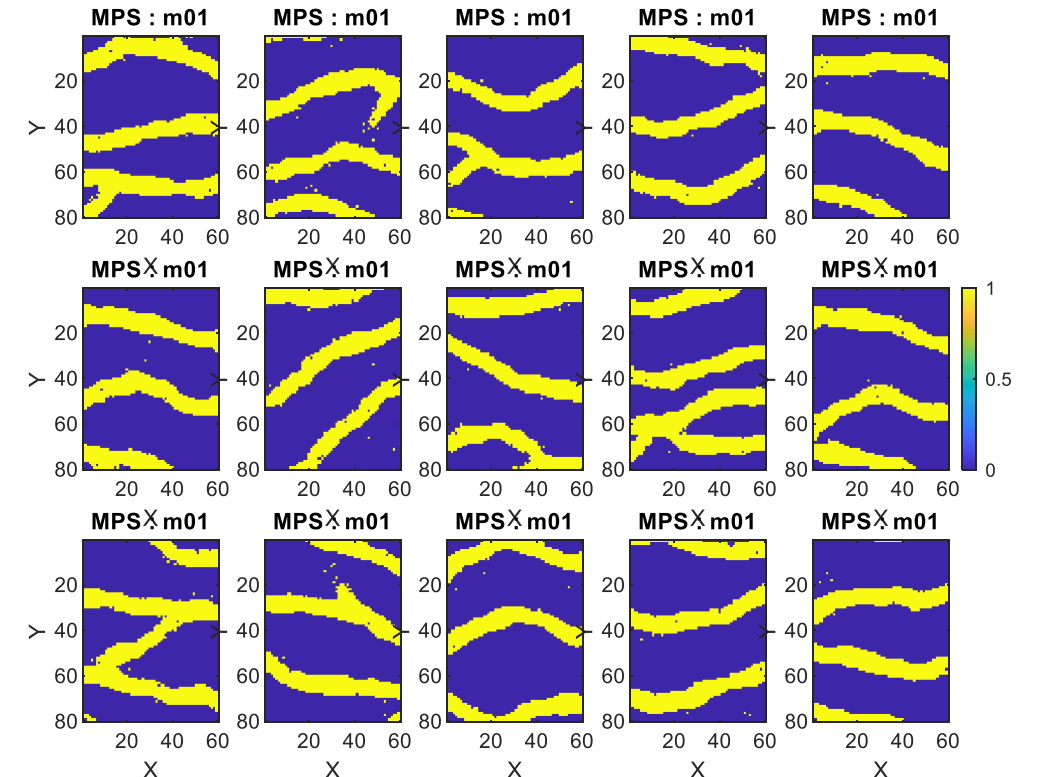
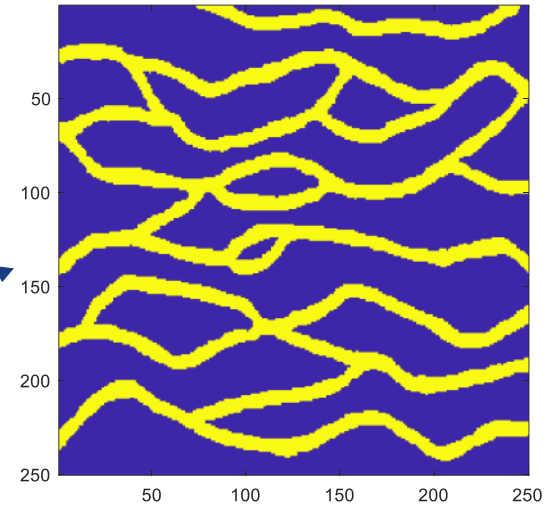


2D Multiple Point Statistical prior

```
clear prior
ip=1;
prior{ip}.type='mps';
prior{ip}.method='mps_genesim';
prior{ip}.x=1:1:60;
prior{ip}.y=1:1:80;
prior{ip}.ti=channels;

figure(1);
imagesc(prior{ip}.ti);axis image

sippi_plot_prior_sample(prior);
```



PRIOR ASSUMPTION \leftrightarrow SUBJECTIVE INFORMATION

The use/need of prior information has been debated, Scales and Tenorio (1997)

we try to find a **noninformative**, or conservative, **prior** that **injecting a minimum of artificial information**, that is, information not justified by the physical process.

Can, and should, we avoid prior information?



HOW TO AVOID PRIOR/STRUCTURAL INFORMATION?

$\mathbf{m}=[m_1, m_2 \dots, m_{10000}]$ represents a 2D gray-scale image.

$f(\mathbf{m})$ represents our information about \mathbf{m} .

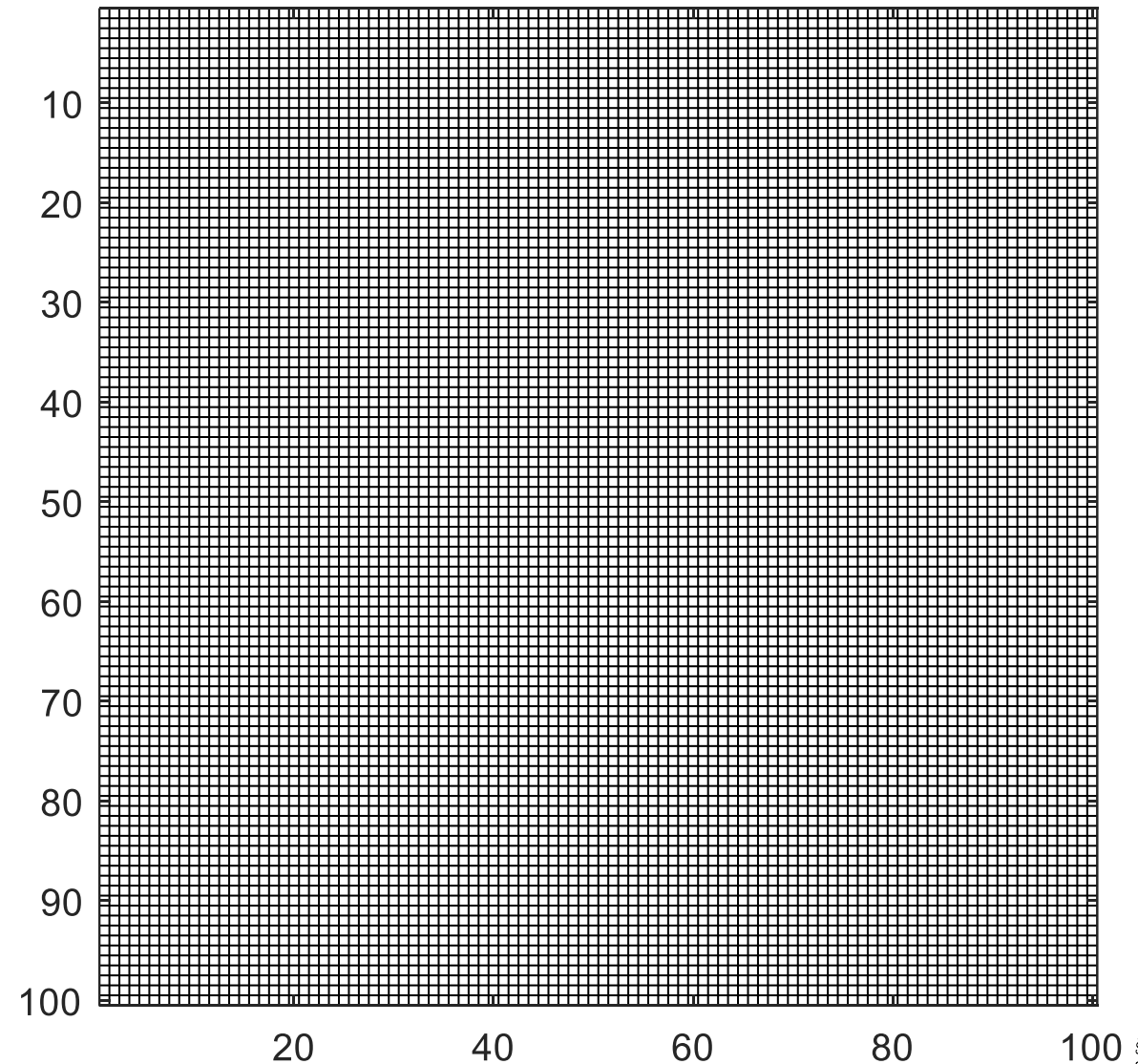
A realization \mathbf{m}^* of $f(\mathbf{m})$ represents a specific image.

1. Assume there is no dependence between individual model parameters

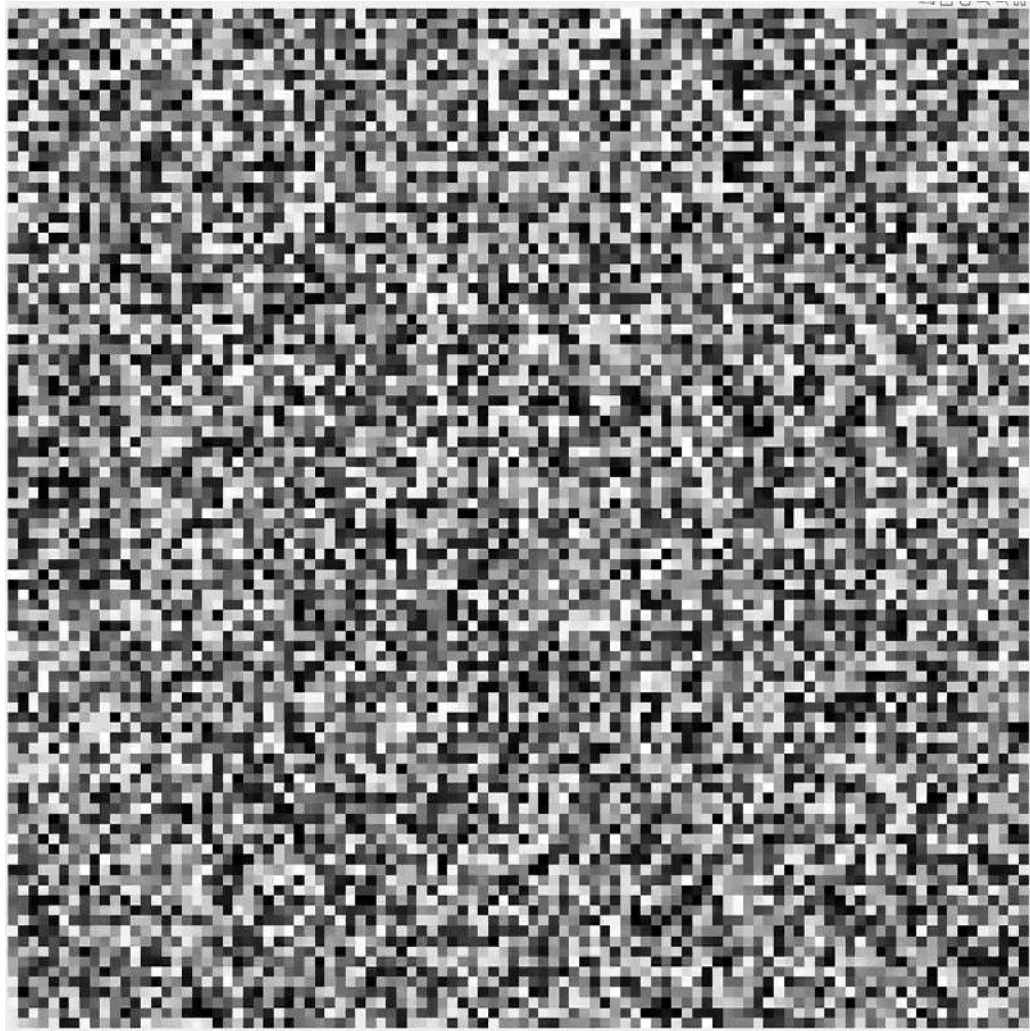
$$f(m_i, m_j)=f(m_i)f(m_j)$$

2. Use a uniform prior distribution

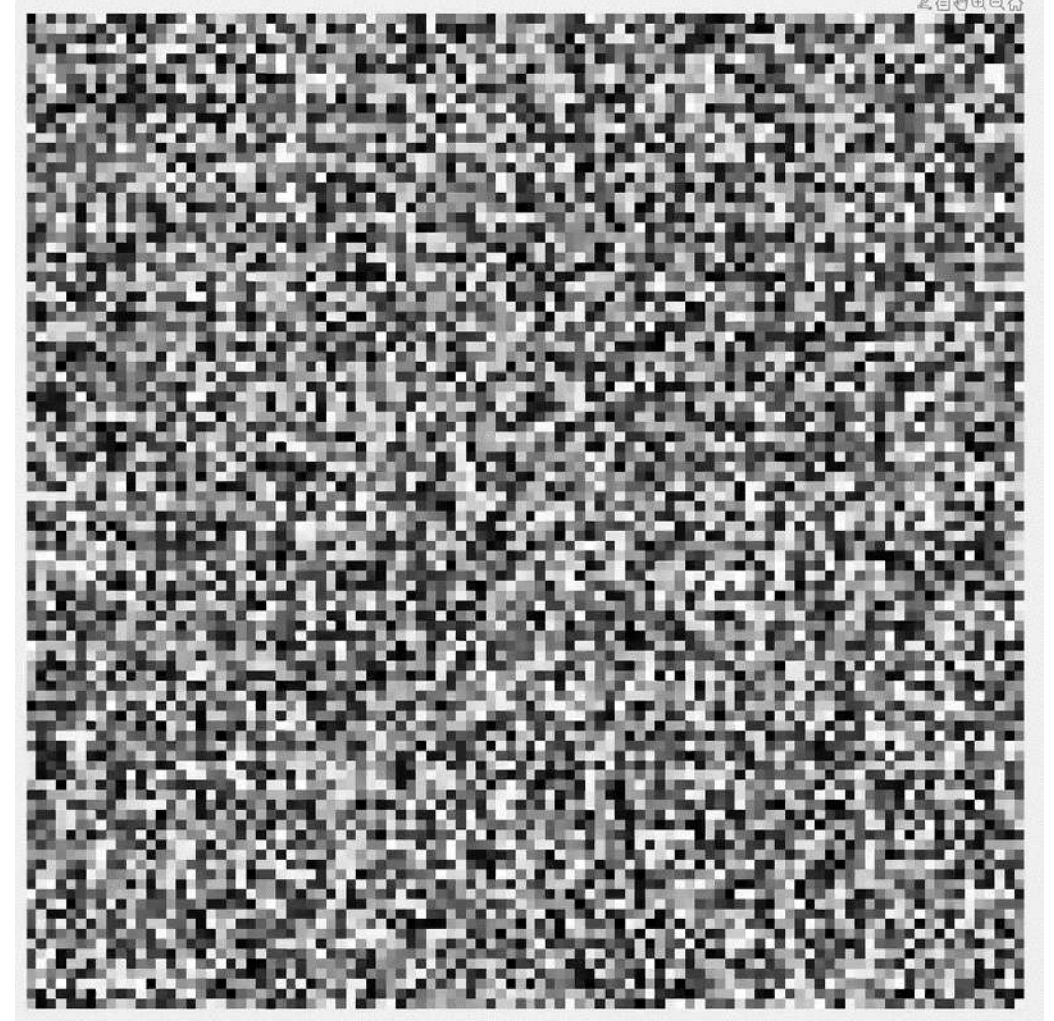
$$f(m_i) = U[0,255]$$

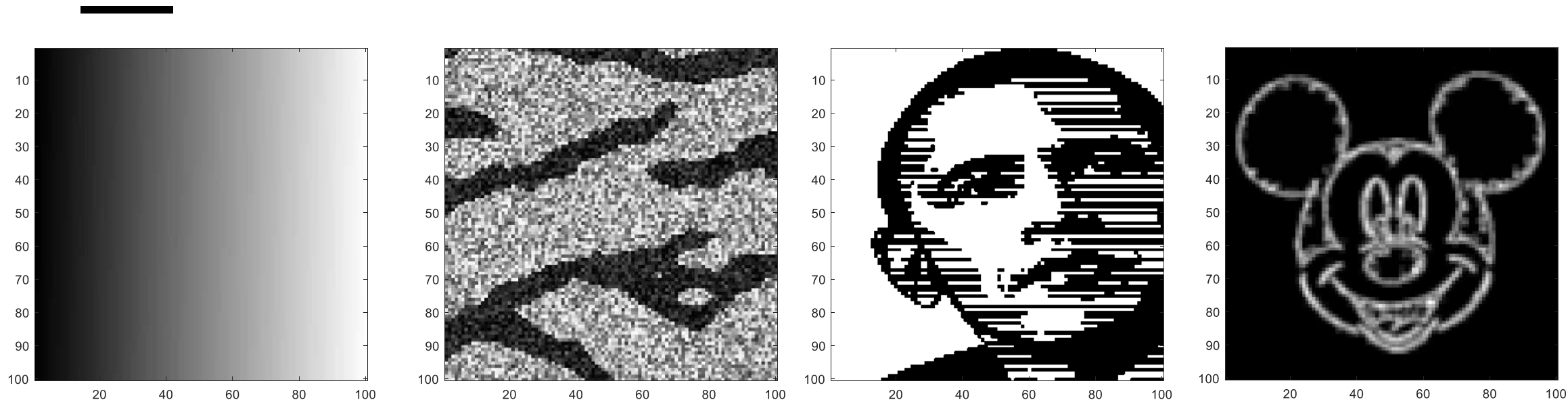


Random walk in $f(\mathbf{m})$



Guided random walk in $f(\mathbf{m})$





$P(\text{Image looks somewhat like Obama}) < 1/250000 = 0.000004$

$P(\text{Image looks somewhat like any structure}) < 0.000004$

$P(\text{Image look like random noise}) > 0.999996$

Key take home message:

YES, an uniform prior can represent in principle any image/structure.

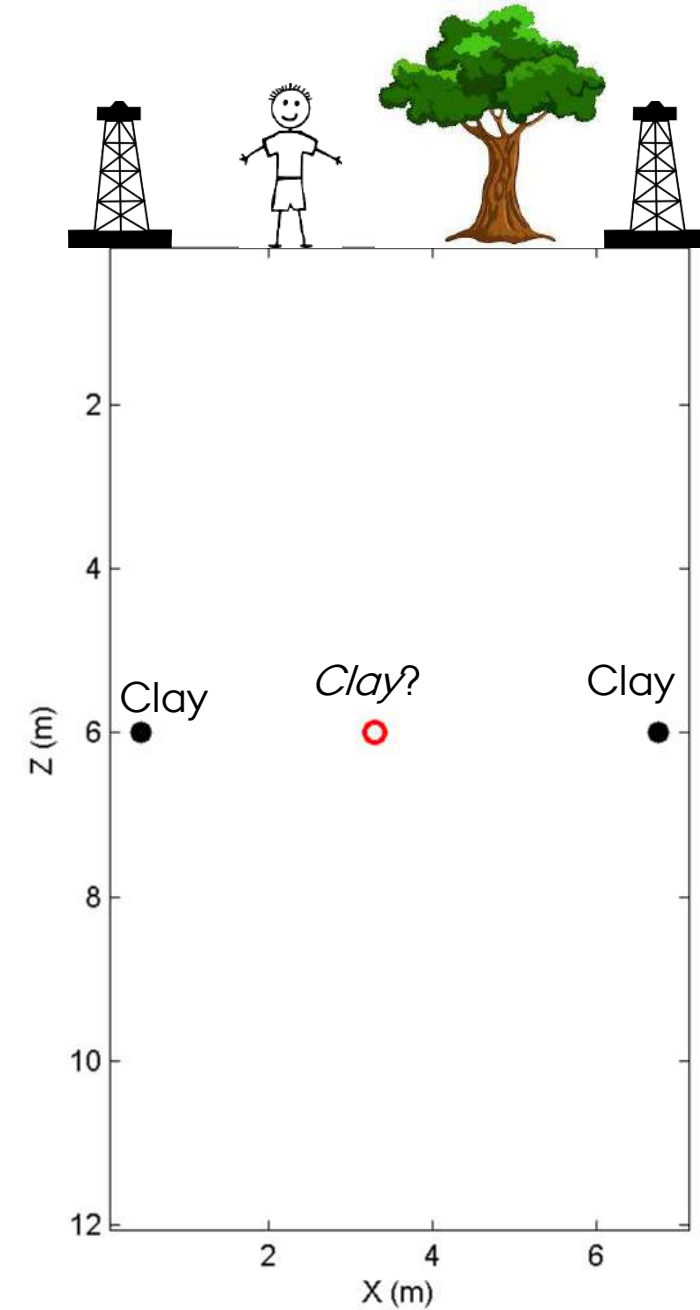
BUT, the probability of any other feature than random noise exists, tend to zero (as the pixel size becomes smaller).

One cannot assume a uniform prior in order to avoid choosing a prior

A uniform prior is not an “uninformed prior” but a specific choice of maximum disorder!

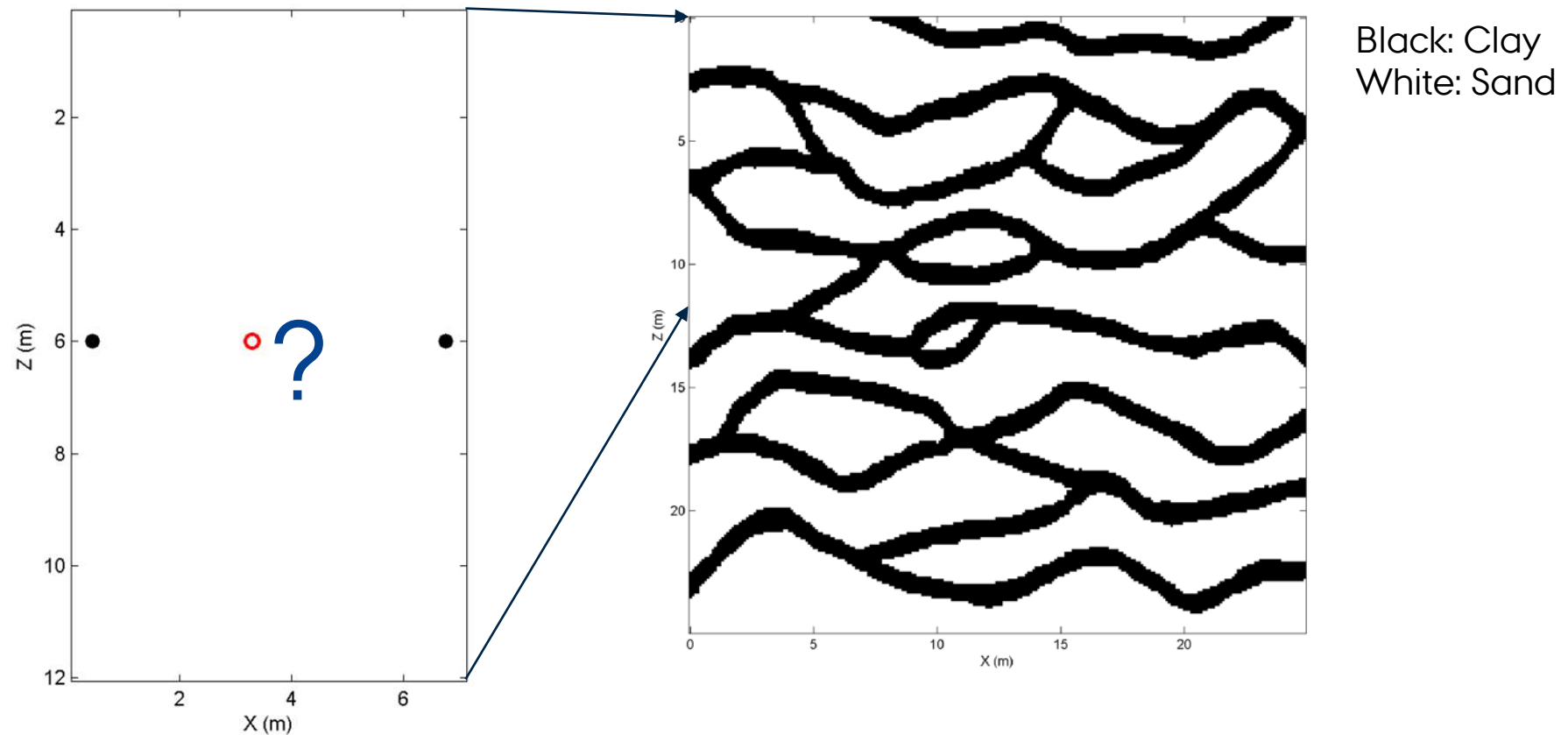
CASE: CLAY OR NOT

EXAMPLE 1: PROBABILITY OF CLAY

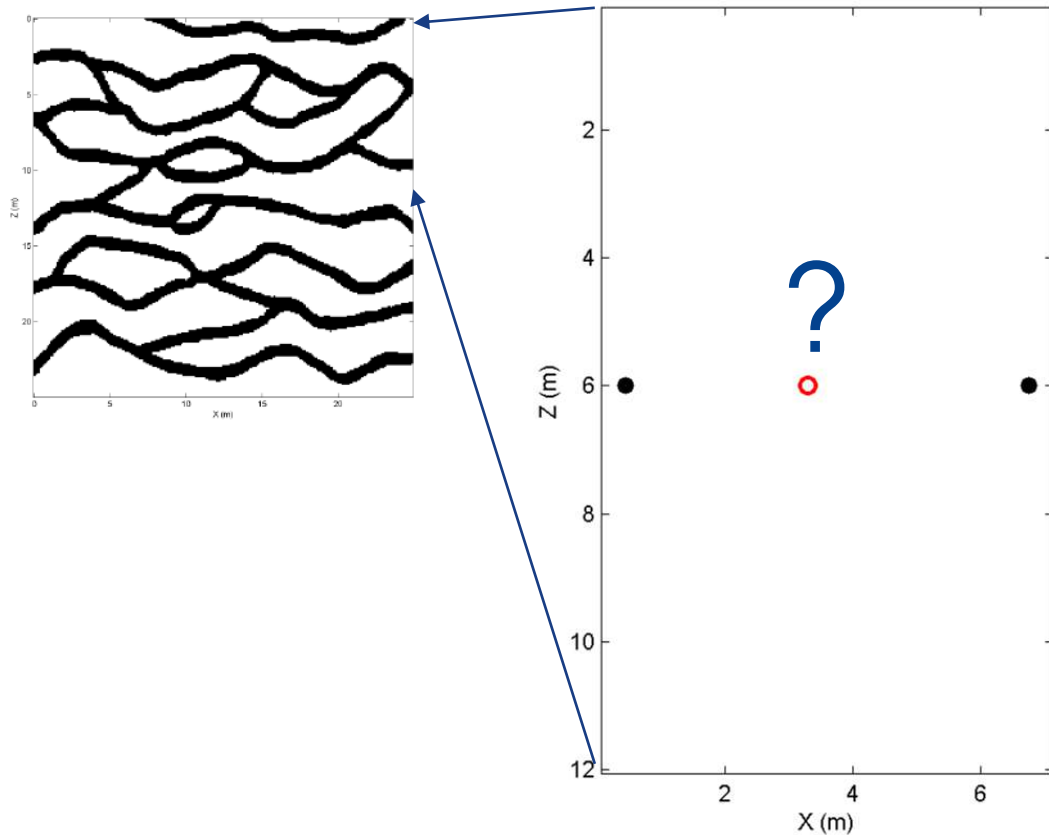


A simple application of geostatistics and inverse problem theory

- We have 2 observations:
- I_1 : Structural information from a training image:



A simple application of geostatistics and inverse problem theory



Prior probability of channel/clay:

$$P(\text{red circle} == \text{clay} \mid \text{prior}) = 0.30$$

Posterior probability of channel:

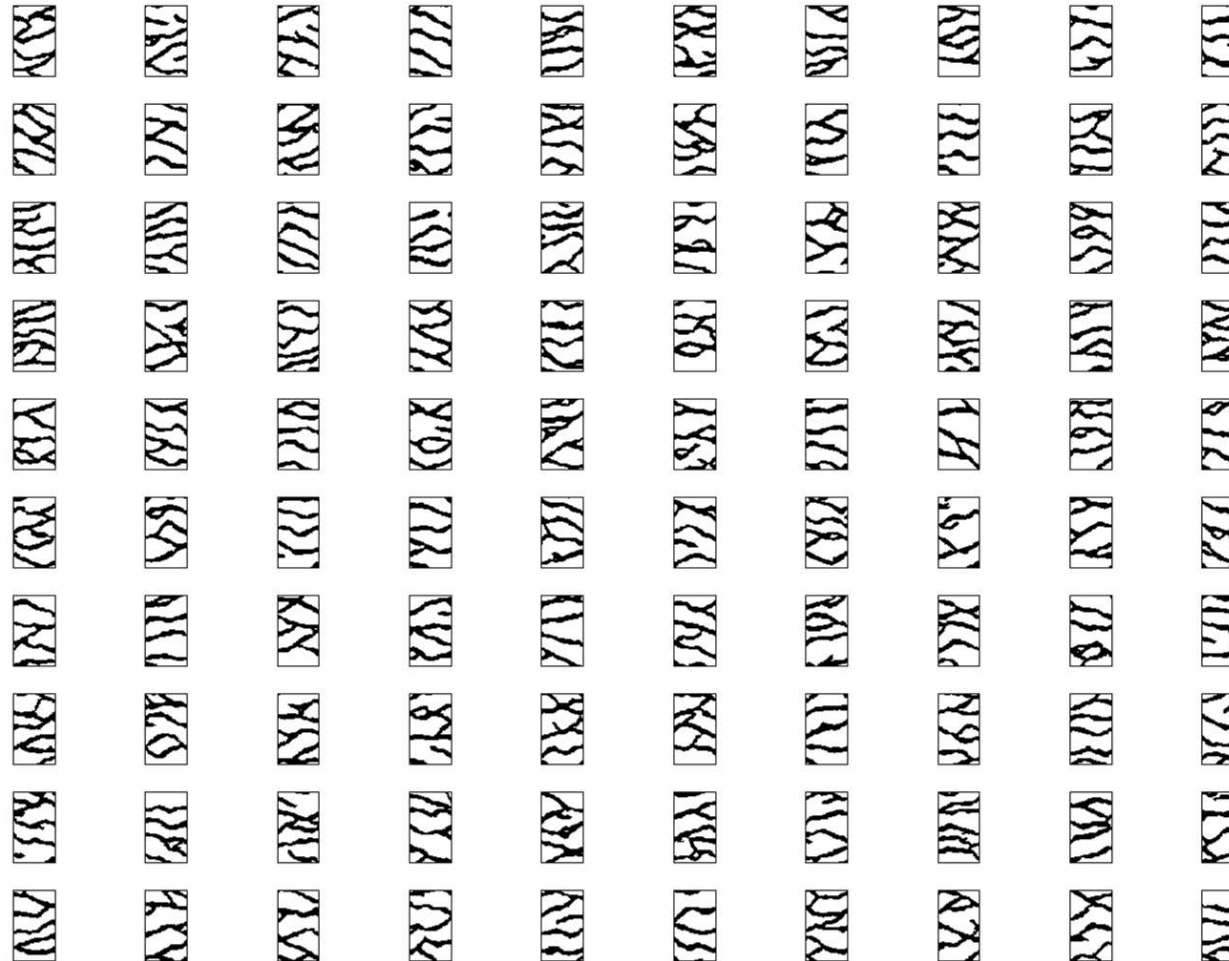
$$P(\text{red circle} == \text{clay} \mid \text{prior, data}) = ?$$

QUIZ:

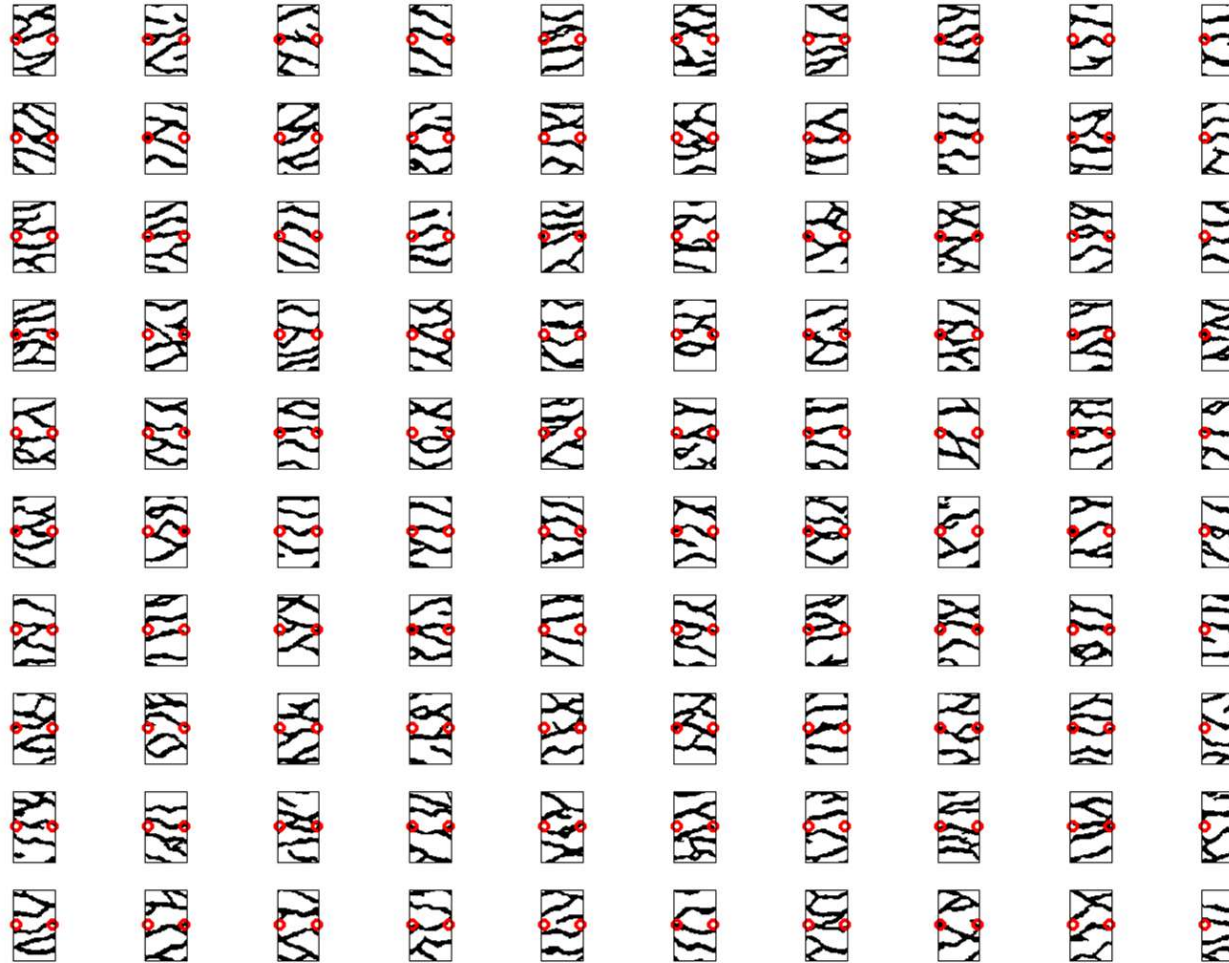
$$P(\text{red circle} == \text{clay} \mid \text{prior, data}) < 0.30$$

$$P(\text{red circle} == \text{clay} \mid \text{prior, data}) \geq 0.30$$

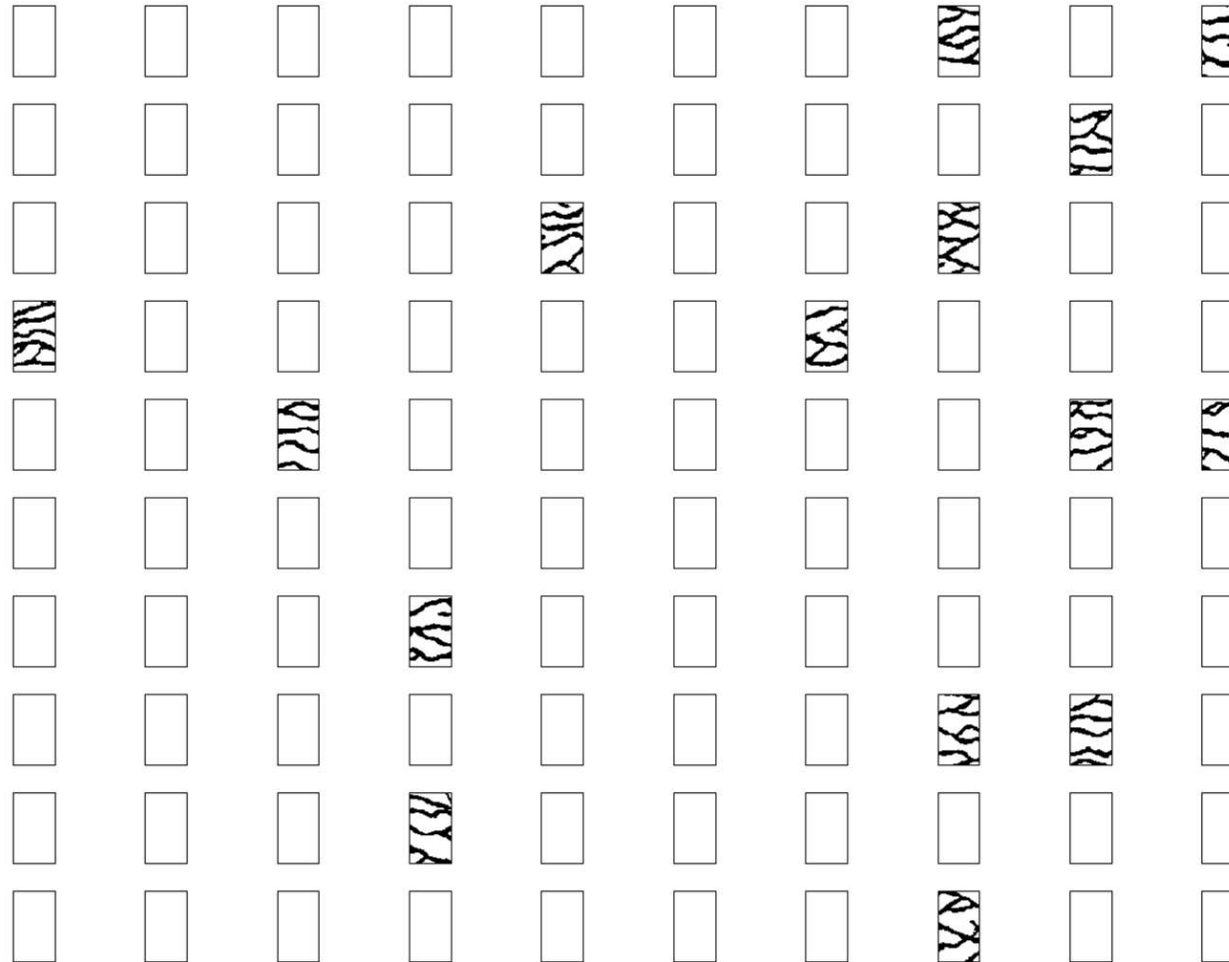
A simple application of geostatistics and inverse problem theory



A simple application of geostatistics and inverse problem theory

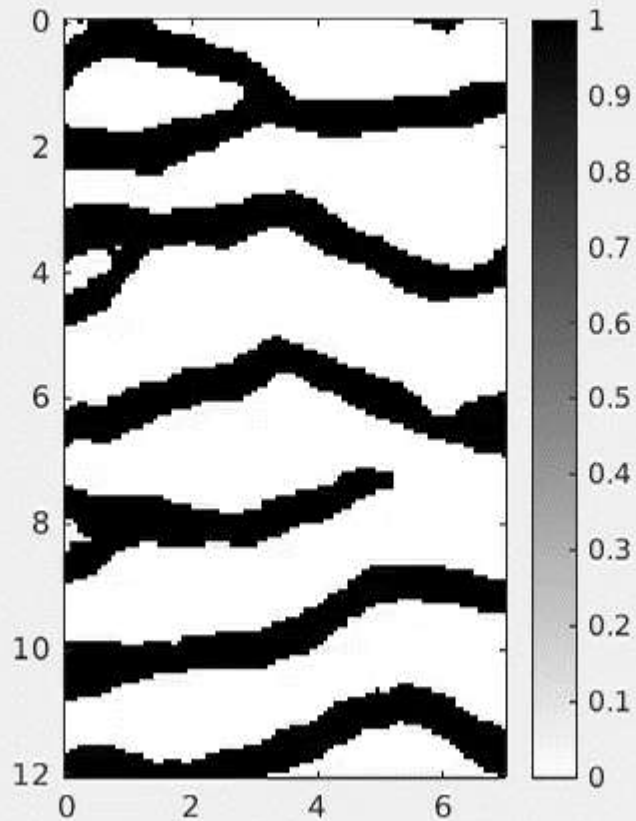


A simple application of geostatistics and inverse problem theory

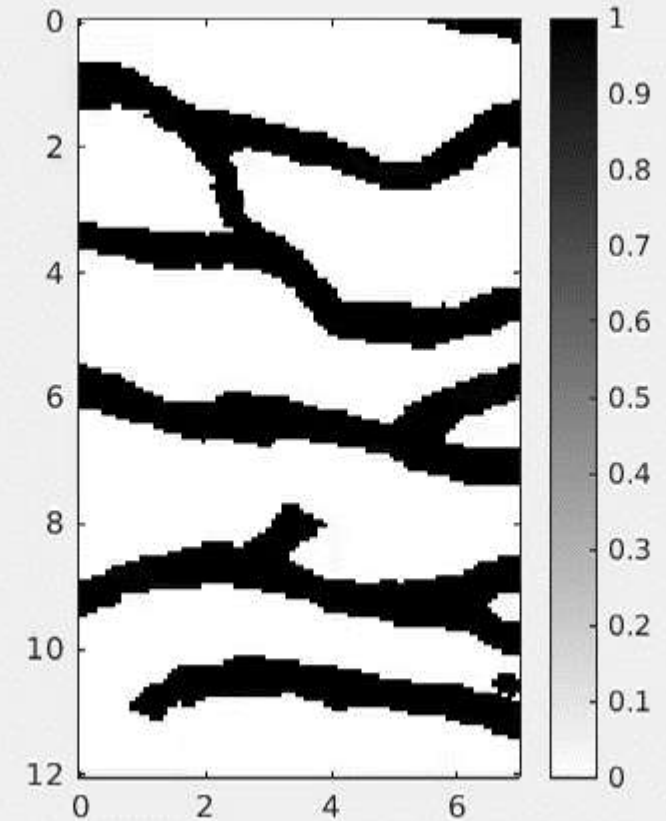


A simple application of geostatistics and inverse problem theory

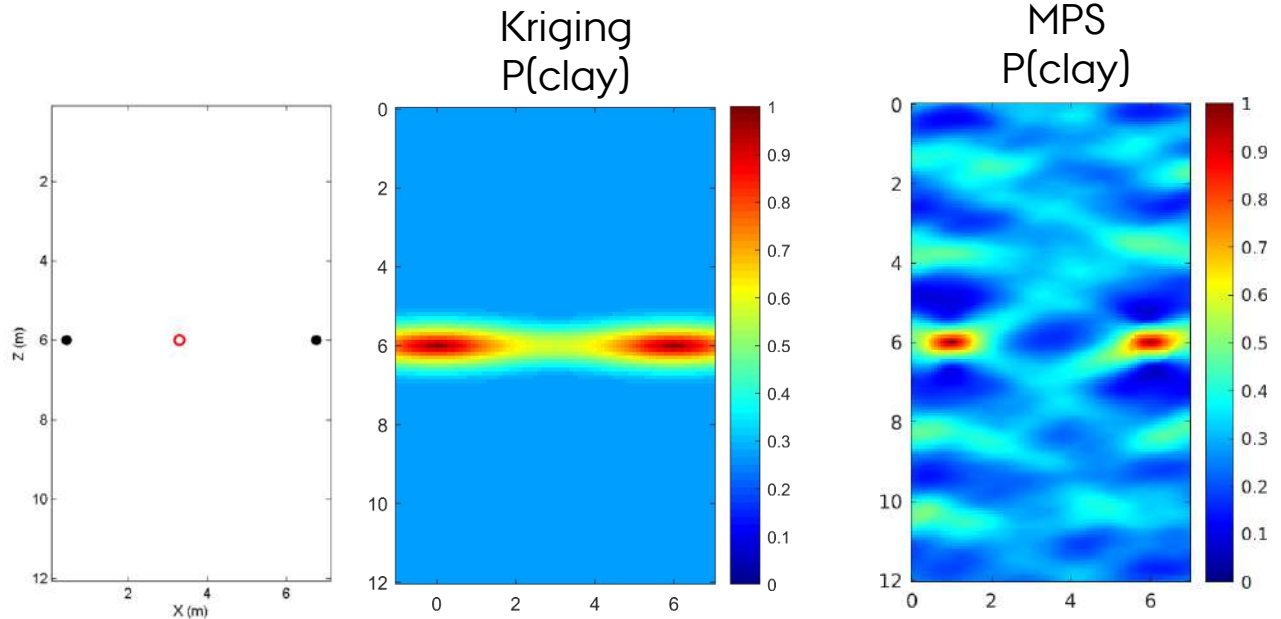
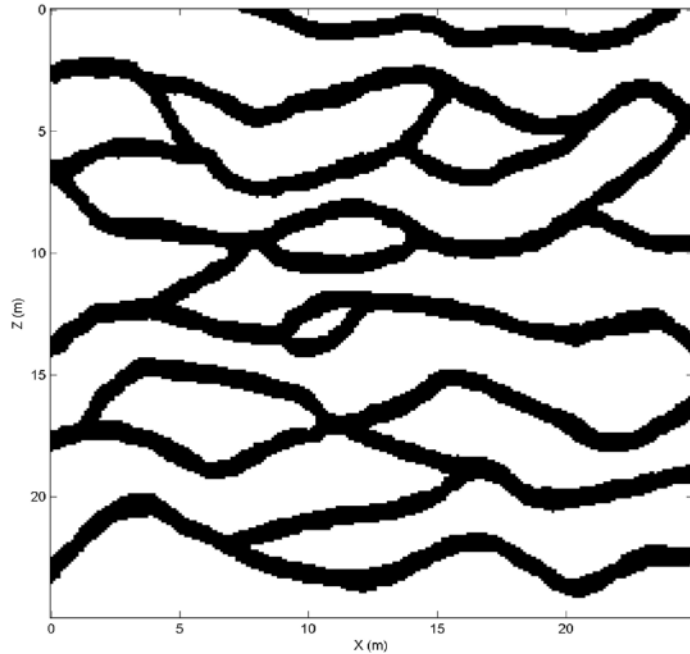
$$f(m \mid \text{prior}) \rightarrow m^*$$



$$f(m \mid \text{prior}, \text{data}) \rightarrow m^*$$



A simple application of geostatistics and inverse problem theory



$$P(\text{red circle} == \text{clay} \mid \text{prior}) = 0.30$$

$$P(\text{red circle} == \text{clay} \mid \text{prior, data}) = 0.56$$

$$P(\text{red circle} == \text{clay} \mid \text{prior, data}) = 0.17$$

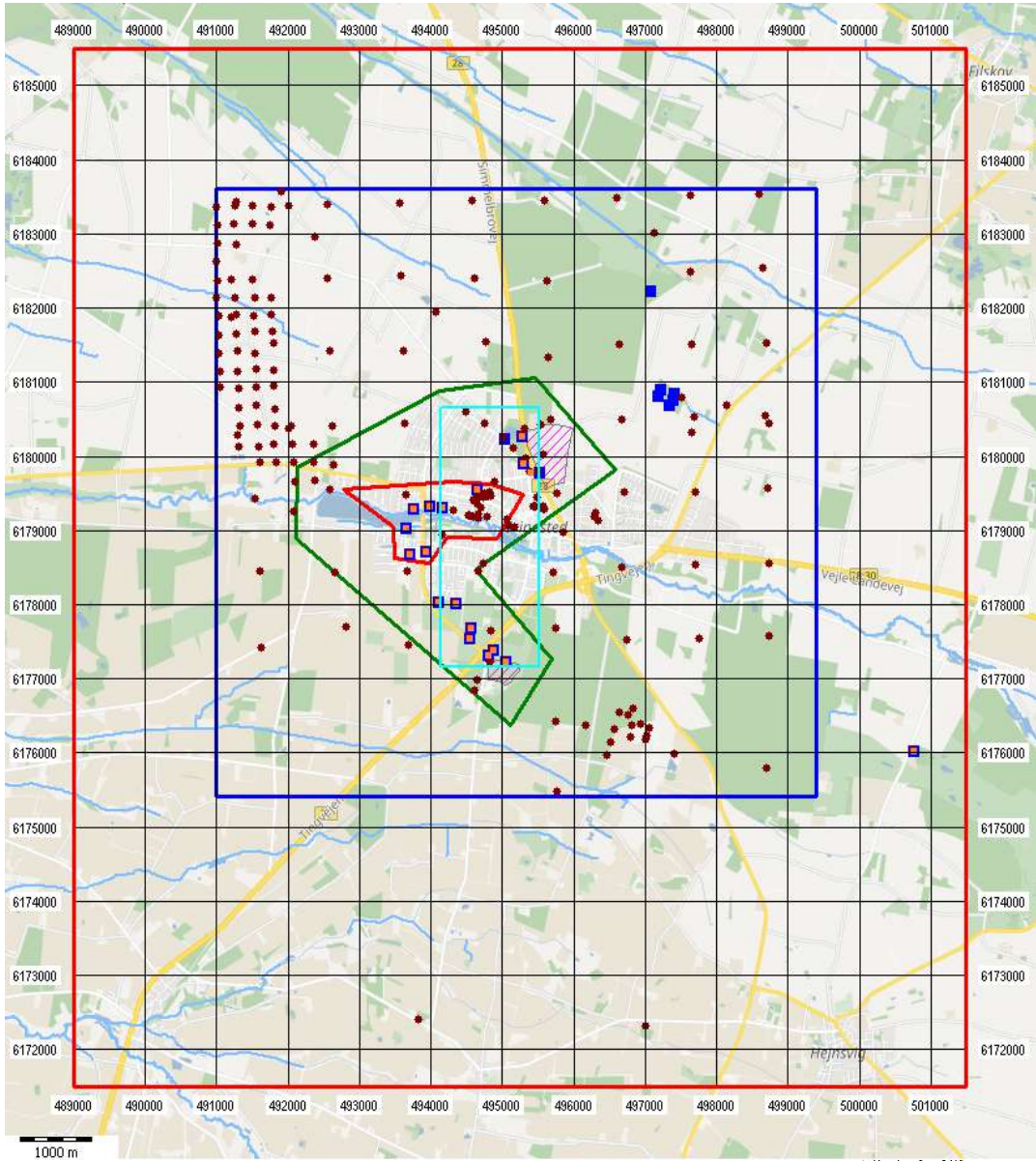
Kriging

MPS

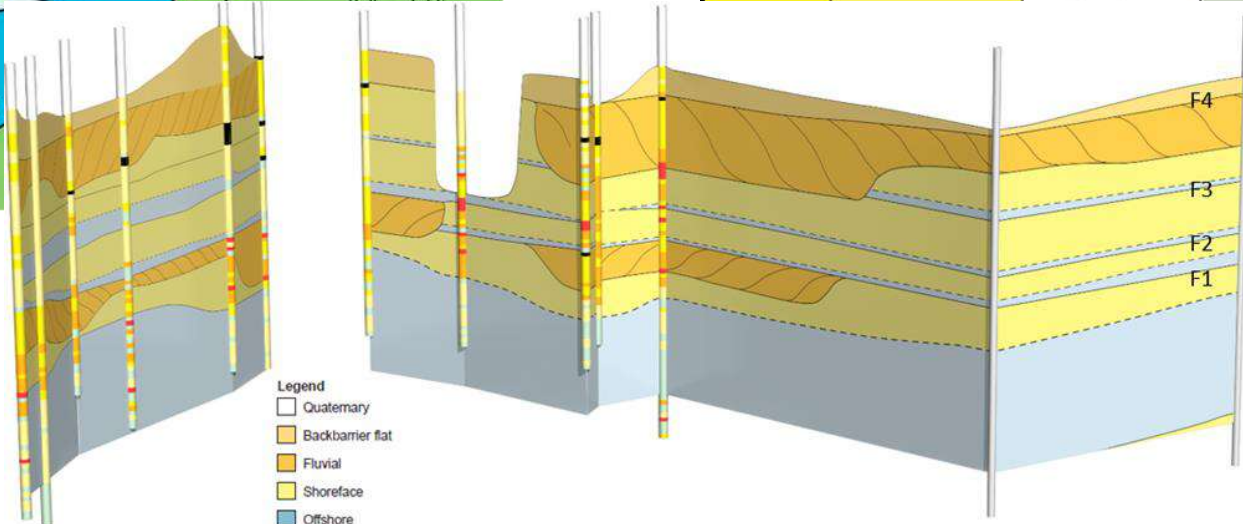
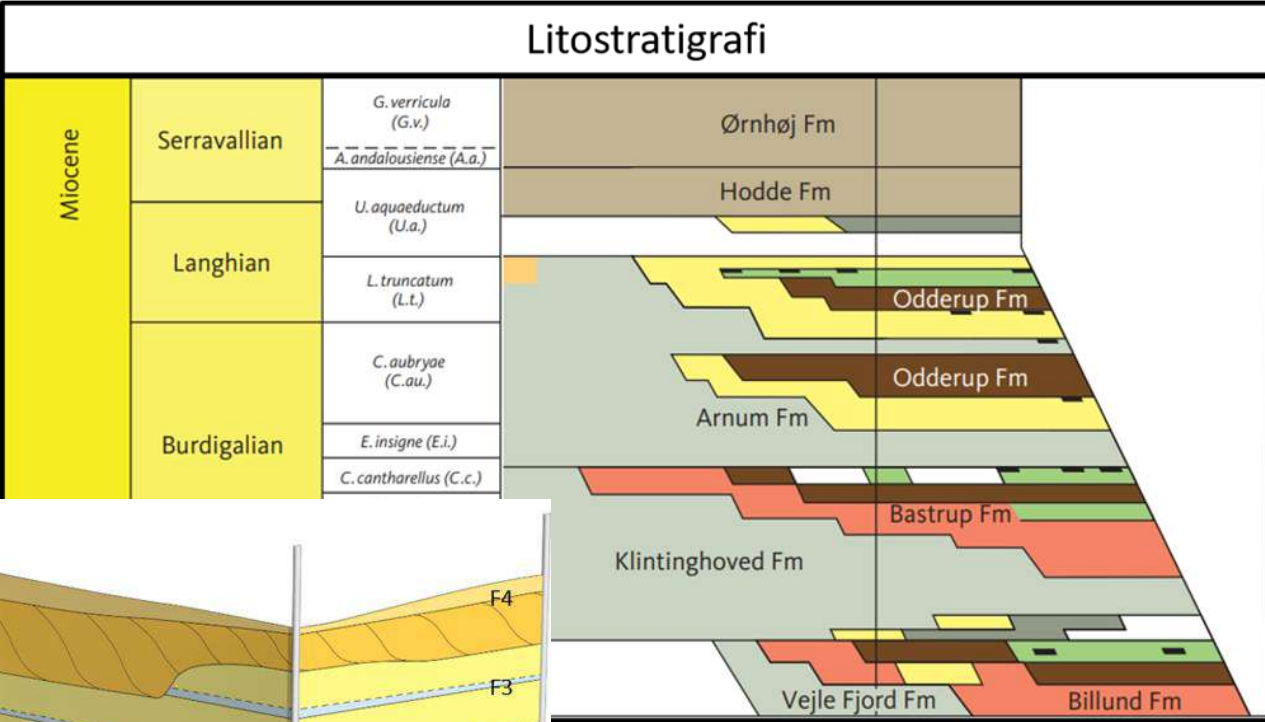
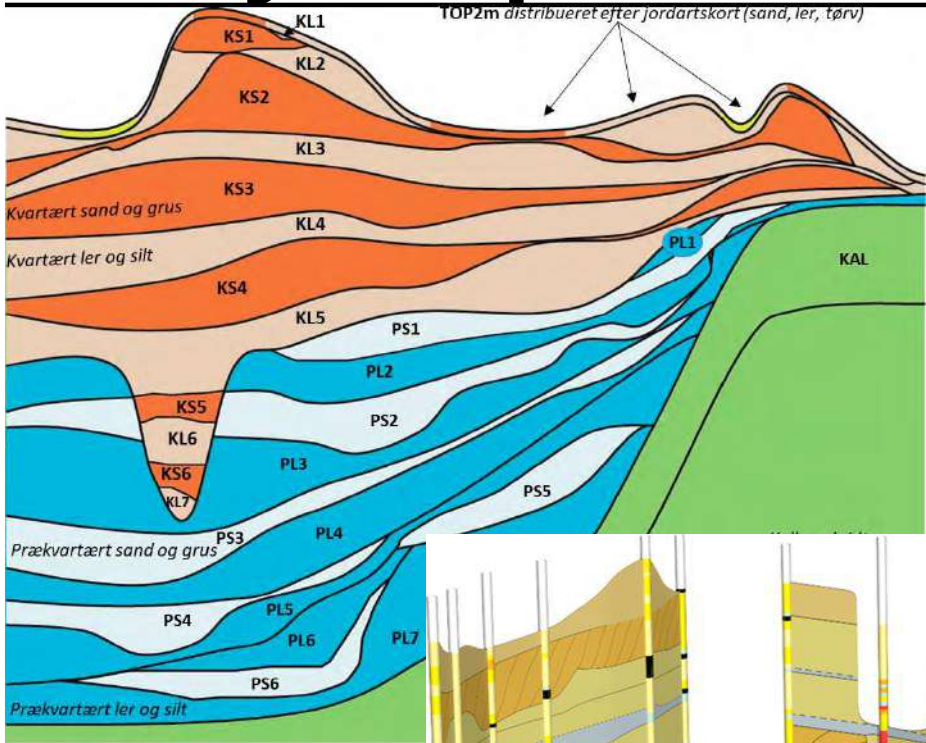
CASE: MAPING POLLUTION USING GEOLOGICALLY INFORMED

WITH, RASMUS BØDKER MADSEN, INGELISE BALLING, GEUS, REGION MIDT.

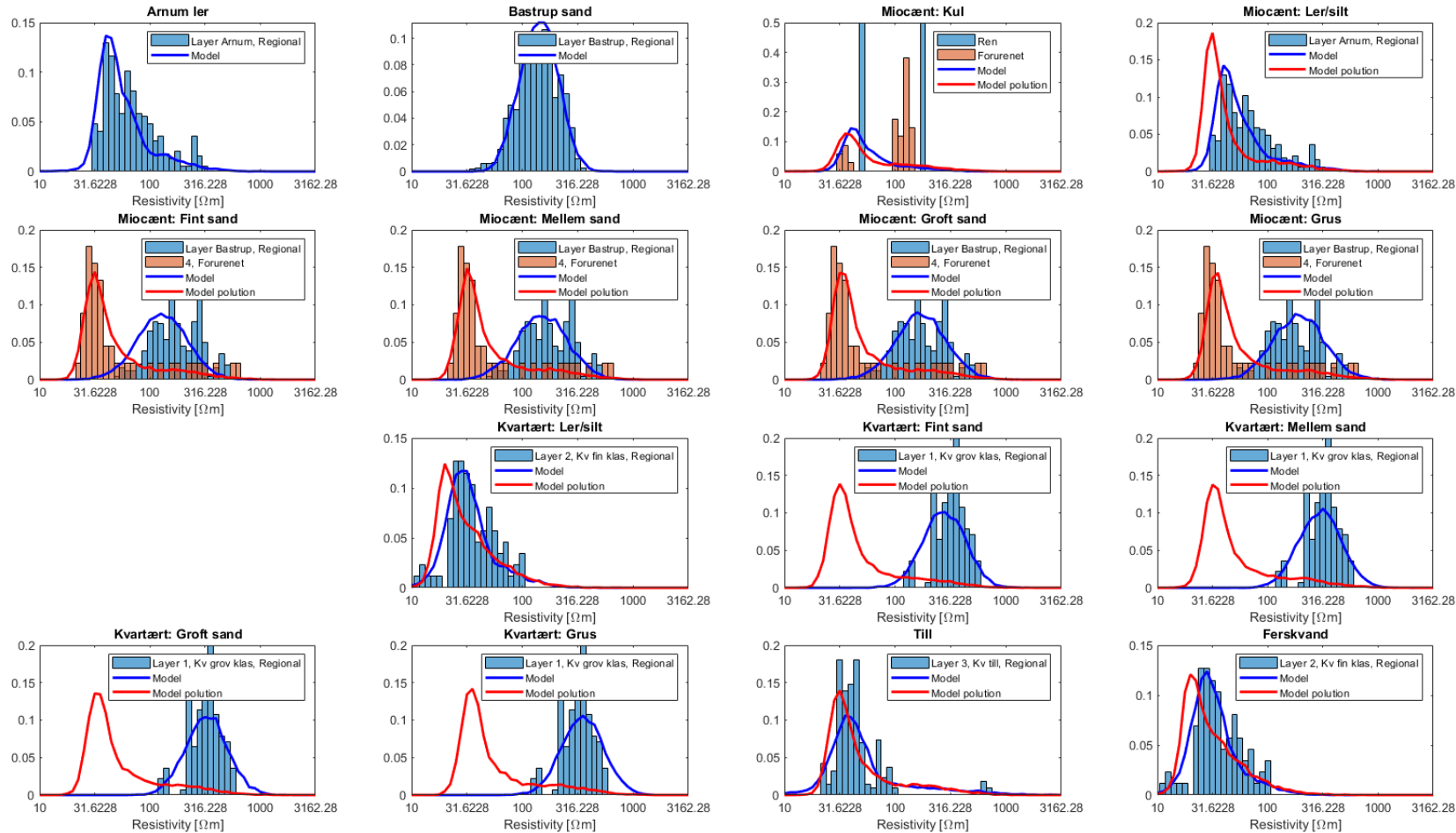
GRINDSTED



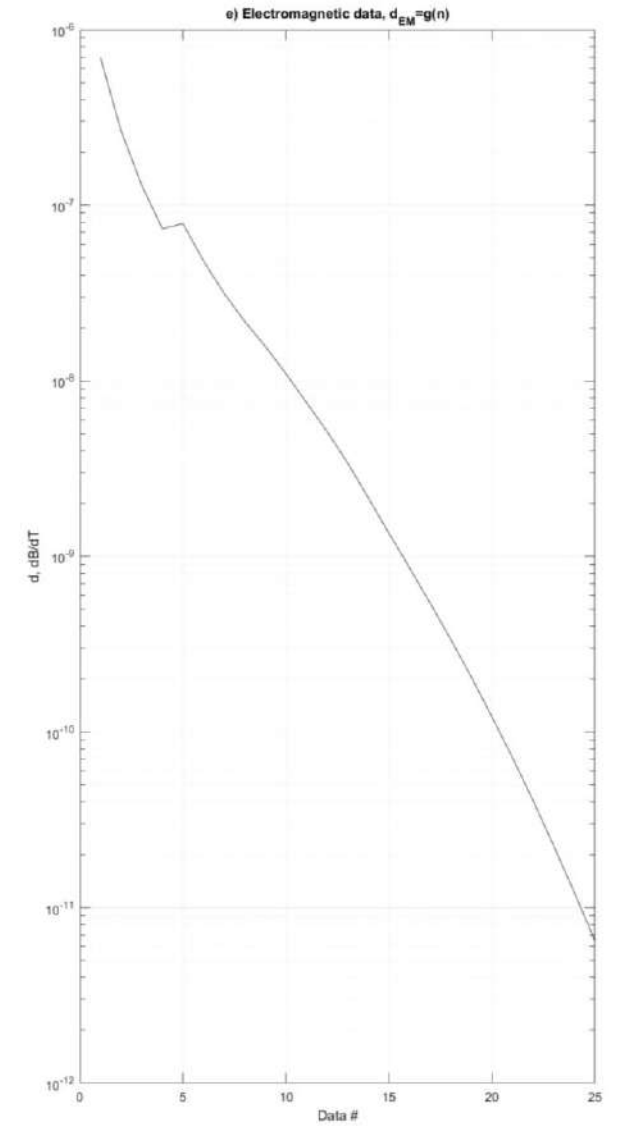
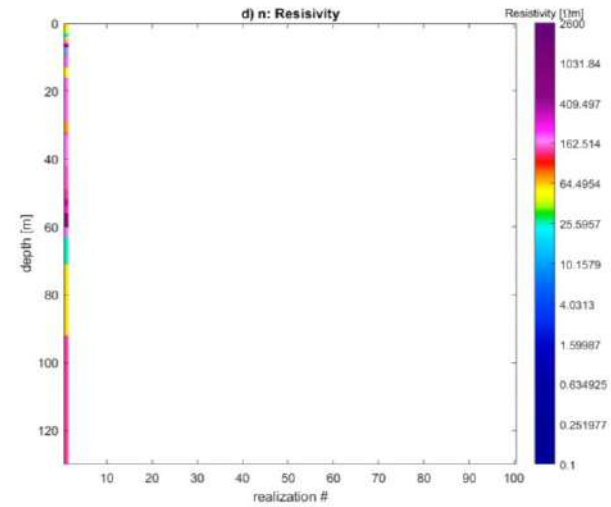
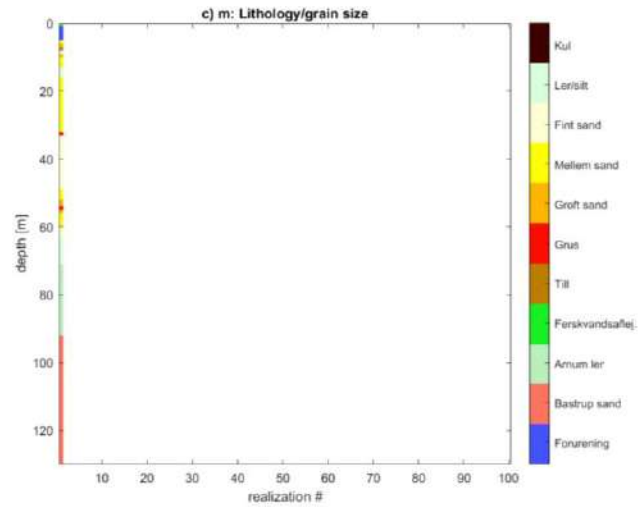
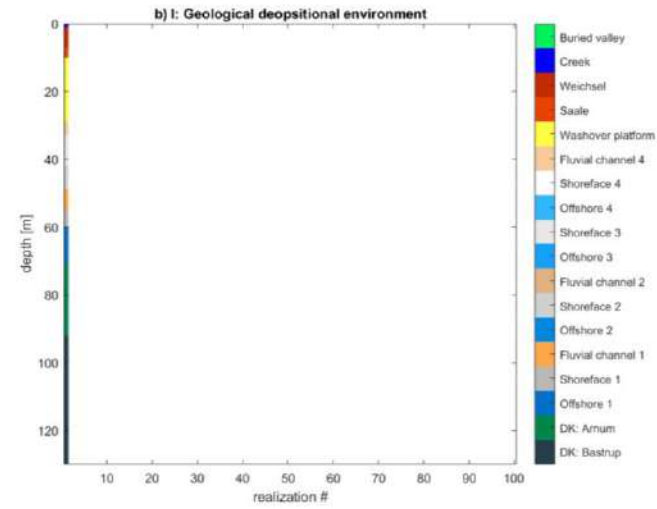
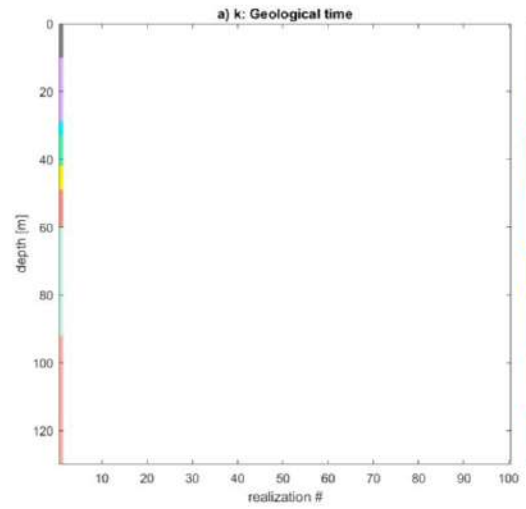
Geological prior information



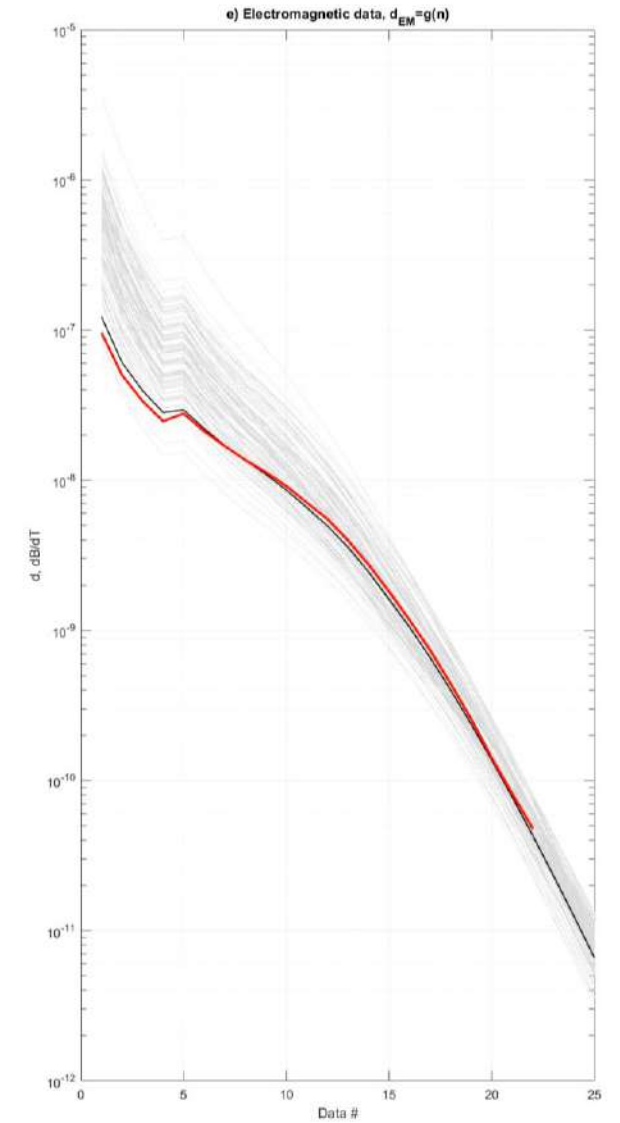
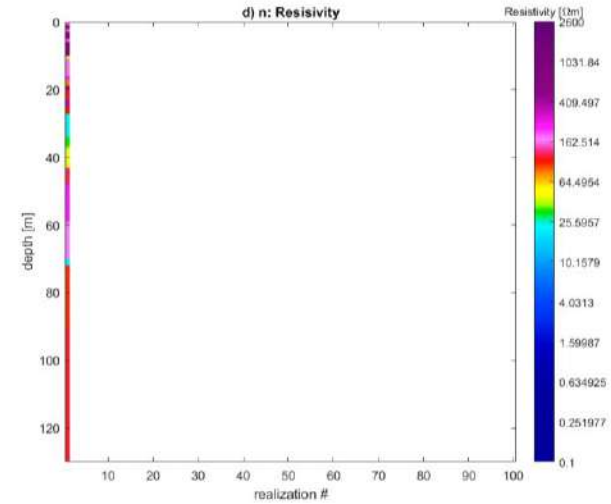
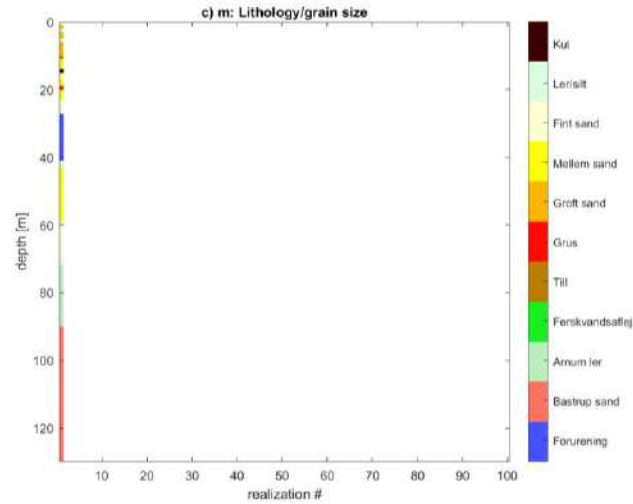
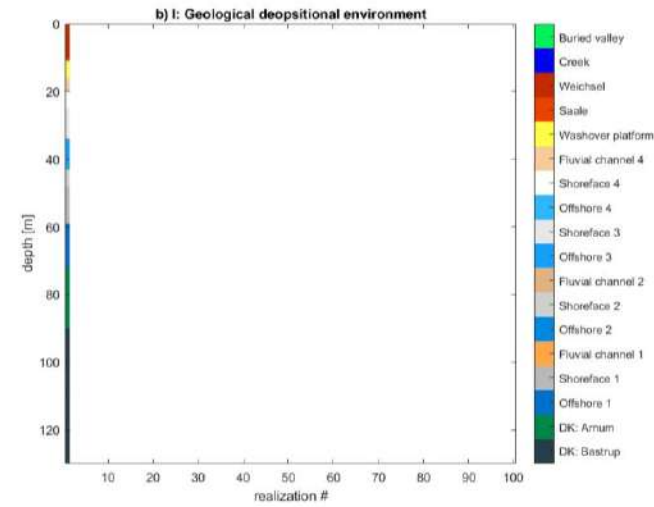
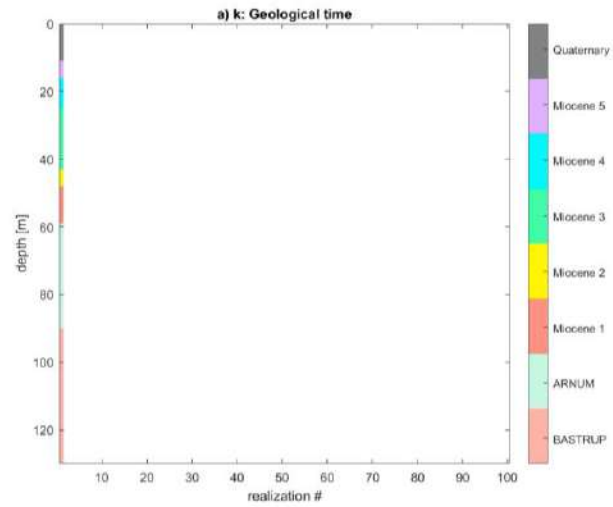
Grainsize -> resistivity modeling f(resistivity | grainsize)

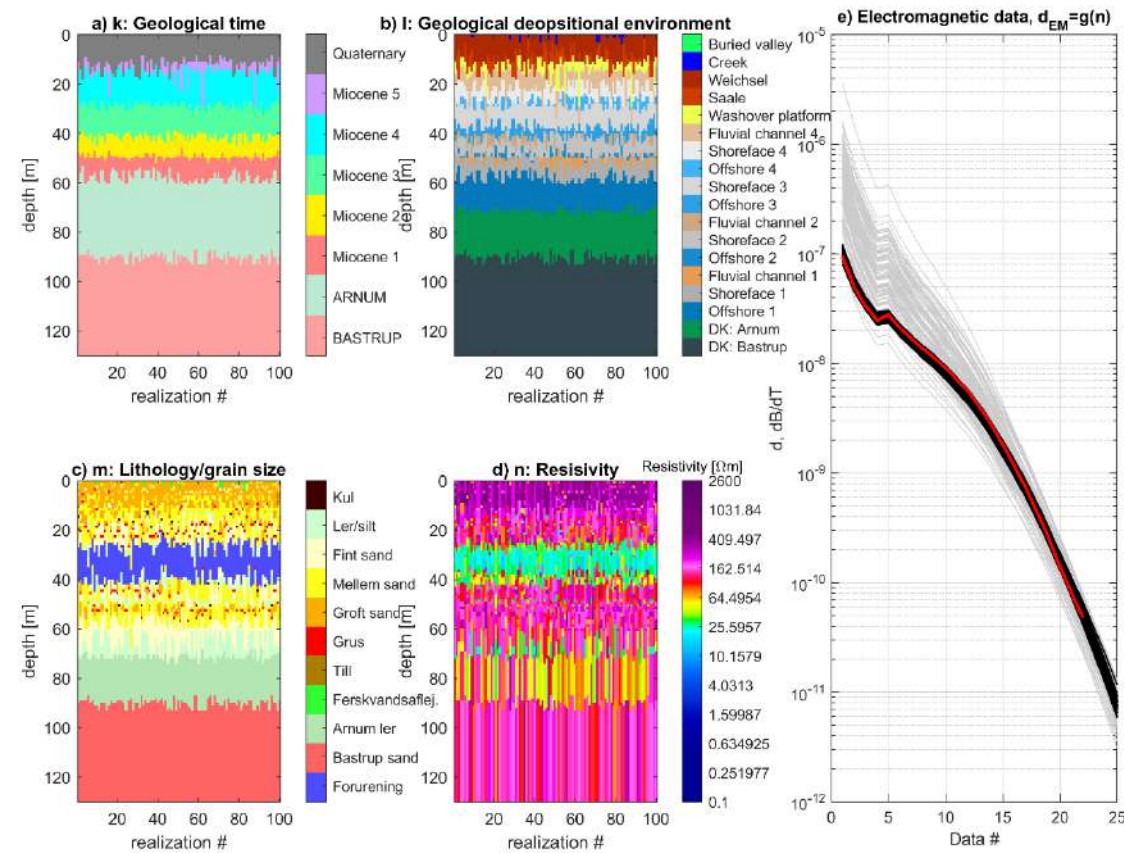
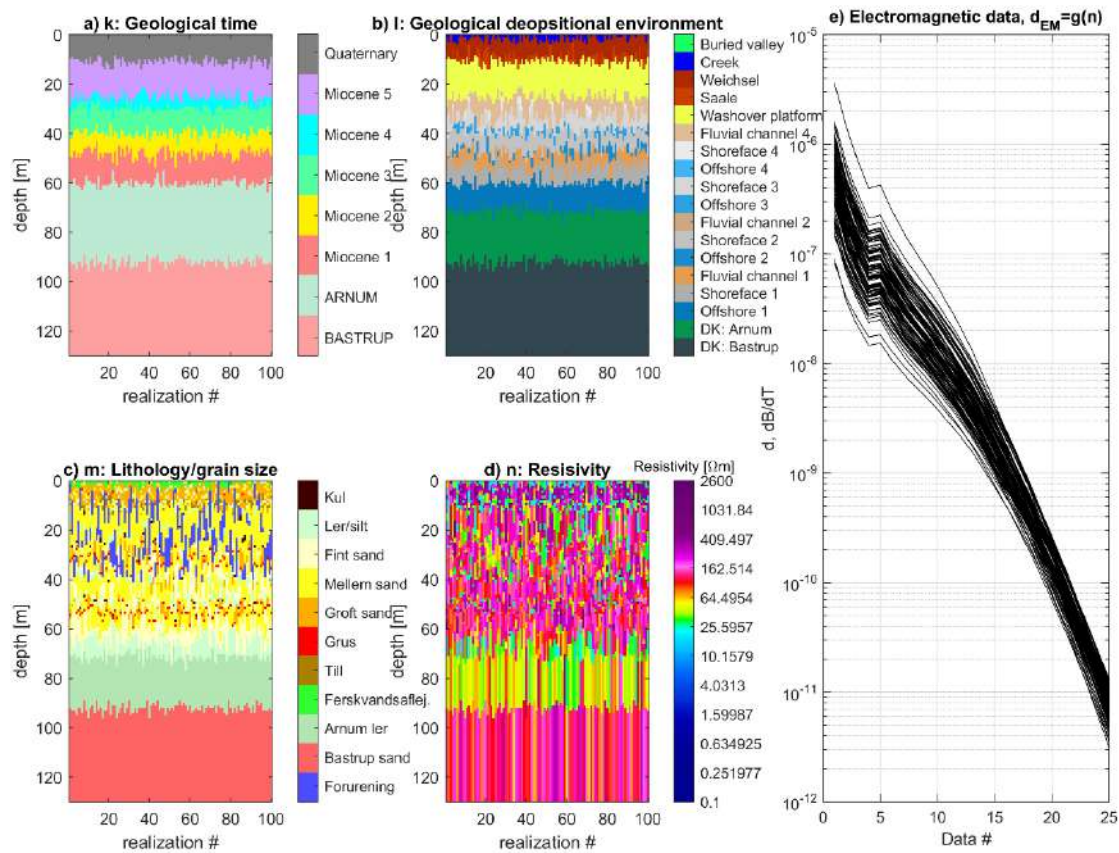


PRIOR

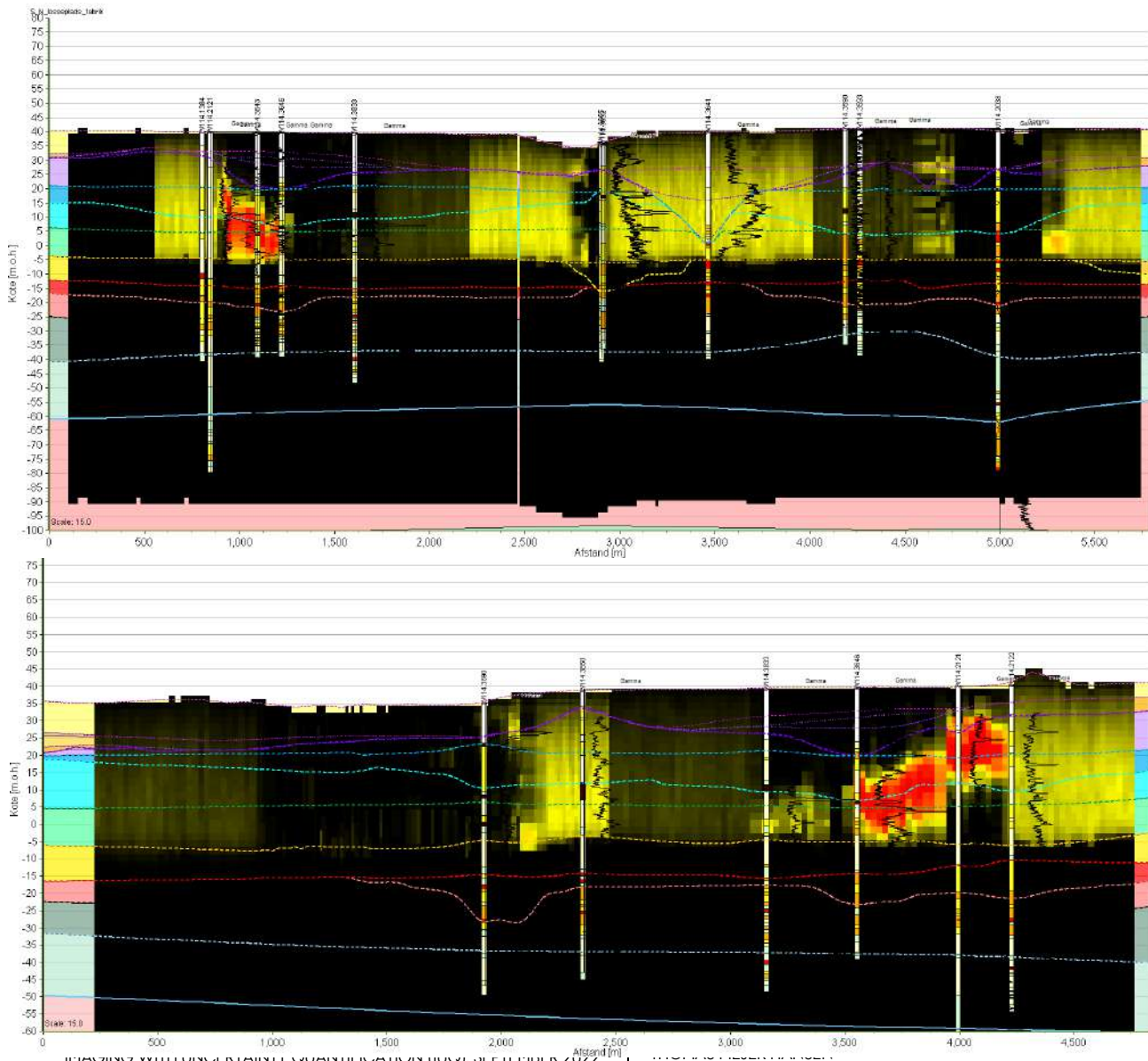


POST





Probability of contaminated plume

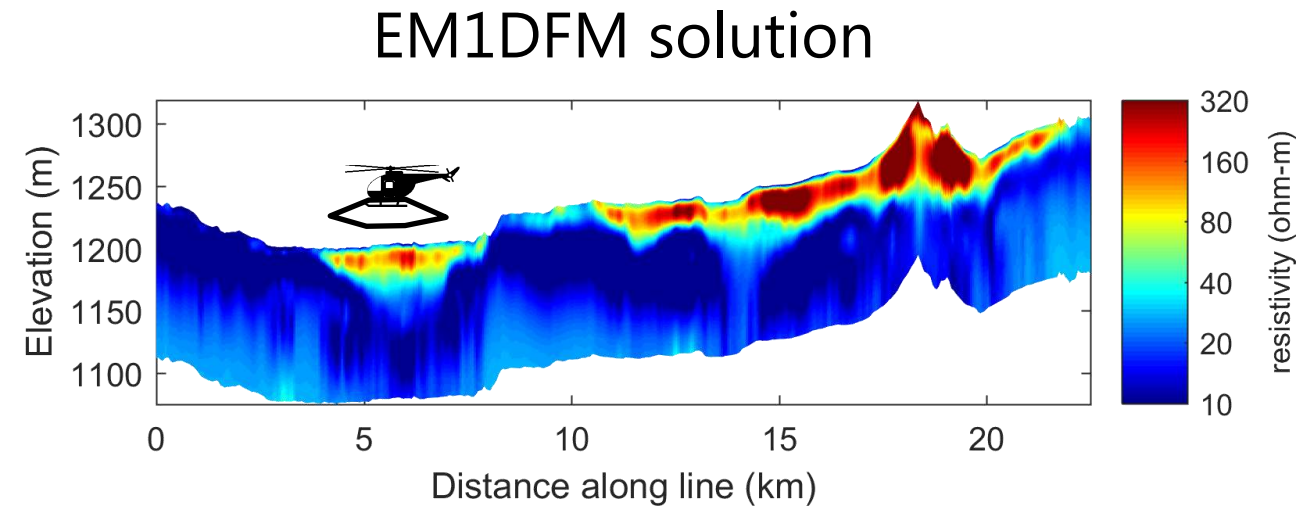
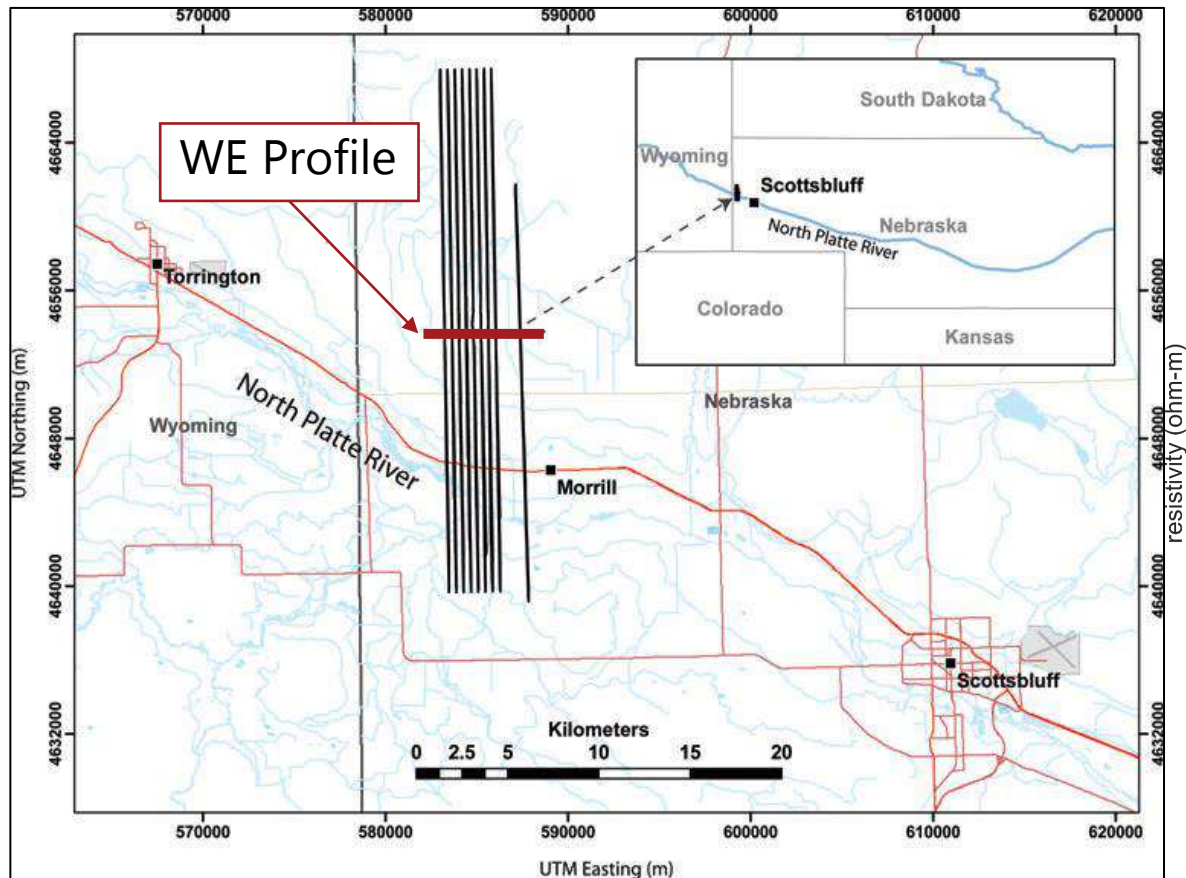


CASE: AIRBORNE EM

Sampling the posterior distribution using the extended Metropolis algorithm

THOMAS MEJER HANSEN AND BURKE MINSLEY

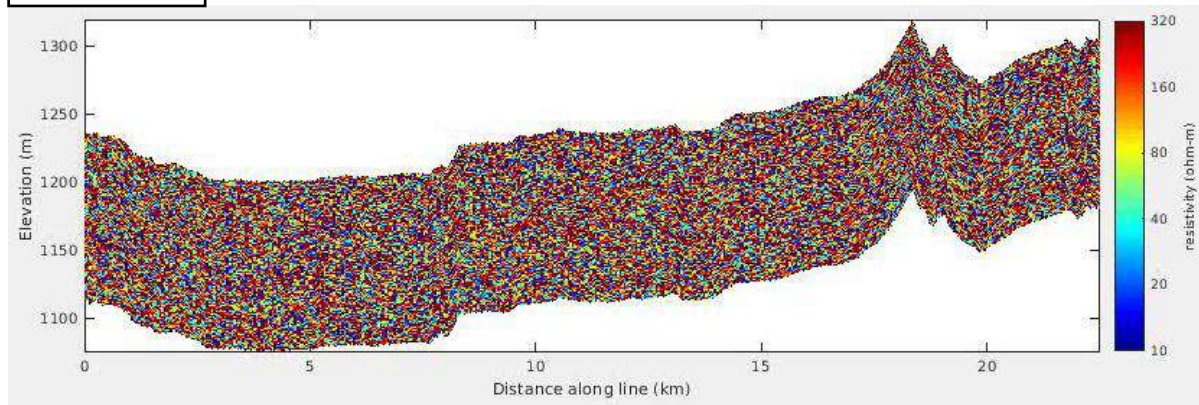
Airborne FD-EM data from Morill, Nebraska



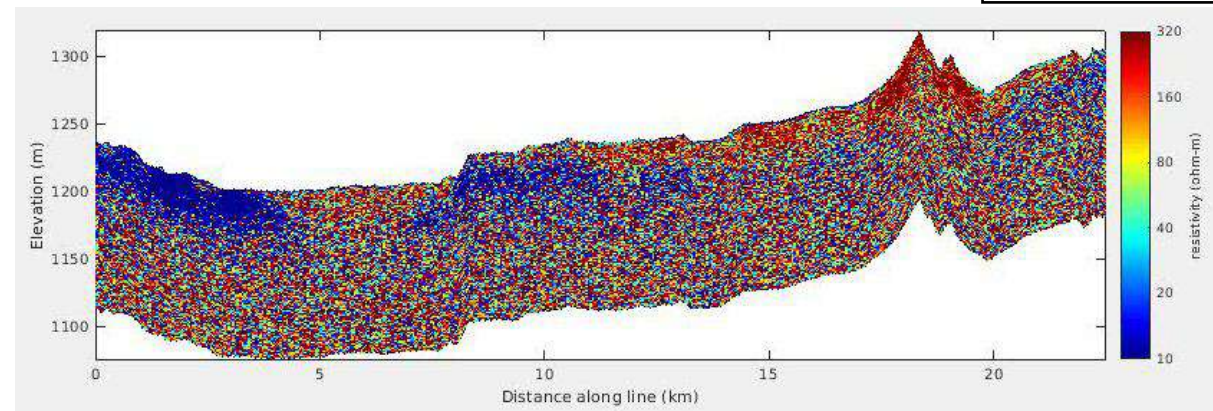
1D inversion along profile – ‘no prior’ $f_{\text{GEO}}(\mathbf{m}) = \text{constant}$

Purely data driven

Prior

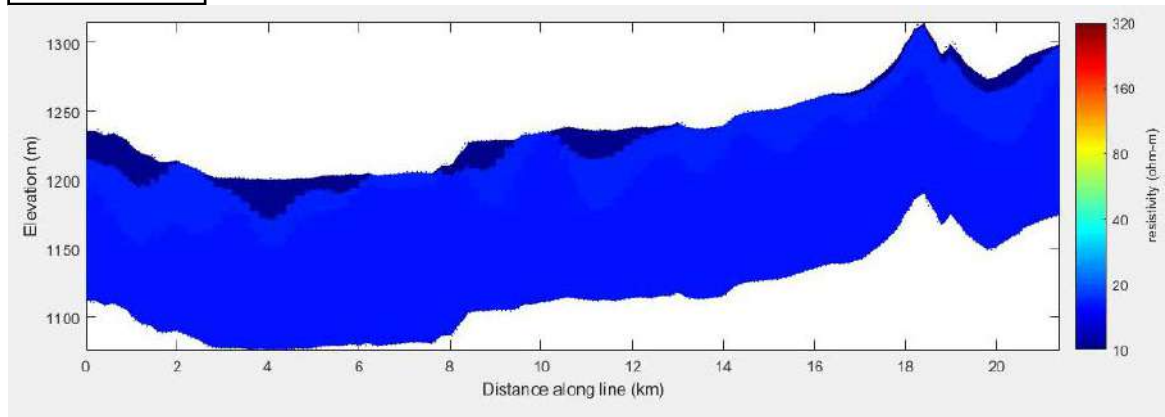


Posterior

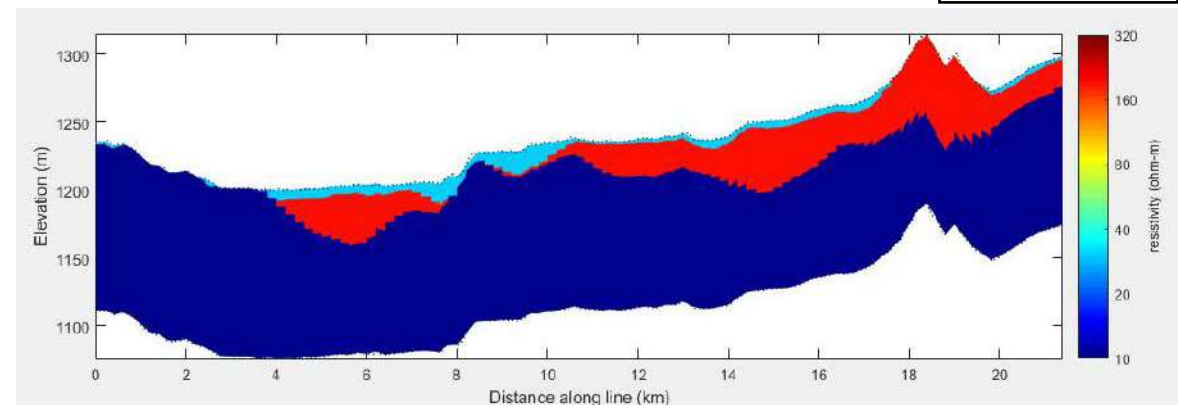


1D inversion along profile – $f_{\text{GEO}}(\mathbf{m})$: 3 layer. Too informative prior

Prior

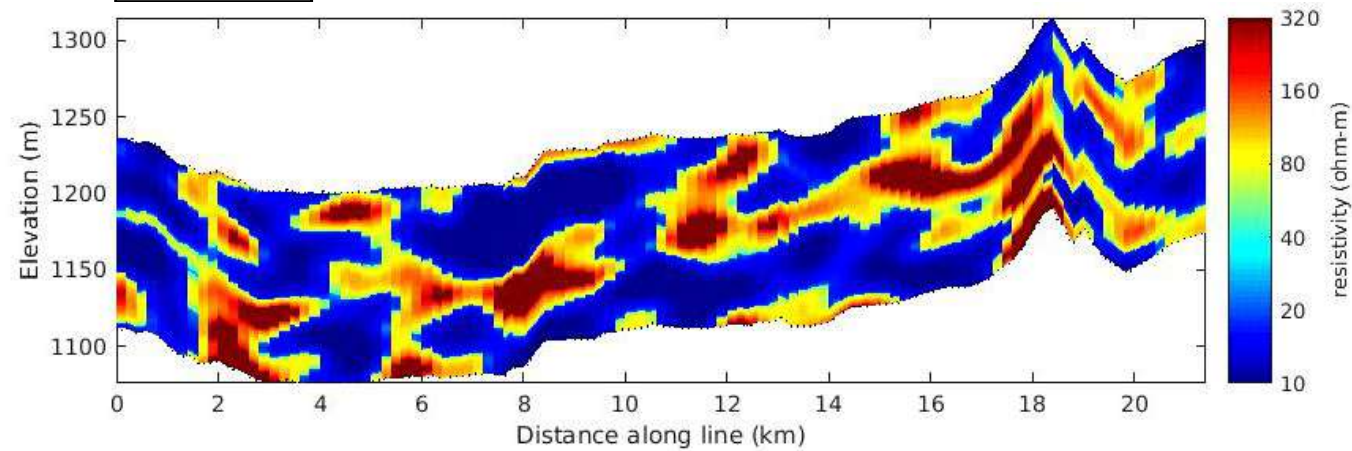


Posterior

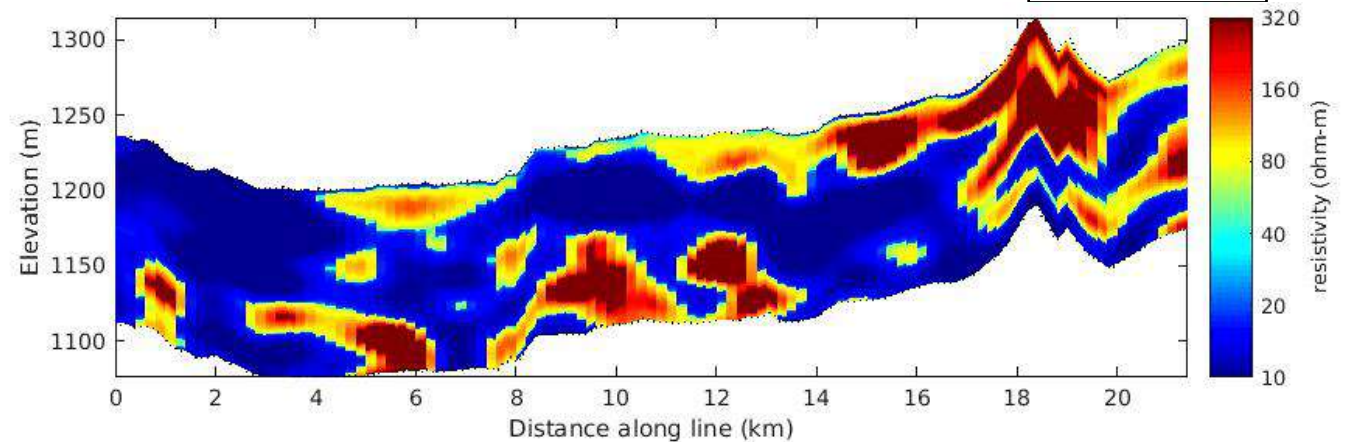


1D inversion along profile $-f_{\text{GEO}}(\mathbf{m})$: realistic Trimodal prior inferred from well logs

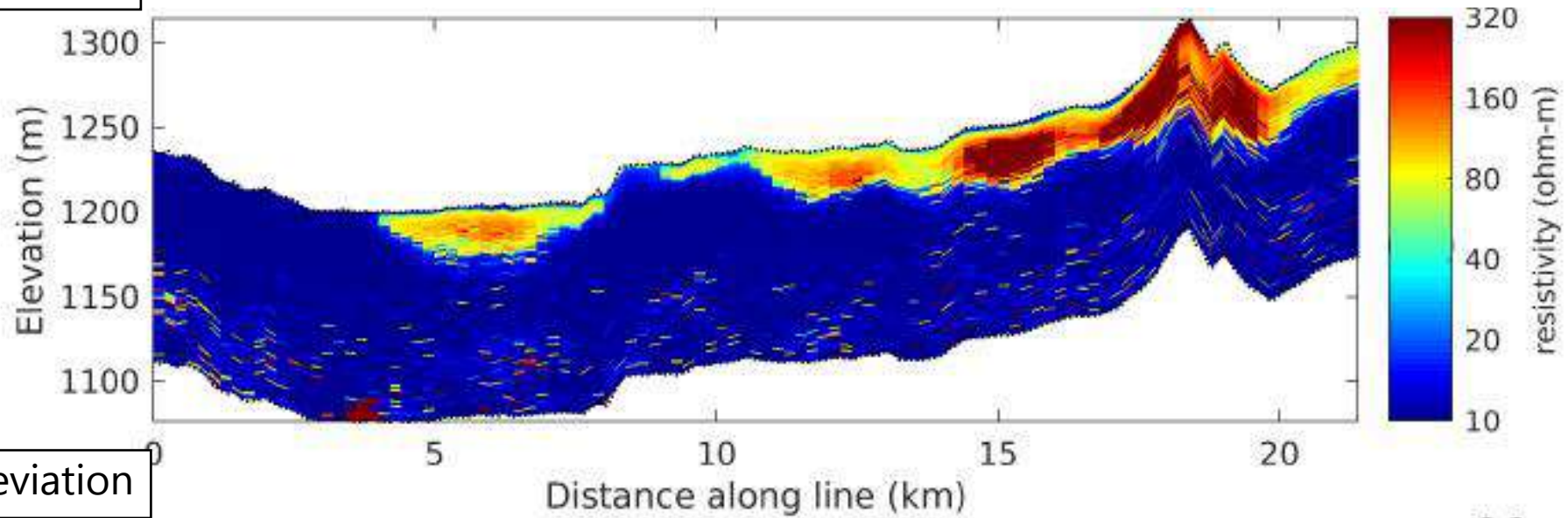
Prior



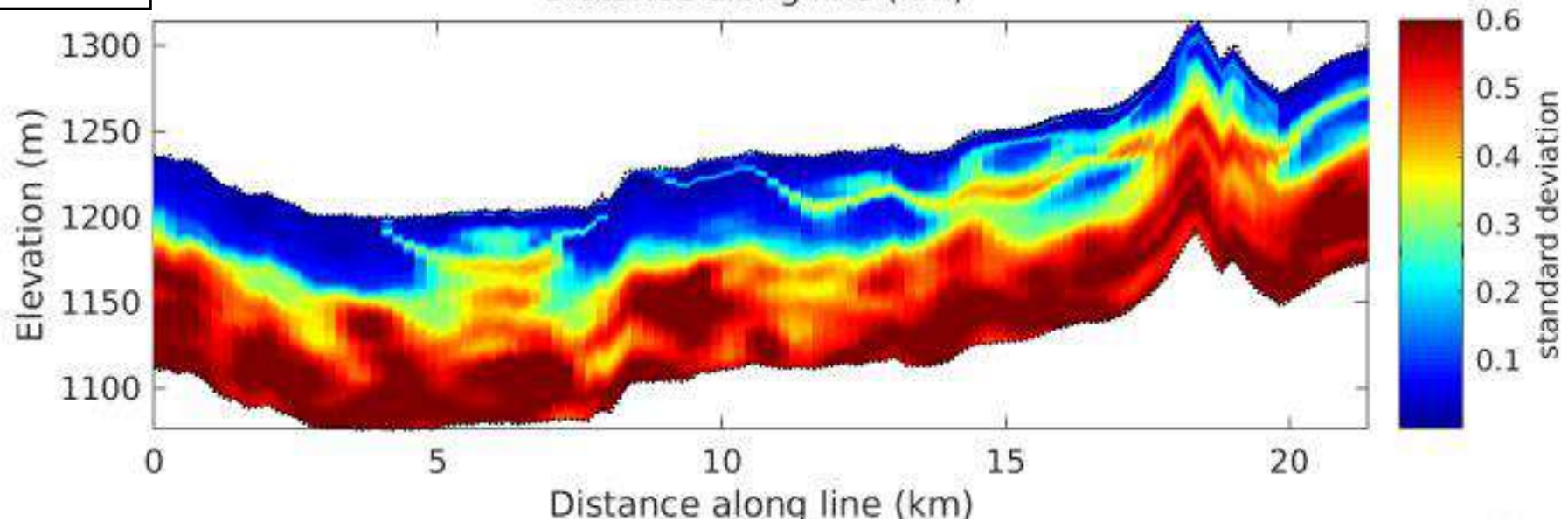
Posterior



Mode



Standard deviation



CASE: AIRBORNE EM

Sampling the posterior distribution using the
extended rejection sampler

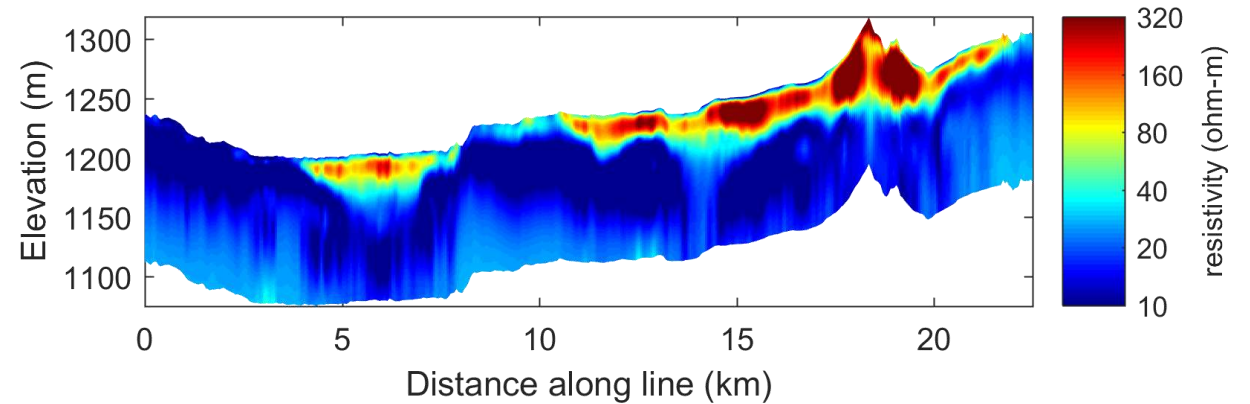
LOCALIZED PROBABILISTIC INVERSION

$[N^*, M^*, D^*]$

$\rho(m)$



$n^{*1}, m^{*1}, d^{*1},$
 $n^{*2}, m^{*2}, d^{*2},$
...
 $n^{*nr}, m^{*nr}, d^{*nr},$



Localized inversed problem:

When the prior and the forward problem are the same, for different data sets!

THE EXTENDED REJECTION SAMPLER WITH LOOKUP TABLES

The goal: Sample from $\sigma(\mathbf{m}) = k \rho(\mathbf{m}) L(\mathbf{m})$

1. Propose a model from $\rho(\mathbf{m}) \rightarrow \mathbf{m}^*$
2. Accept \mathbf{m}^* as a realization from the posterior with probability

$$P_{\text{acc}} = L(\mathbf{d}_{\text{obs}} - \mathbf{g}(\mathbf{m}^*)) / \max(L)$$

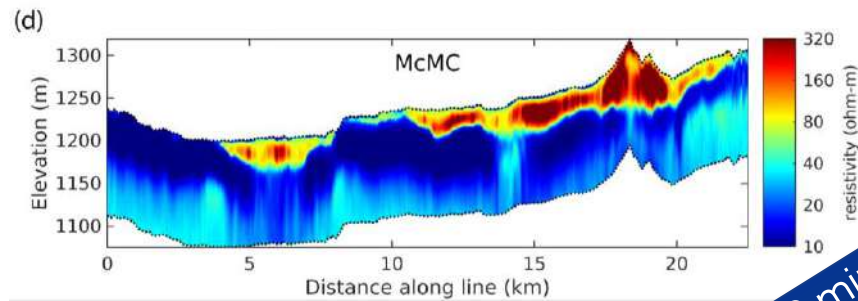
To apply the extended rejection sampler one must be able to

- 1) Sample the prior $\sigma(\mathbf{m})$
- 2) Evaluate the likelihood $L(\mathbf{m})$ for any model, (solve the forward problem, evaluate the noise)

Lookup table: $[\mathbf{M}^*, \mathbf{D}^*] \rightarrow$ Compute once, apply for any \mathbf{d}_{obs}

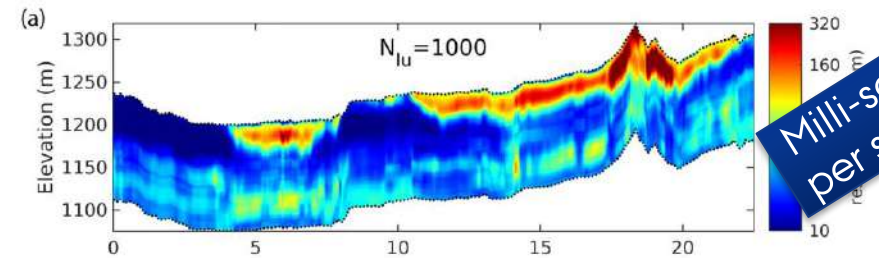
POSTERIOR MEAN

Extended Metropolis algorithm

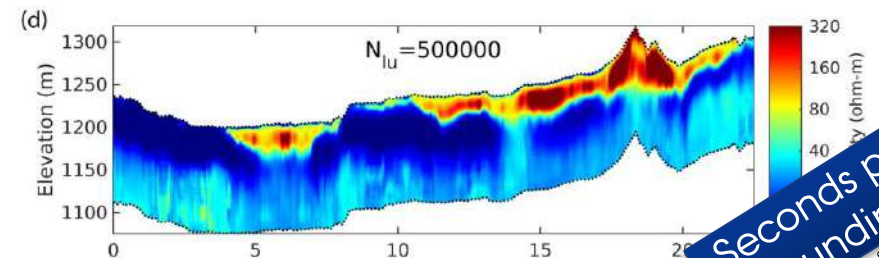
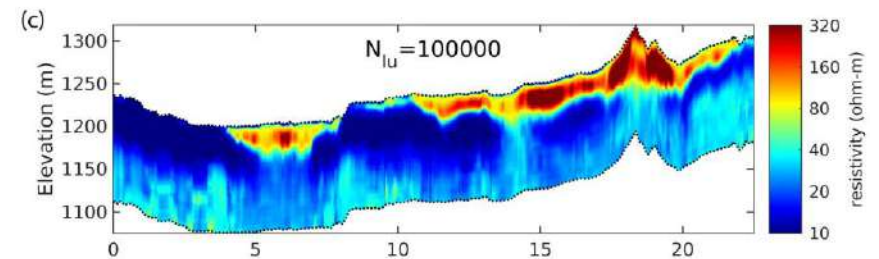
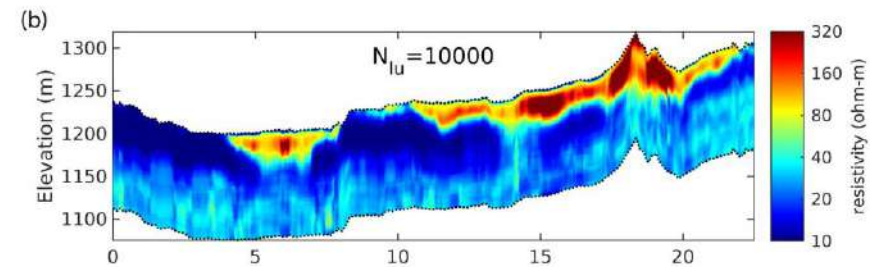


10 minutes
per sounding

Extended rejection sampler
with lookup tables



Milli-seconds
per sounding



Seconds
per sounding

CASE: AIRBORNE EM

Computing statistics of the posterior
distribution directly using machine learning,
without sampling the posterior!

THOMAS MEJER HANSEN AND CHRIS FINLAY

$$\sigma(\mathbf{m}) = k \rho(\mathbf{m}) L(\mathbf{d}_{\text{obs}} - \mathbf{g}(\mathbf{m}))$$

The forward problem

$$\mathbf{d} = \mathbf{g}(\mathbf{m})$$

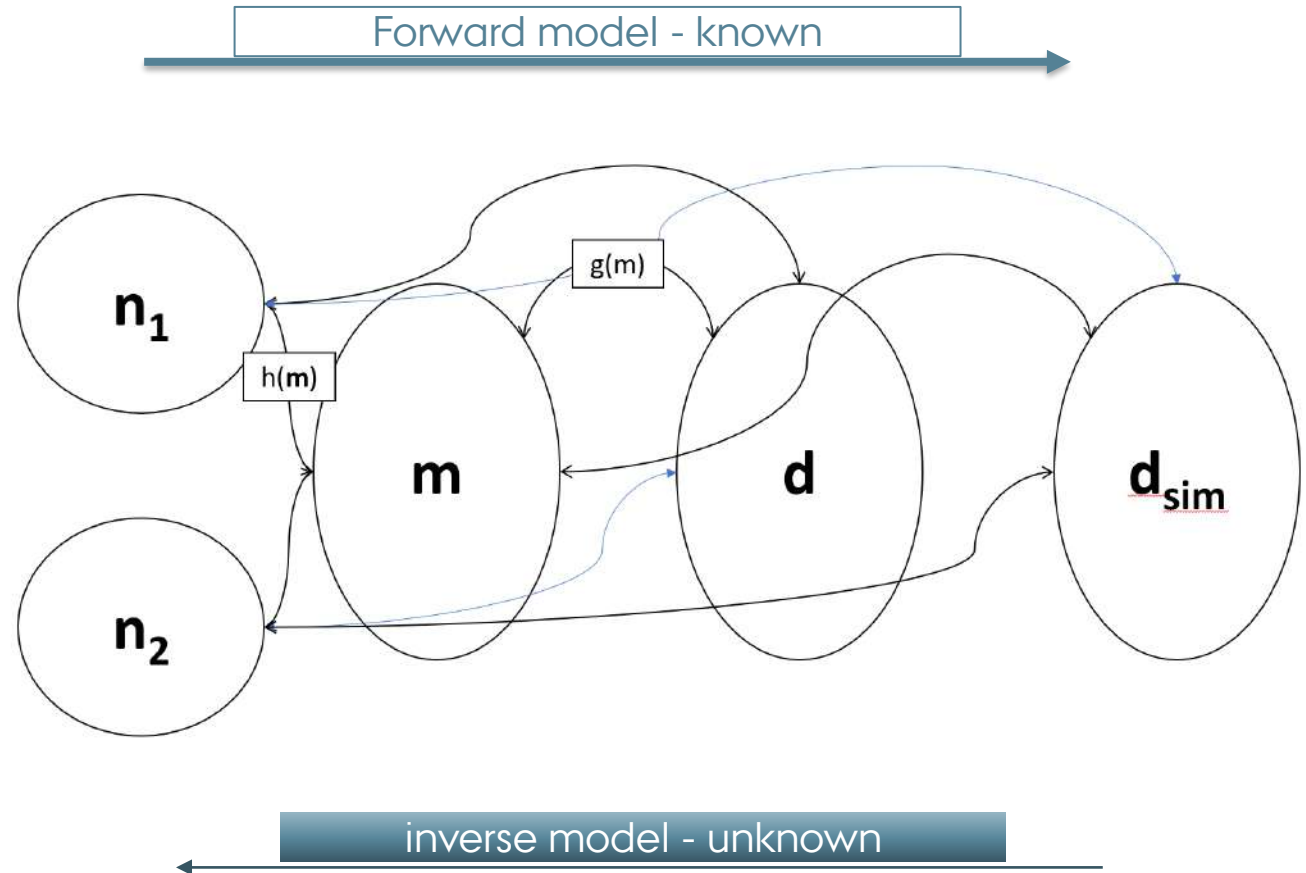
The forward problem -> simulation

$$\mathbf{d}_{\text{sim}} = \mathbf{d} + \mathbf{e}(\mathbf{m}) = \mathbf{g}(\mathbf{m}) + \mathbf{e}(\mathbf{m})$$

Sometimes one is not interested in \mathbf{m} ,
but in a property \mathbf{n} related to \mathbf{m}

$$\mathbf{n} = \mathbf{h}(\mathbf{m})$$

$$\rightarrow \sigma(\mathbf{n})$$



PROBABILISTIC INVERSION USING MACHINE LEARNING

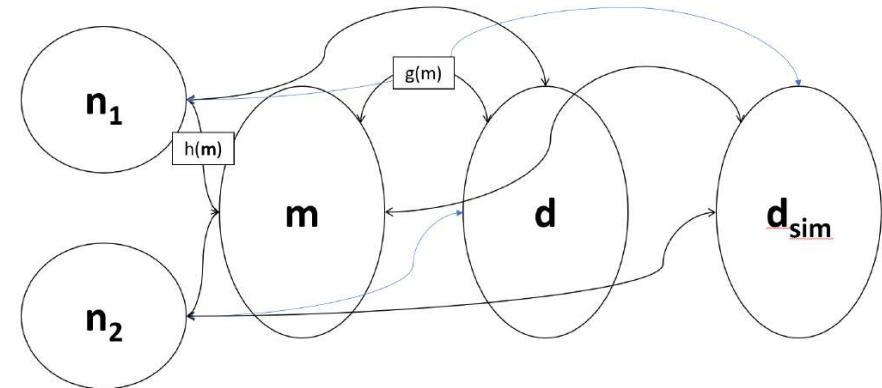
$$[N^*, M^*, D^*, D_{sim}^*]$$

$$d_{sim} = d + e(m) = g(m) + e(m)$$

$$n = h(m)$$

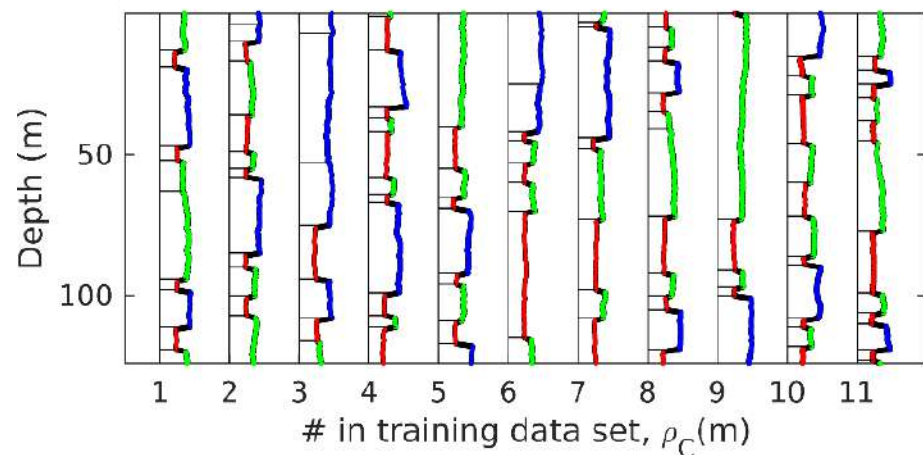
$$\begin{bmatrix} n^{*1}, m^{*1}, d^{*1}, d_{sim}^{*1} \\ n^{*2}, m^{*2}, d^{*2}, d_{sim}^{*2} \\ \dots \\ n^{*nr}, m^{*nr}, d^{*nr}, d_{sim}^{*nr} \end{bmatrix}$$

$\rho(m)$
↓

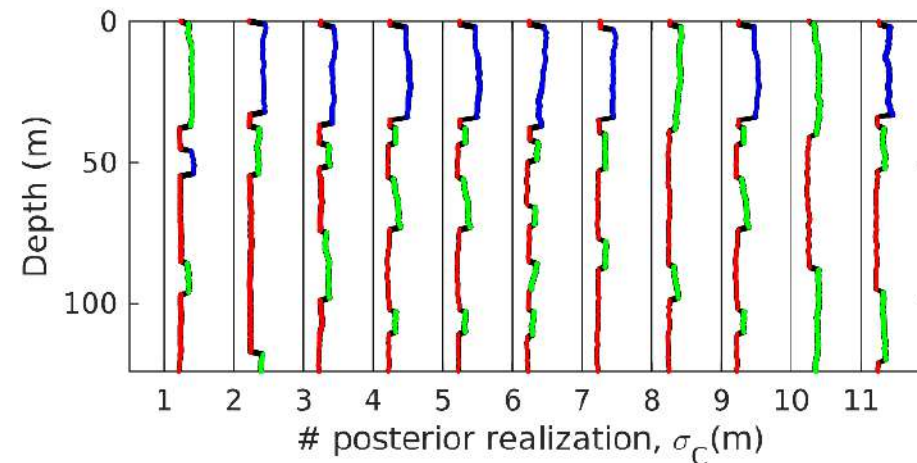


MODEL

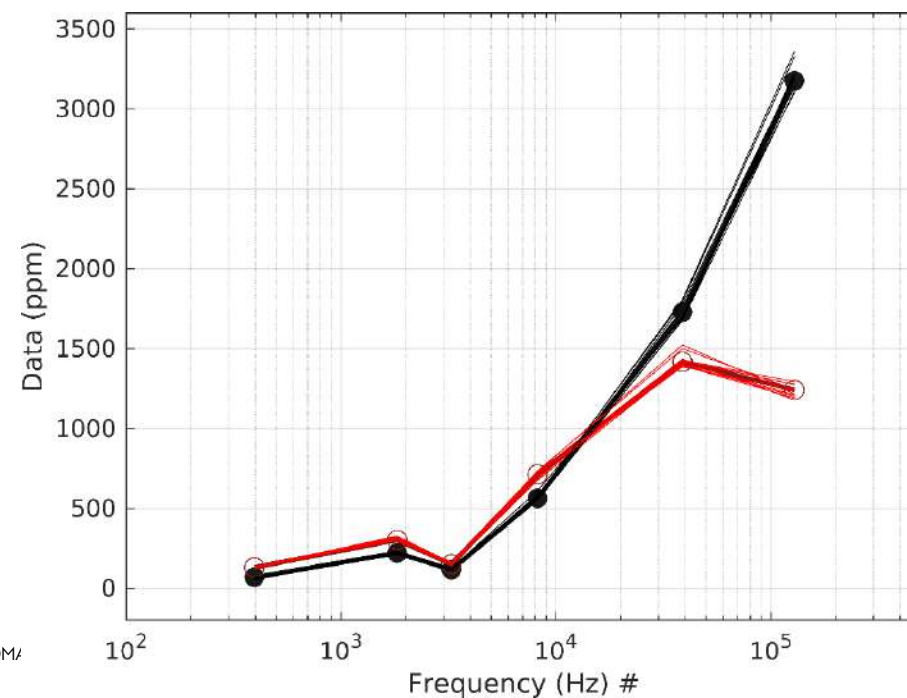
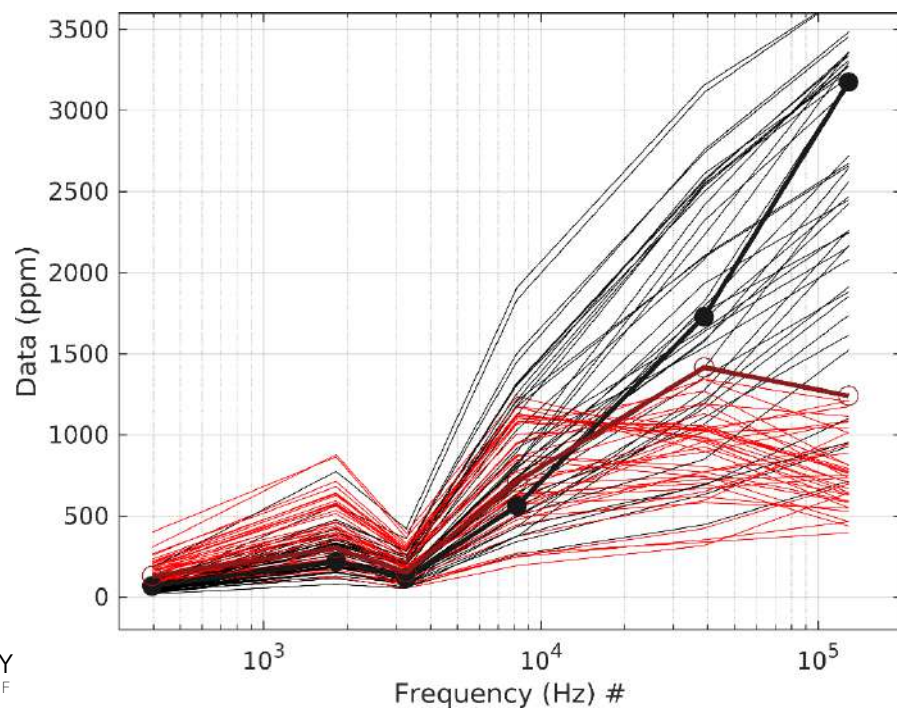
PRIOR $\rho(\mathbf{m})$



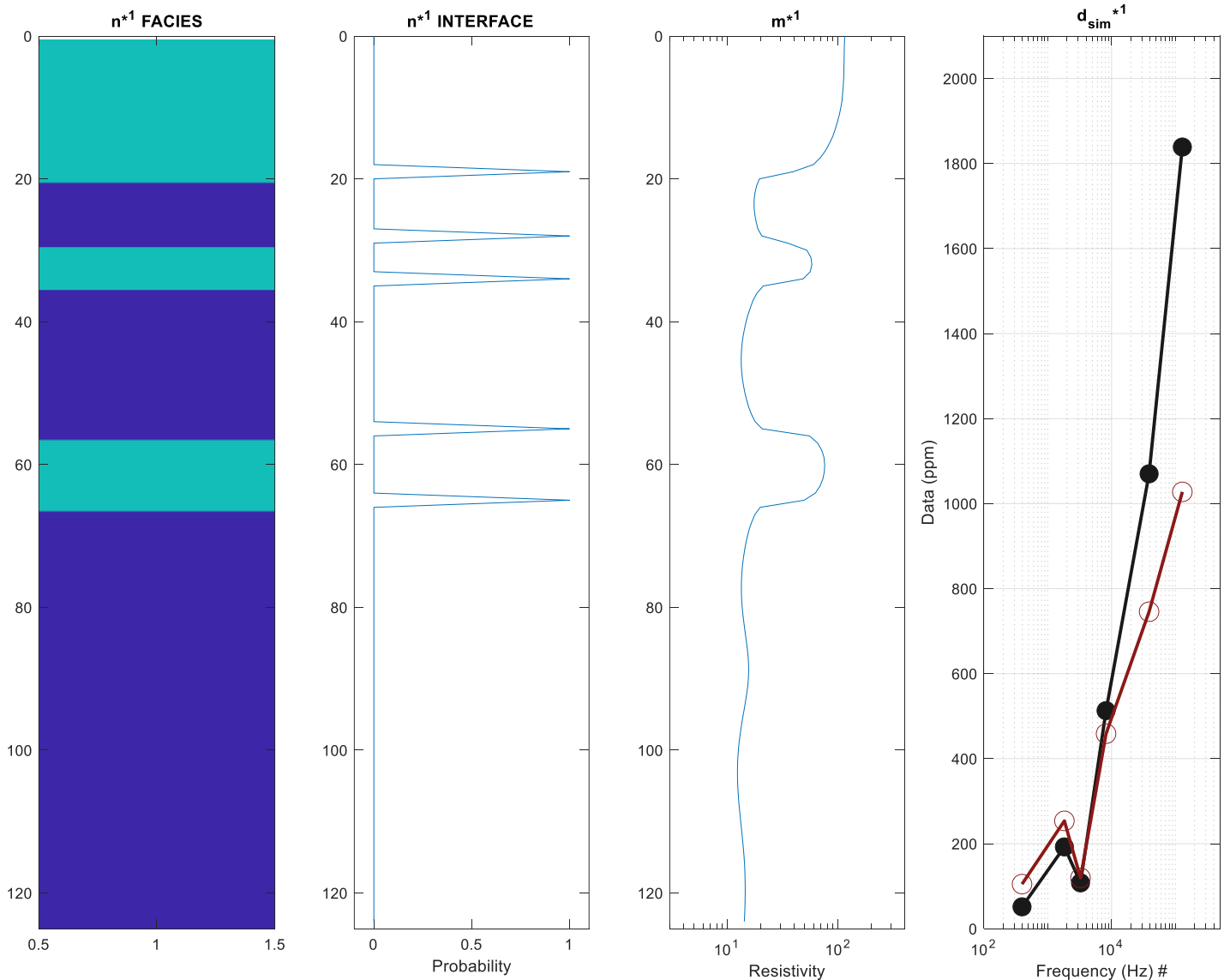
POSTERIOR $\sigma(\mathbf{m})$



DATA



CONSTRUCTING A TRAINING DATA SET



CONSTRUCTING A TRAINING DA

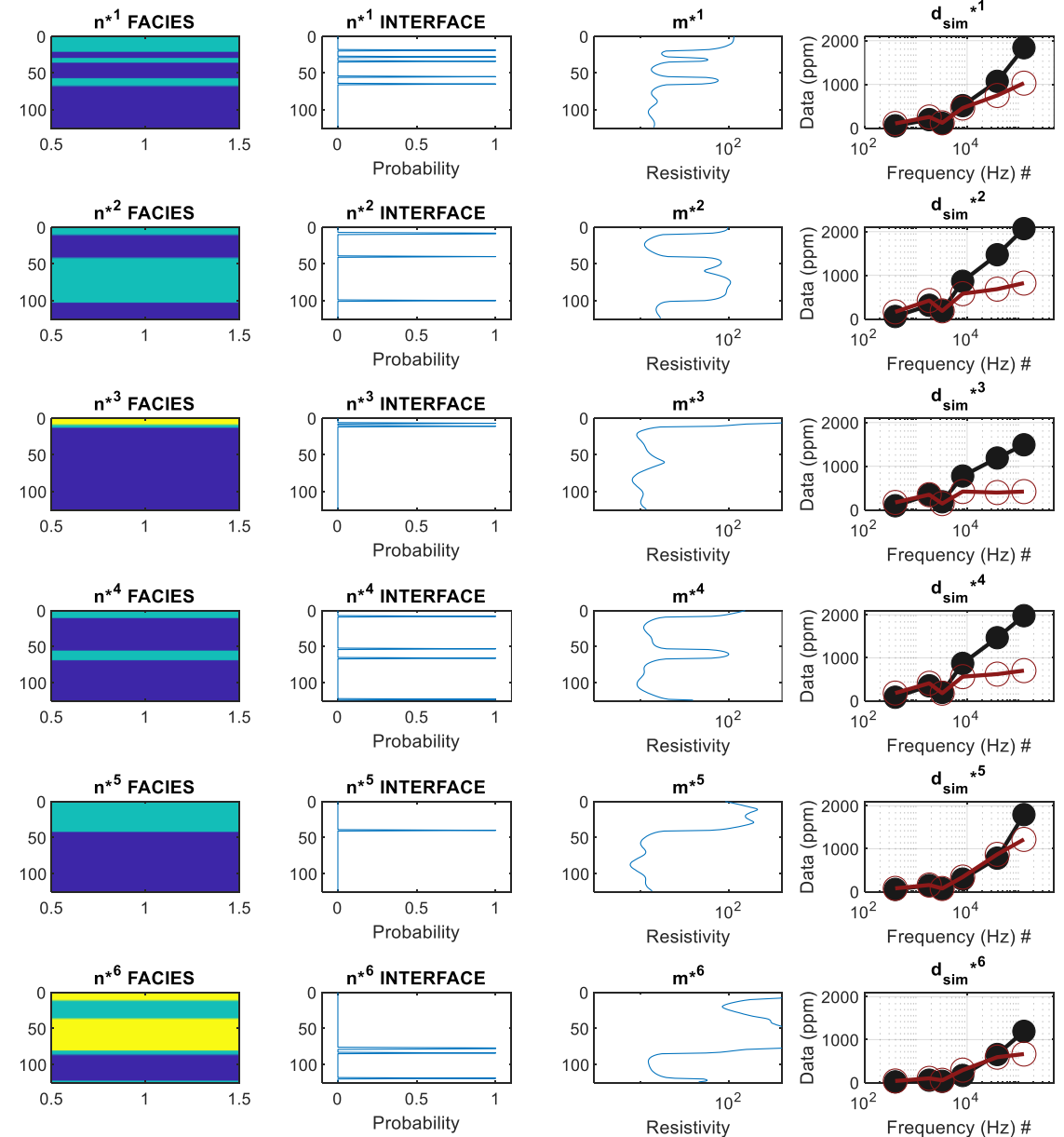
The training data can, in principle, be arbitrarily large

Here the training data set consists of up to 5000000 sets of \mathbf{n} , \mathbf{m} and \mathbf{d}_{sim}

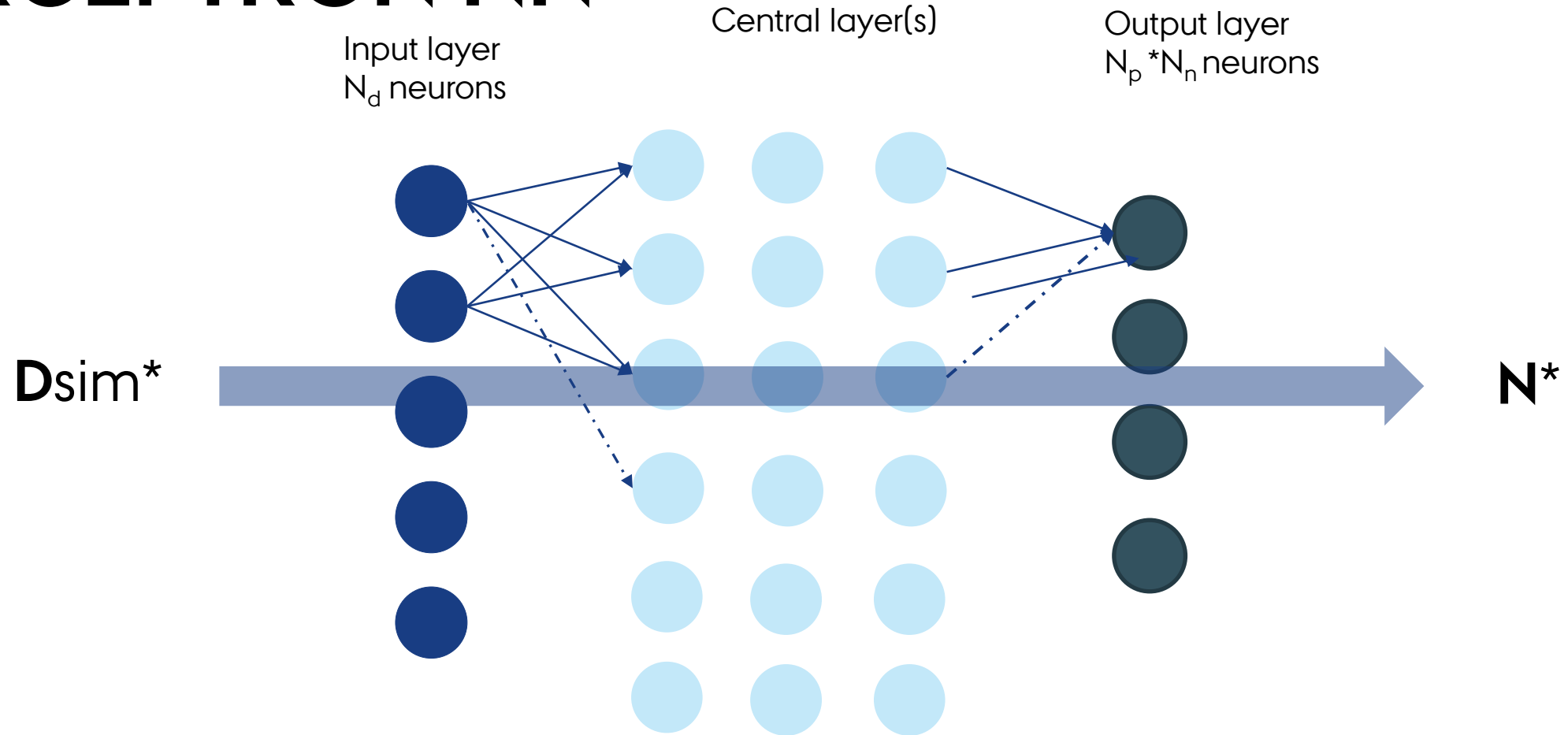
$[\mathbf{N}^*, \mathbf{M}^*, \mathbf{D}_{\text{sim}}^*]$

This training set express everything we know (as much as can be represented by a finite size sample) about

- The prior
- The forward
- The noise



DESIGNING A SIMPLE MULTILAYER PERCEPTRON NN



CHOOSING A LOSS FUNCTION TO MINIMIZE WHEN ADJUSTING THE FREE PARAMETERS OF THE NN

Regression, \mathbf{n} is a continuous parameter

Linear activation function in output layer

The output of the NN can for example be the parameters of a multivariate normal distribution $N(\mathbf{m}', \mathbf{Cm}')$:

$$f(\mathbf{m}^*) = k \exp(-0.5 (\mathbf{m}' - \mathbf{m}^*) \mathbf{Cm}'^{-1} (\mathbf{m}' - \mathbf{m}^*)^T)$$

$$\text{LOSS} = -\log(f(\mathbf{m}^*)) = -0.5 (\mathbf{m}' - \mathbf{m}^*) \mathbf{Cm}'^{-1} (\mathbf{m}' - \mathbf{m}^*)^T$$

A loss function minimizing this, and many more, distributions can easily be implemented using TensorFlow probability.

<https://www.tensorflow.org/probability>

Classification, \mathbf{n} is a discrete parameter

Softmax activation function in output layer

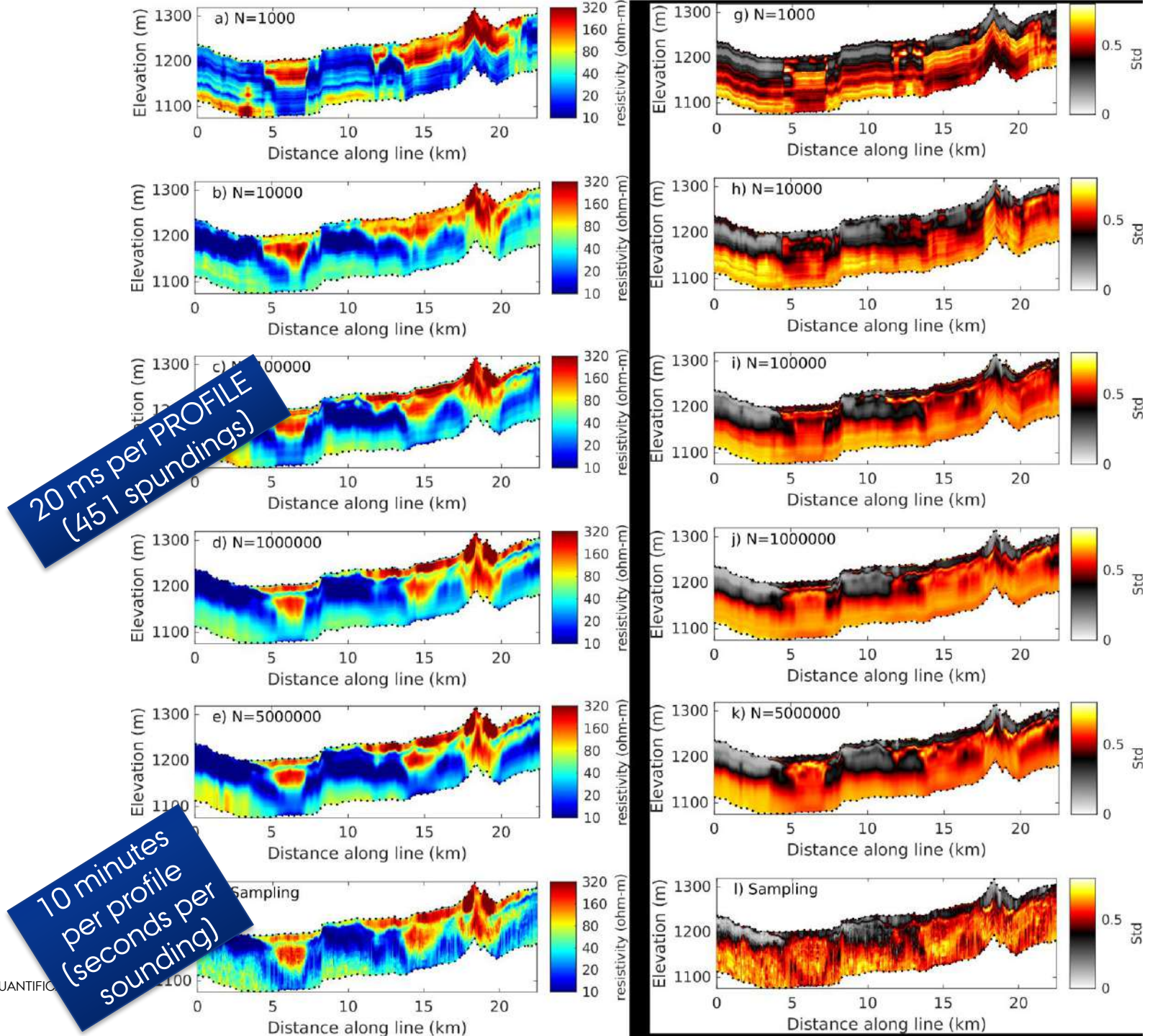
Loss function: Binary cross entropy

Ensures the output of the NN can be interpreted as $P(\mathbf{n} = \mathbf{n}^o)$

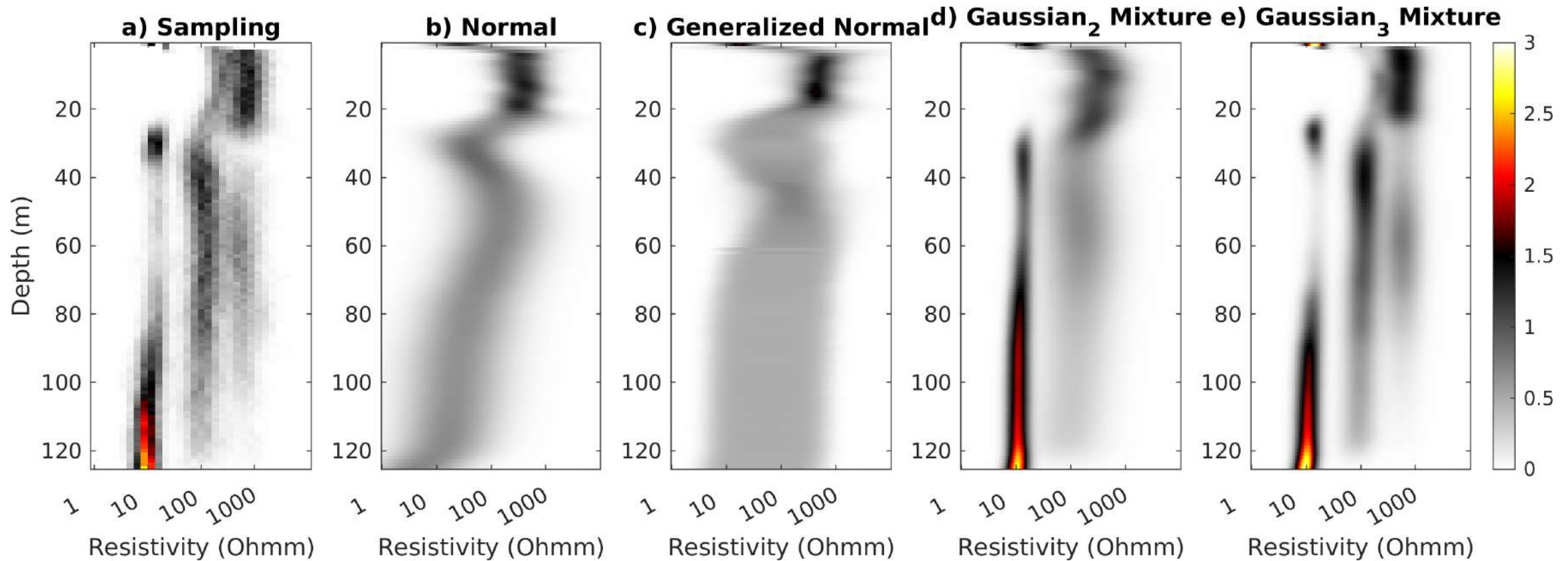
Direct estimation of the pointwise mean and standard deviation for different size of training data set, compared to using the extended rejection sampler.

Realistic sized EM surveys may contain +100.000 soundings, and they are getting larger.!

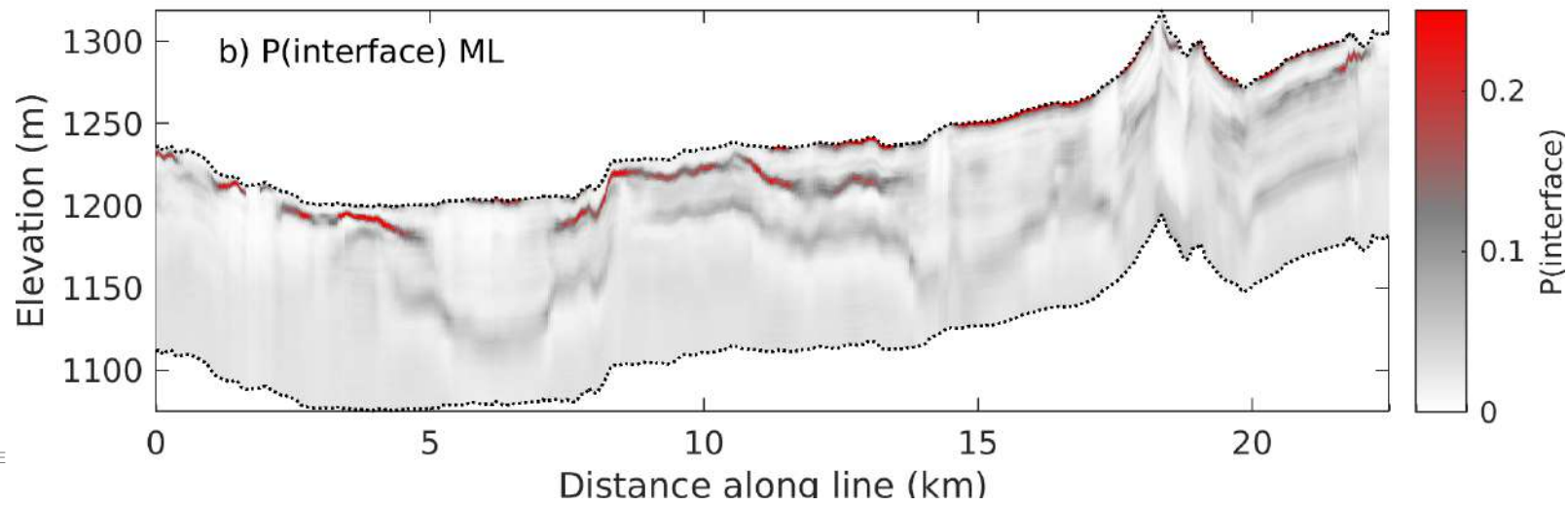
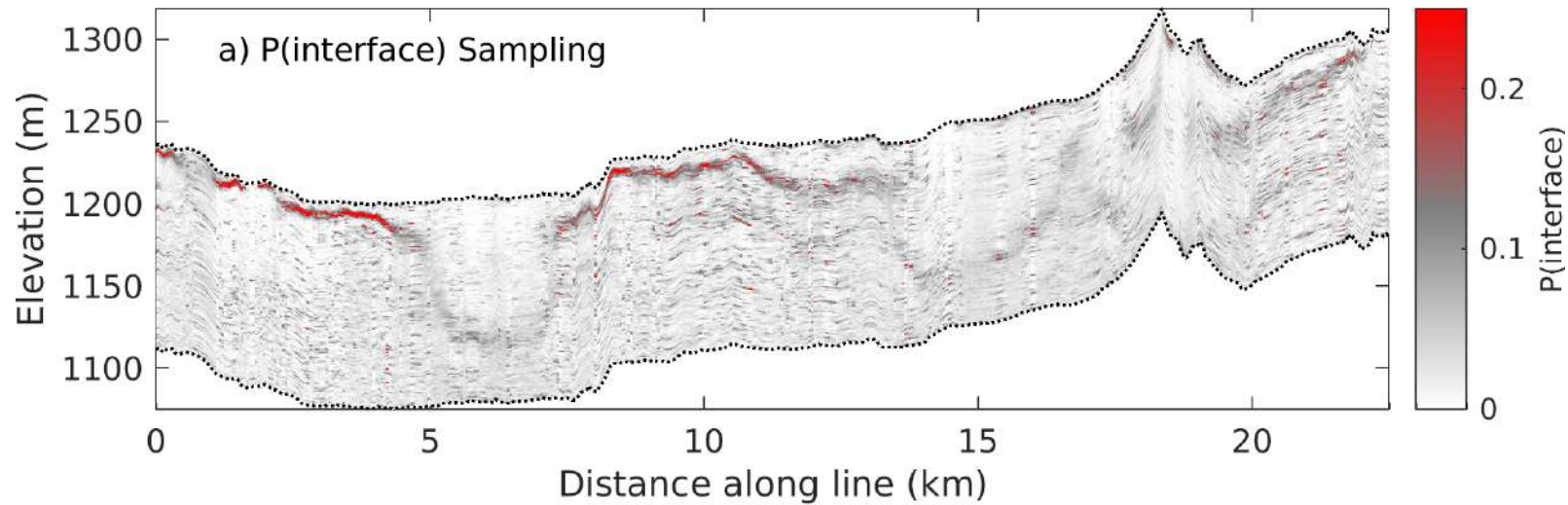
Training time: seconds (N=1000) to 15 minutes (N=5.000.000)



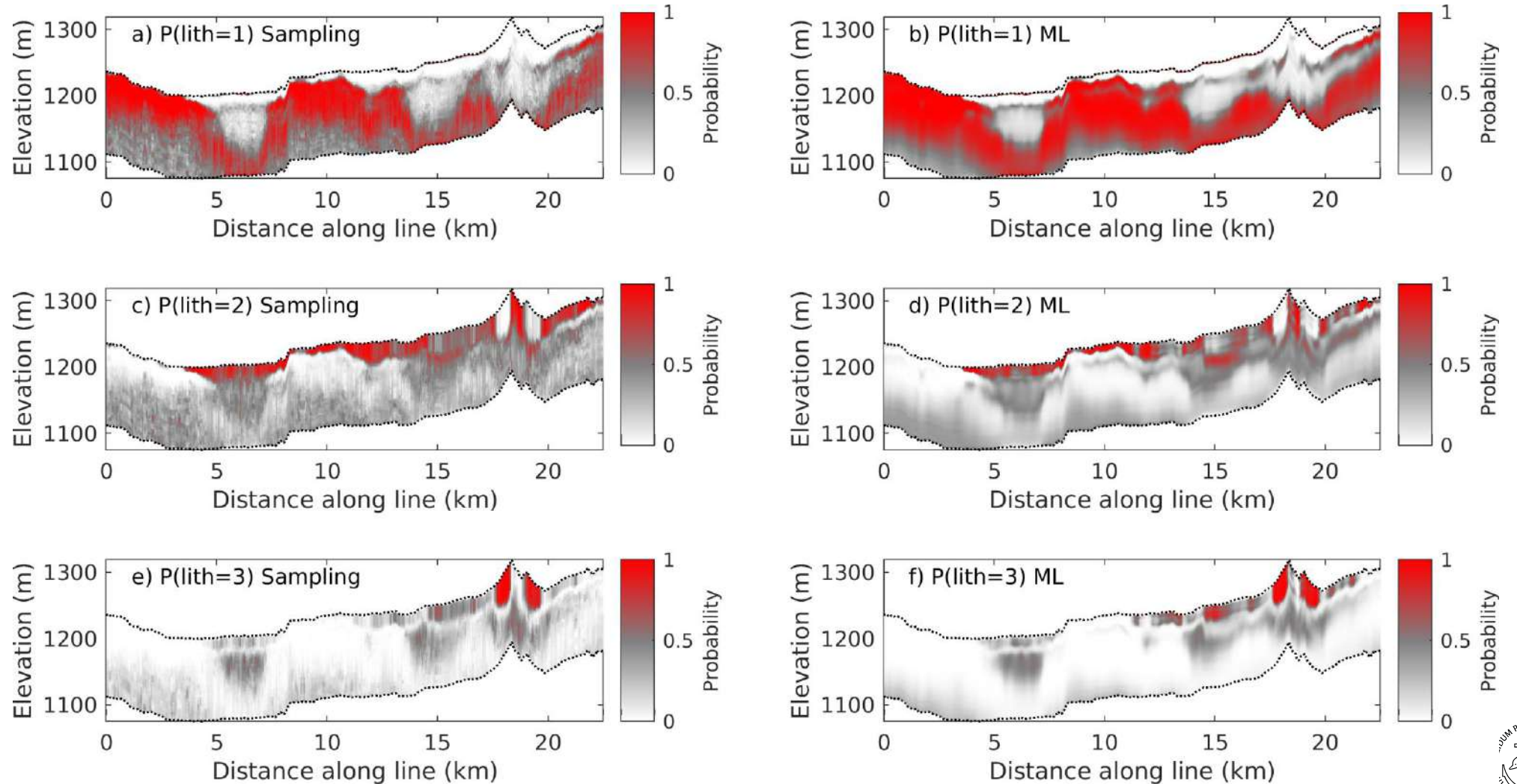
MARGINAL POSTERIOR STATISTICS



Posterior probability of a layer interface



Posterior probability of lithology



ALGORITHMS FOR SAMPLING

Hansen, Thomas M., and Burke J. Minsley. "Inversion of airborne EM data with an explicit choice of prior model." *Geophysical Journal International* 218.2 (2019): 1348-1366.

Minutes per EM sounding

Hansen, Thomas M. "Efficient probabilistic inversion using the rejection exemplified on airborne EM data." *Geophysical Journal International* 224.1 (2021): 543-557.

Seconds per EM sounding

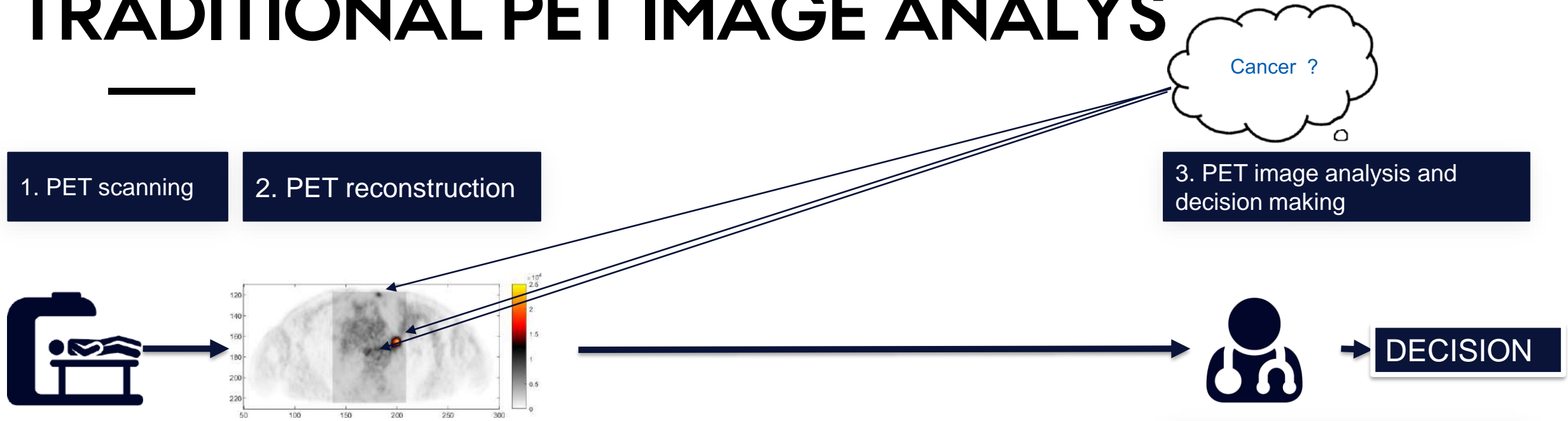
Hansen, Thomas M., and Christopher C. Finlay. "Use of machine learning to estimate statistics of the posterior distribution in probabilistic inverse problems - an application to airborne EM data." in review/arxiv. <https://doi.org/10.31223/X5JS56>

Milli-seconds per EM sounding

CASE: PET IMAGE ANALYSIS

THOMAS MEJER HANSEN, KLAUS MOSEGAARD,
SØREN HOLM, FLEMMING ANDERSEN, BARBARA MALENE FISCHER, AND
ADAM ESPE HANSEN

TRADITIONAL PET IMAGE ANALYSIS



Small cancer cells are difficult to identify due to

- Resolution limitations
- Noise

Identification of large features with high activity is easy.

Early identification of small features with high activity concentration is difficult.

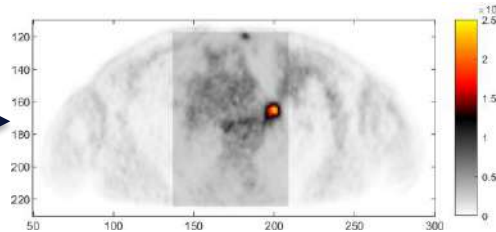
Early identification is key to successful treatment

PROBABILISTIC PET IMAGE ANALYSIS

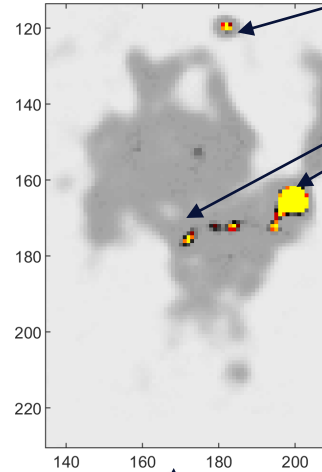
1. PET scanning

2. PET reconstruction

3. PET image analysis and decision making



m : activity concentration,
tissue type, cancer



DECISION

Cancer
99% probability??



$g(m)$



$L(m)$

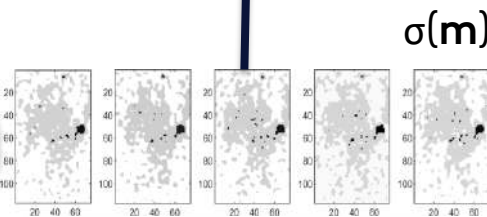
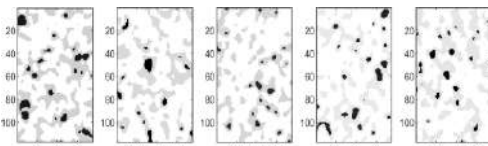


$\rho(m)$

PHYSICS (smoothing)

NOISE model

Medical EXPERTISE



$\sigma(m)$

In **ProPET** all information
available about the PET
image is combined into
one statistical model

One result is an “Enhanced
image that shows the average
intensity (high is indicative of
cancer) in each cell
This aids the medical expert to
take more informed decisions
especially regarding the
existence of small cancer cells

DEFINITIONS

$\Phi = [\Phi_1, \Phi_2, \dots]$: The real activity concentration

Φ_{PET} : The reconstructed PET image

The forward model assumption (convolution)

$$\Phi_{\text{PET}} = G \Phi + n(\Phi)$$

$\rho(\Phi)$: The prior distribution of activity concentration

$\rho_{\Phi}(\Phi_{\text{PET}})$: The noise distribution of Φ_{PET}

$$L(\Phi) = \rho_{\Phi}(G\Phi)$$

PROBABILISTIC PET IMAGE ANALYSIS

A: Quantify information

- A1: $\rho(\Phi)$ Quantify prior information
- A2: $L(\Phi) = \rho_{\Phi}(G\Phi)$ Quantify noise and the forward model

B: Combine information

- Sample from $\sigma(\Phi) \propto \rho(\Phi) L(\Phi)$

C: Clinical quantitative analysis

- Statistical analysis of the obtained sample from $\sigma(\Phi)$

The *extended* Metropolis sampler:

We need to be able to

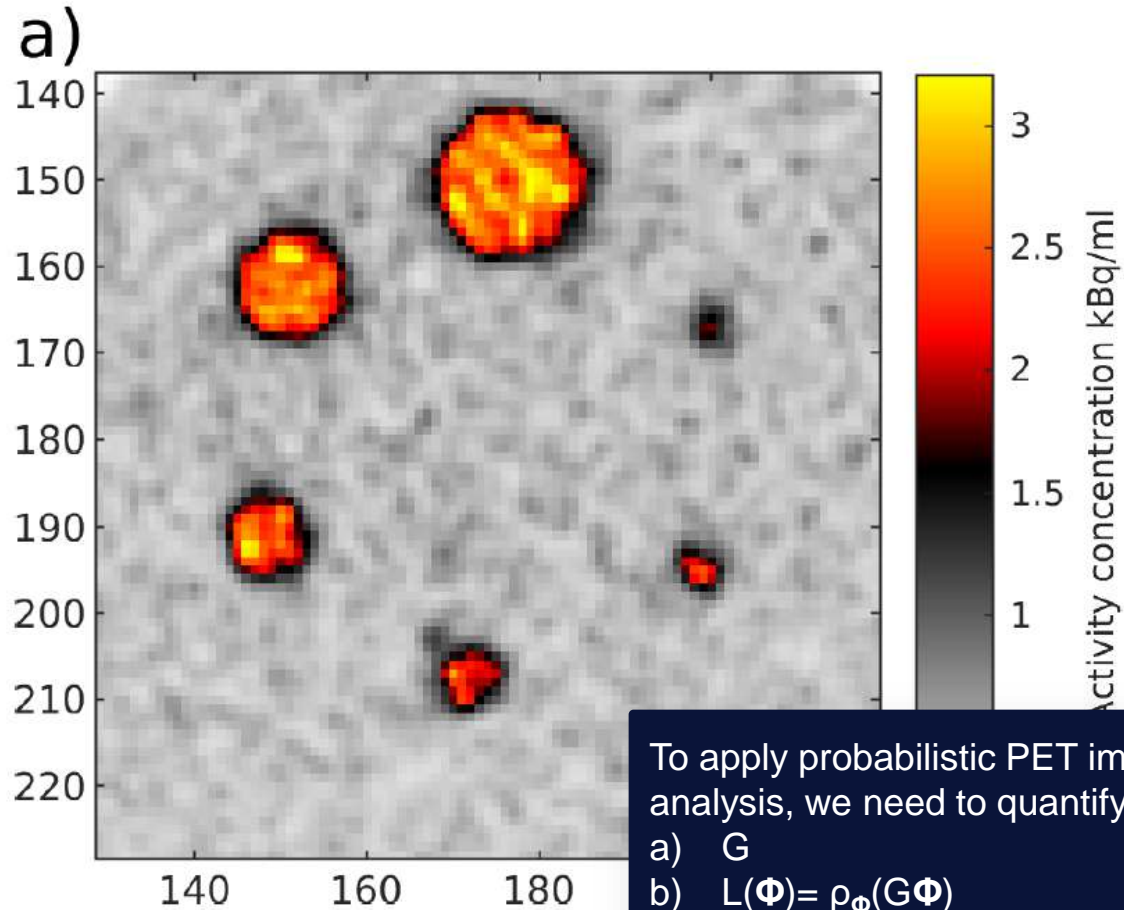
- 1) sample $\rho(\Phi)$, and
- 2) evaluate $L(\Phi)$.

$\rho(\Phi)$ can be represented by an algorithm!

[Mosegaard and Tarantola, 1995]

PHANTOM CASE DATA

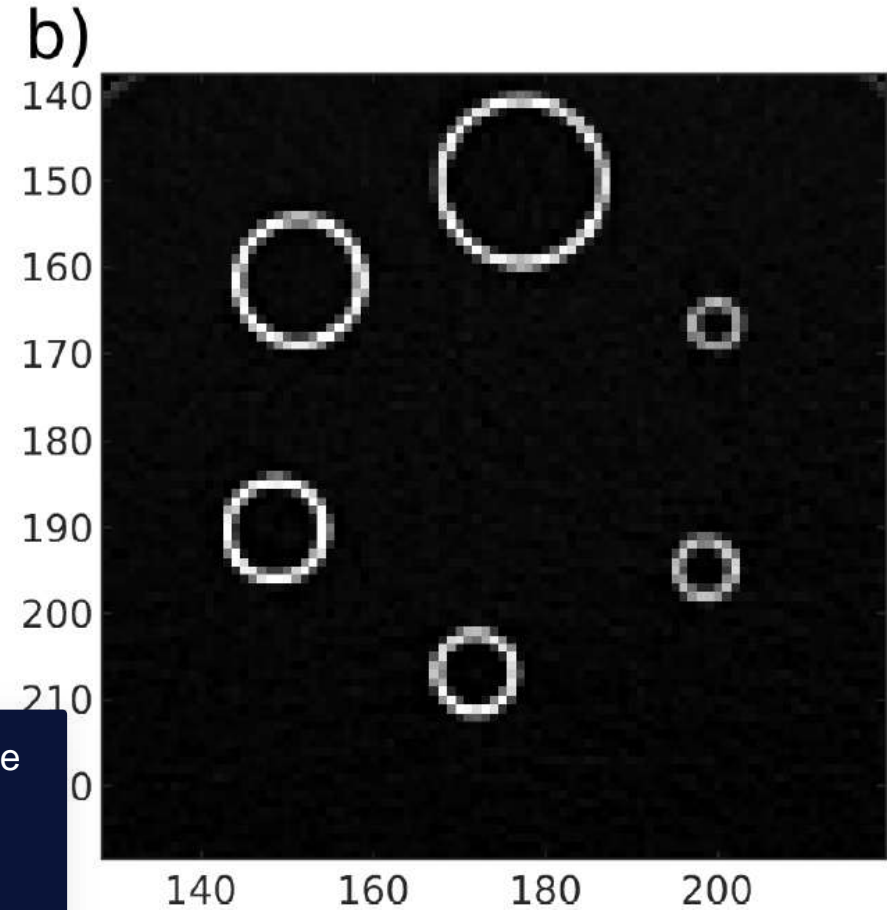
PET



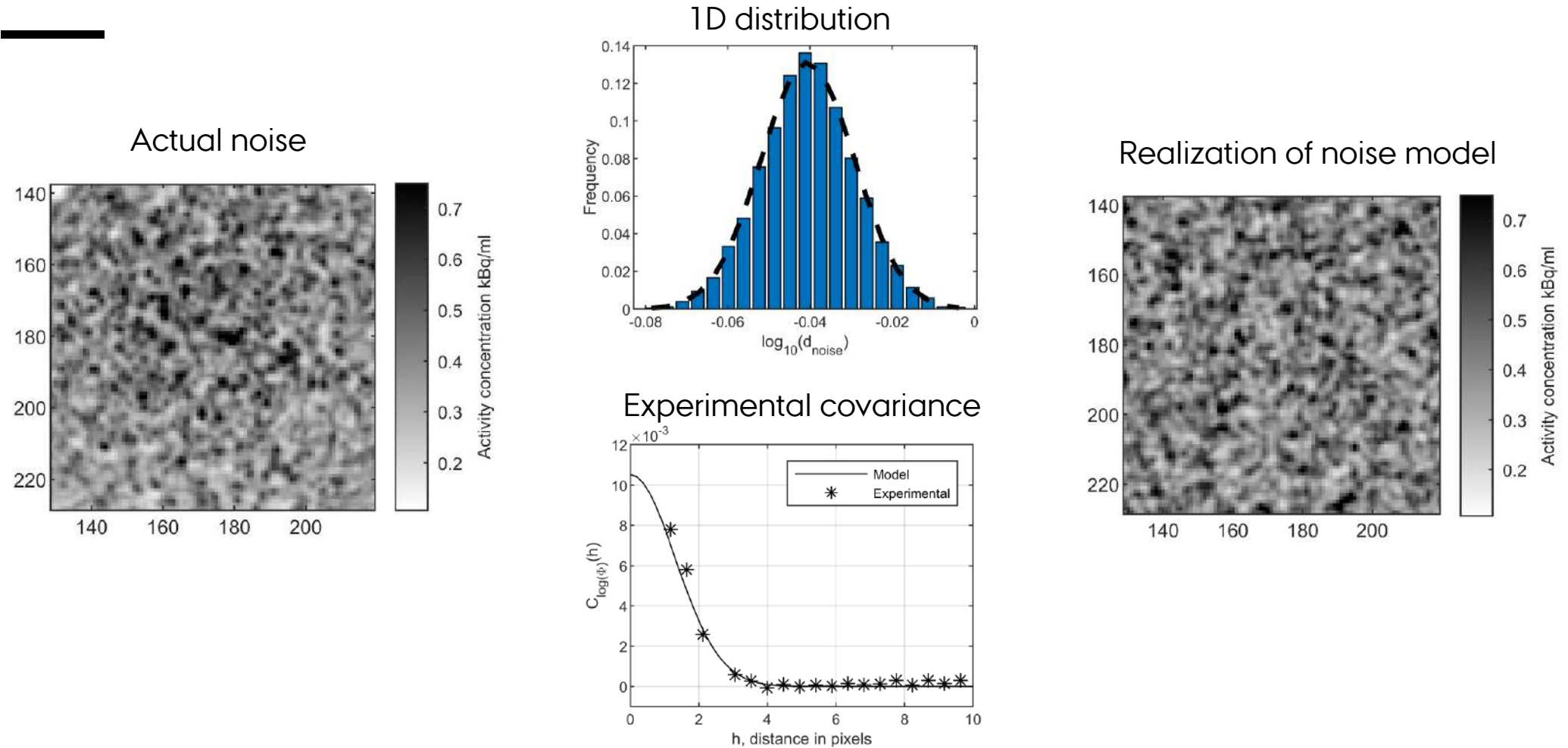
To apply probabilistic PET image analysis, we need to quantify

- a) G
- b) $L(\Phi) = \rho_\Phi(G\Phi)$
- c) $\rho(\Phi)$

CT

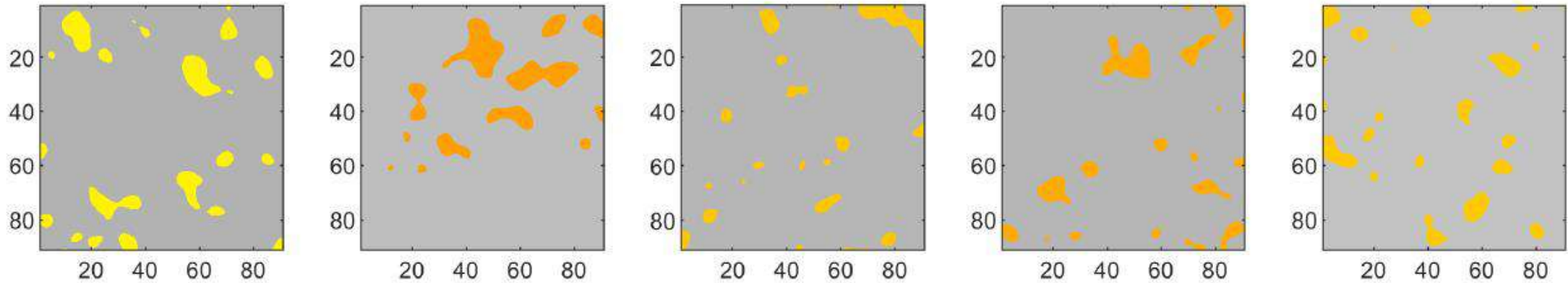


PHANTOM CASE: THE NOISE, MULTIVARIATE CORRELATED LOG-NORMAL

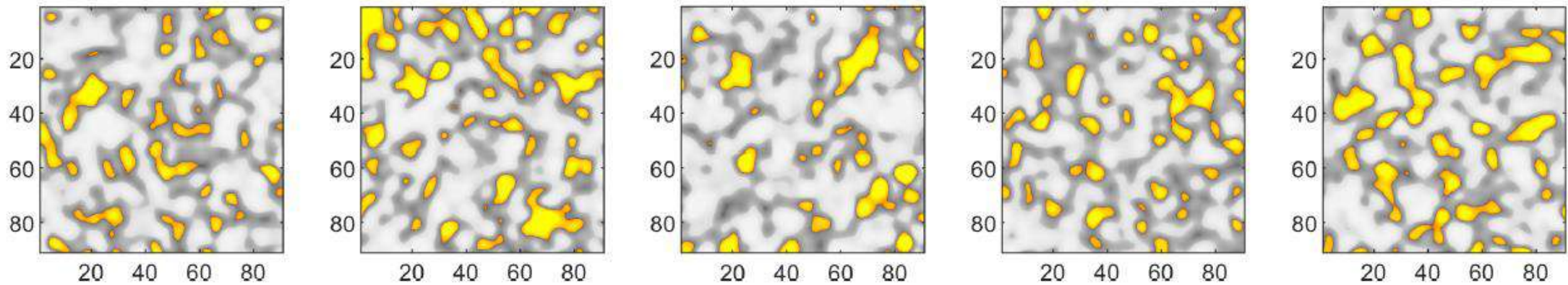


PHANTOM CASE, THE PRIOR

a) $\rho_1(\Phi)$

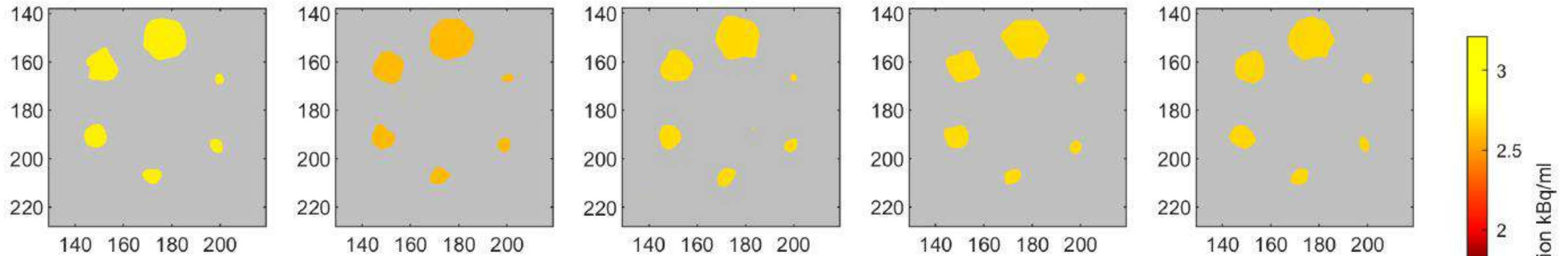


b) $\rho_2(\Phi)$

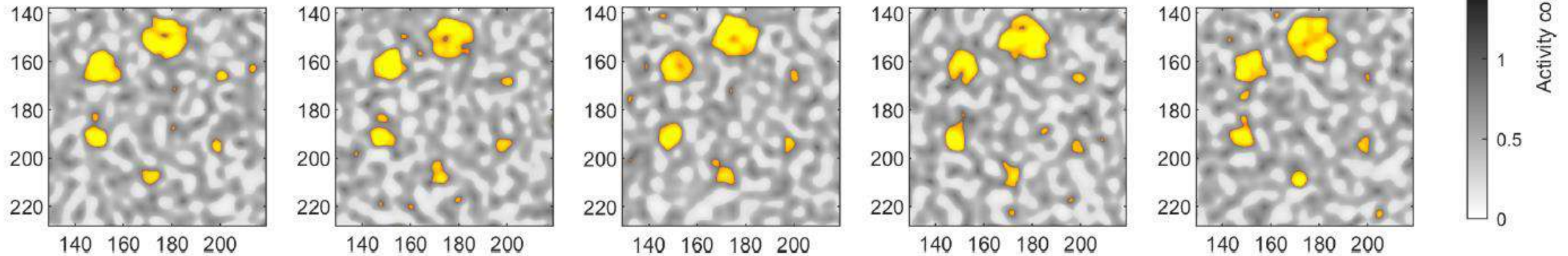


PHANTOM CASE, SAMPLING THE

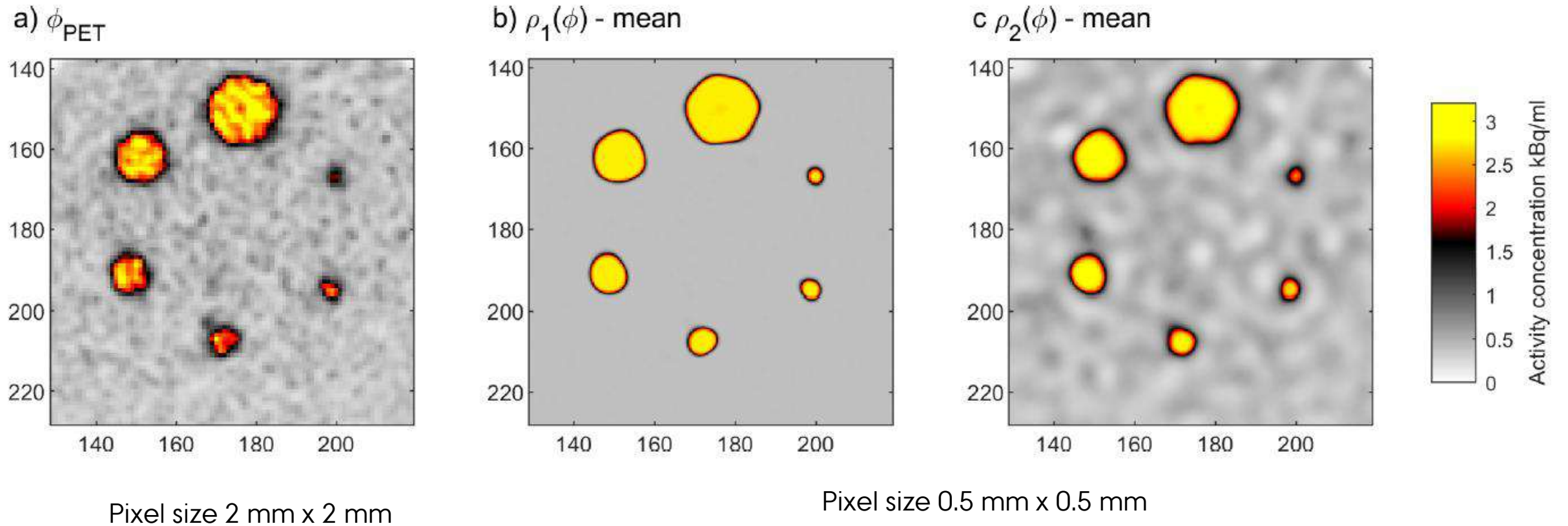
a) $\sigma_1(\Phi)$



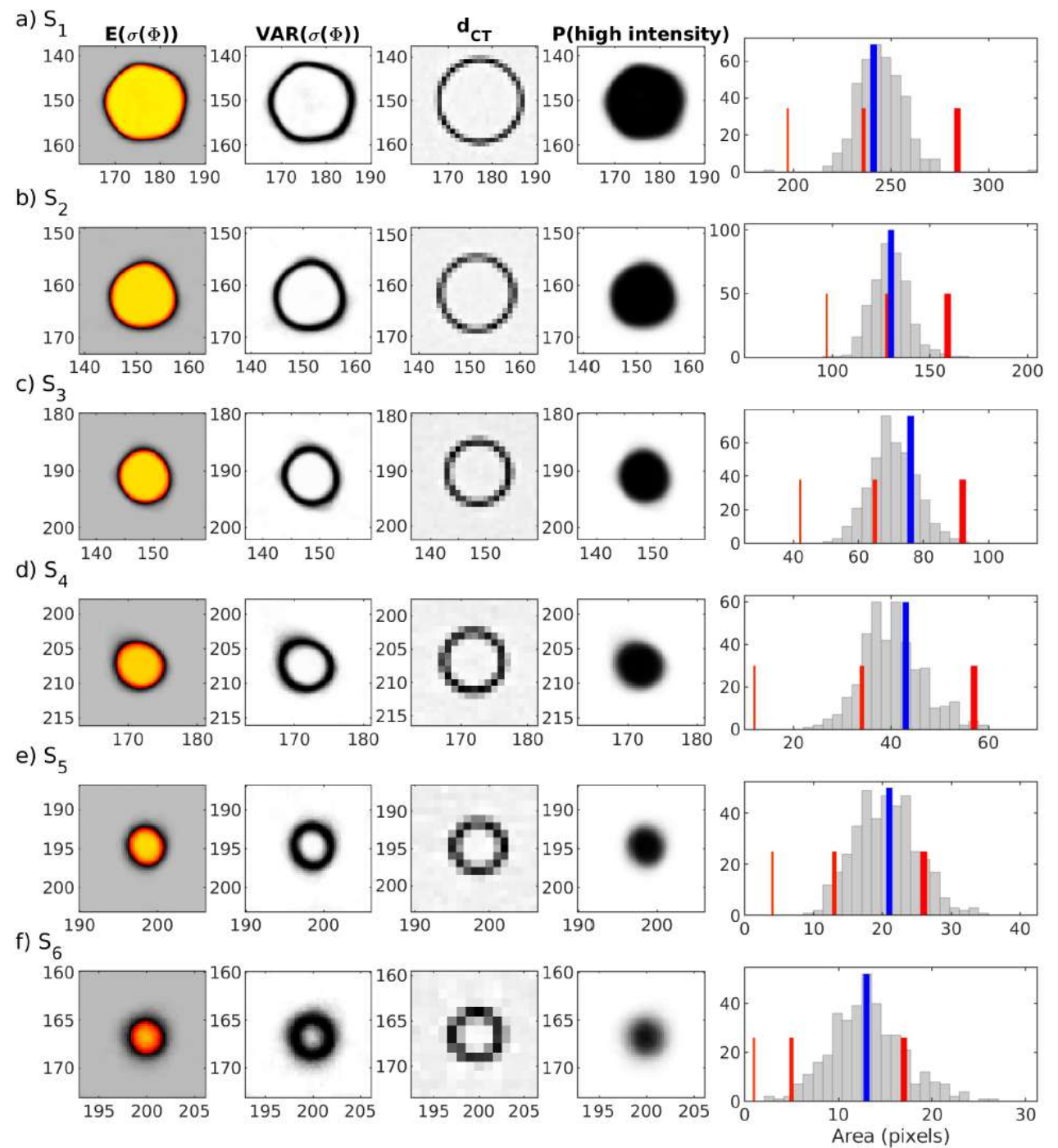
b) $\sigma_2(\Phi)$



PHANTOM CASE – STATISTICS OF THE POSTERIOR



SUB PIXEL RESOLUTION



CONCLUSIONS

Tarantola and Valette (1982) provide a framework for probabilistic integration of information, that naturally allow accounting for uncertainty information.

In order to implement the method in practice we need to be able to quantify available information:

- (Algorithms that can quantify) prior information
 - The choice of choosing/quantifying prior information cannot be avoided
- Algorithms that compute the forward problem (physics)
- Quantification of modelling errors
- Algorithms for integration of information
 - Linear least squares, rejection sampler, Metropolis algorithm, Machine learning.

המחלקה לביטחון
המחלקה לביטחון