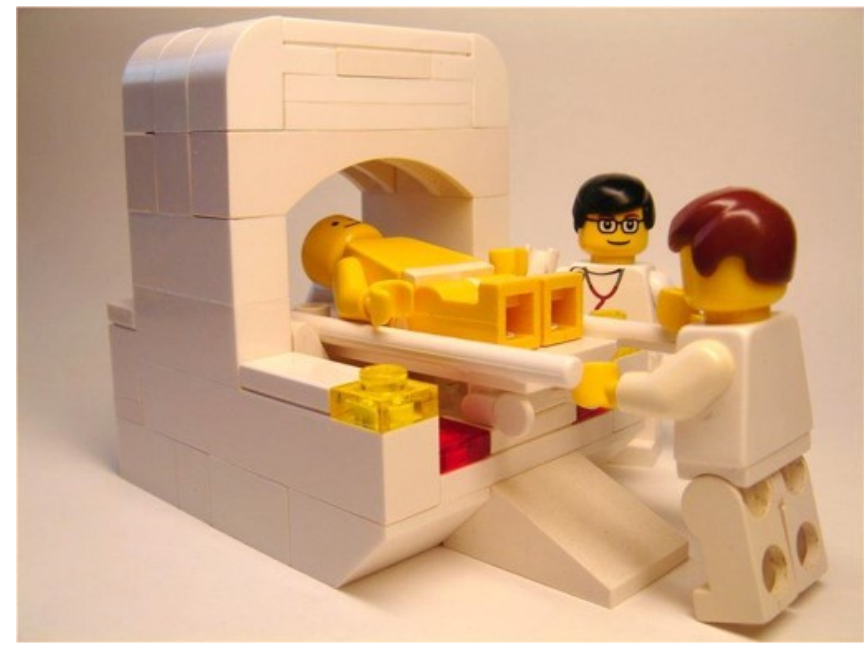


Setting the Stage – True and Nominal View Angles

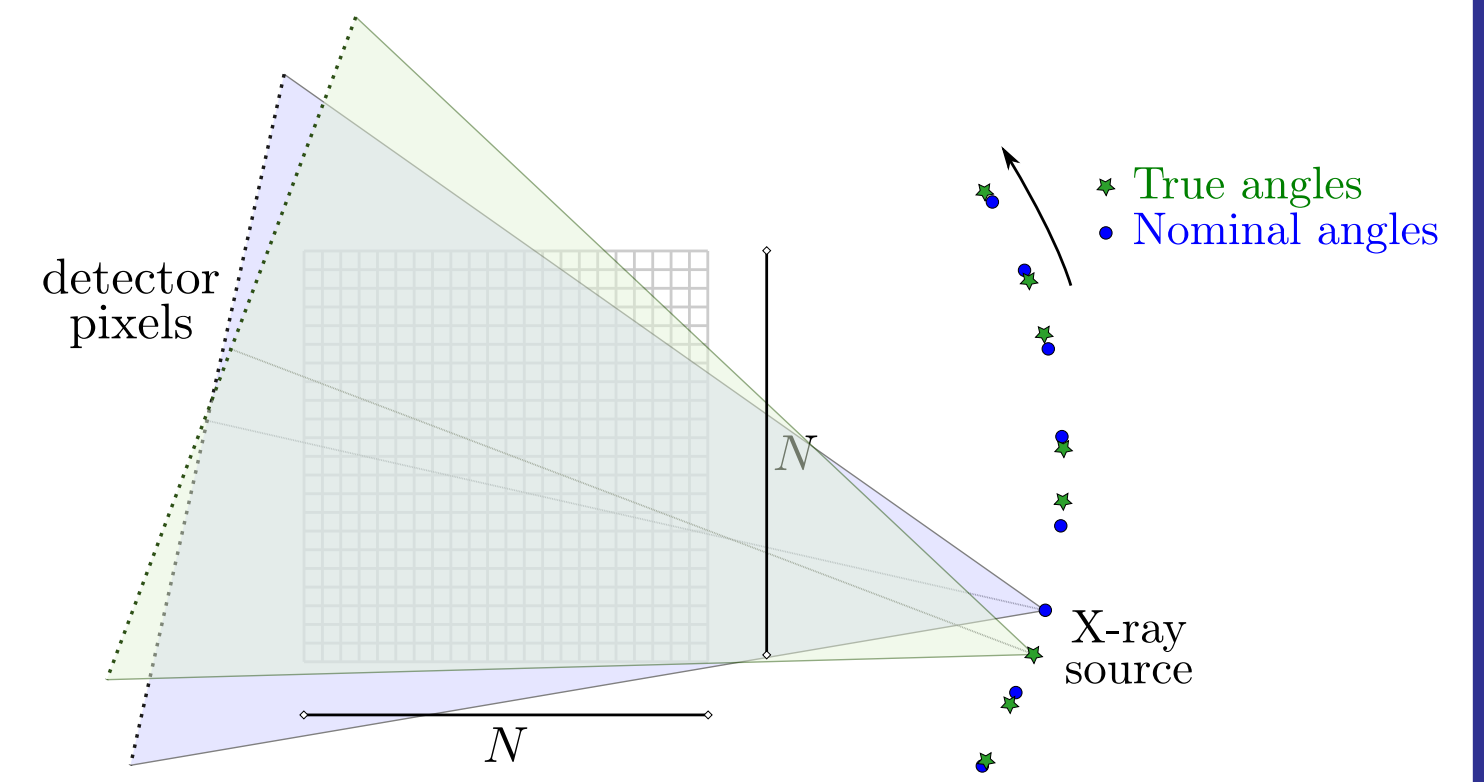


Computed Tomography (CT)

CT data consist of measurements of the attenuation of X-rays passing through an object. We reconstruct an image of the linear attenuation coefficient of the object's interior. For each position of the X-ray source, we measure a set of data referred to as a **view**.

The **true view angles** may differ from the assumed **nominal view angles**:

- The model for the measured data is $\mathbf{b} = \mathbf{A}_{\text{tru}} \mathbf{x} + \mathbf{e}$, where \mathbf{e} is the measurement noise, \mathbf{x} represents the image, and \mathbf{A}_{tru} is the forward model for the *unknown* true angles.
- A “naive” and bad reconstruction uses the matrix \mathbf{A}_{nom} based on the nominal angles.



How to Handle Uncertain View Angles – The Bayesian Framework

We consider the true view angles as unknowns θ , together with the image \mathbf{x} :

$$\text{find } (\mathbf{x}, \theta) \text{ such that } \mathbf{b} = \mathbf{A}(\theta) \mathbf{x} + \mathbf{e}.$$

Here, $\mathbf{A}(\theta)$ denotes the forward model corresponding to the view angles θ .

We apply the Bayesian framework with a likelihood that involves both \mathbf{x} and θ :

$$\pi_{\text{pos}}(\mathbf{x}, \theta) \propto \pi_{\text{lik}}(\mathbf{b} | \mathbf{x}, \theta) \times \pi_{\text{pri}}(\mathbf{x}) \times \pi_{\text{pri}}(\theta).$$

- The distribution of \mathbf{e} is determined by the measurements; in CT it is log-Poisson and we approximate it by a Gaussian. Hence, $\pi_{\text{lik}}(\mathbf{b} | \mathbf{x}, \theta)$ is a **Gaussian**.
- For $\pi_{\text{pri}}(\mathbf{x})$ we use a **Laplace distribution of the differences of neighbour pixels** (enables sharp edges in the image; related to total variation (TV) regularization).
- For $\pi_{\text{pri}}(\theta)$ we use the **von Mises distribution** (i.e., a *periodic* normal distribution).

But wait, there's more. We introduce scalar **hyperparameters**: λ in the Gaussian likelihood, δ in the Laplace-difference prior for \mathbf{x} , and κ in the von Mises prior for θ . All three have exponential distributions $\pi_{\text{hpri}}(\cdot) = \beta \exp(-\beta \cdot)$ with $\beta = 10^{-4}$. Thus, the posterior takes the form

$$\pi_{\text{pos}}(\mathbf{x}, \theta, \lambda, \delta, \kappa) \propto \pi_{\text{lik}}(\mathbf{b} | \mathbf{x}, \theta, \lambda) \times \pi_{\text{pri}}(\mathbf{x} | \delta) \times \pi_{\text{pri}}(\theta | \kappa) \times \pi_{\text{hpri}}(\lambda) \pi_{\text{hpri}}(\delta) \times \pi_{\text{hpri}}(\kappa)$$

How to Sample – A Hybrid Gibbs Sampler

Performing statistical inference of the full posterior $\pi_{\text{pos}}(\mathbf{x}, \theta, \lambda, \delta, \kappa)$ is challenging: the number of pixels n is large, the forward model $\mathbf{A}(\theta)$ is nonlinear in the view angles θ , and the prior $\pi_{\text{pri}}(\mathbf{x} | \delta)$ is nondifferentiable due to the 1-norm. We split the posterior and apply different samplers for each parameter, hence the sampler is hybrid.

$$\pi_1(\mathbf{x} | \theta, \lambda, \delta) \propto \exp\left(-\frac{\lambda}{2} \|\mathbf{A}(\theta) \mathbf{x} - \mathbf{b}\|_2^2 - \delta (\|(I \otimes D) \mathbf{x}\|_1 + \|(D \otimes I) \mathbf{x}\|_1)\right)$$

$$\pi_2(\theta | \mathbf{x}, \lambda, \kappa) \propto \exp\left(-\frac{\lambda}{2} \|\mathbf{A}(\theta) \mathbf{x} - \mathbf{b}\|_2^2 + \kappa \mathbf{1}^T \cos(\theta - \bar{\theta})\right)$$

$$\pi_3(\lambda | \mathbf{x}, \theta) \propto \lambda^{m/2} \exp\left(-\lambda \left[\frac{1}{2} \|\mathbf{A}(\theta) \mathbf{x} - \mathbf{b}\|_2^2 + \beta\right]\right)$$

$$\pi_4(\delta | \mathbf{x}) \propto \delta^n \exp(-\delta [\|(I \otimes D) \mathbf{x}\|_1 + \|(D \otimes I) \mathbf{x}\|_1 + \beta])$$

$$\pi_5(\kappa | \theta) \propto I_0(\kappa)^{-p} \exp(-\kappa [-\mathbf{1}^T \cos(\theta - \bar{\theta}) + \beta])$$

$\mathbf{b} \in \mathbb{R}^m$, $\mathbf{x} \in \mathbb{R}^n$, $\theta \in \mathbb{R}^p$, $I_0 = 0$ -order mod. Bessel funct., $D = \text{bidiag}(-1, 1)$, $\bar{\theta} = \text{nominal angles}$.

Initial states $\mathbf{x}^{(0)}, \theta^{(0)}, \lambda^{(0)}, \delta^{(0)}, \kappa^{(0)}$

For $j = 1, 2, \dots, N_{\text{samp}}$

Sample attenuation coefficients

$$\mathbf{x}^{(j)} \sim \pi_1(\cdot | \theta^{(j-1)}, \lambda^{(j-1)}, \delta^{(j-1)})$$

Sample view angles

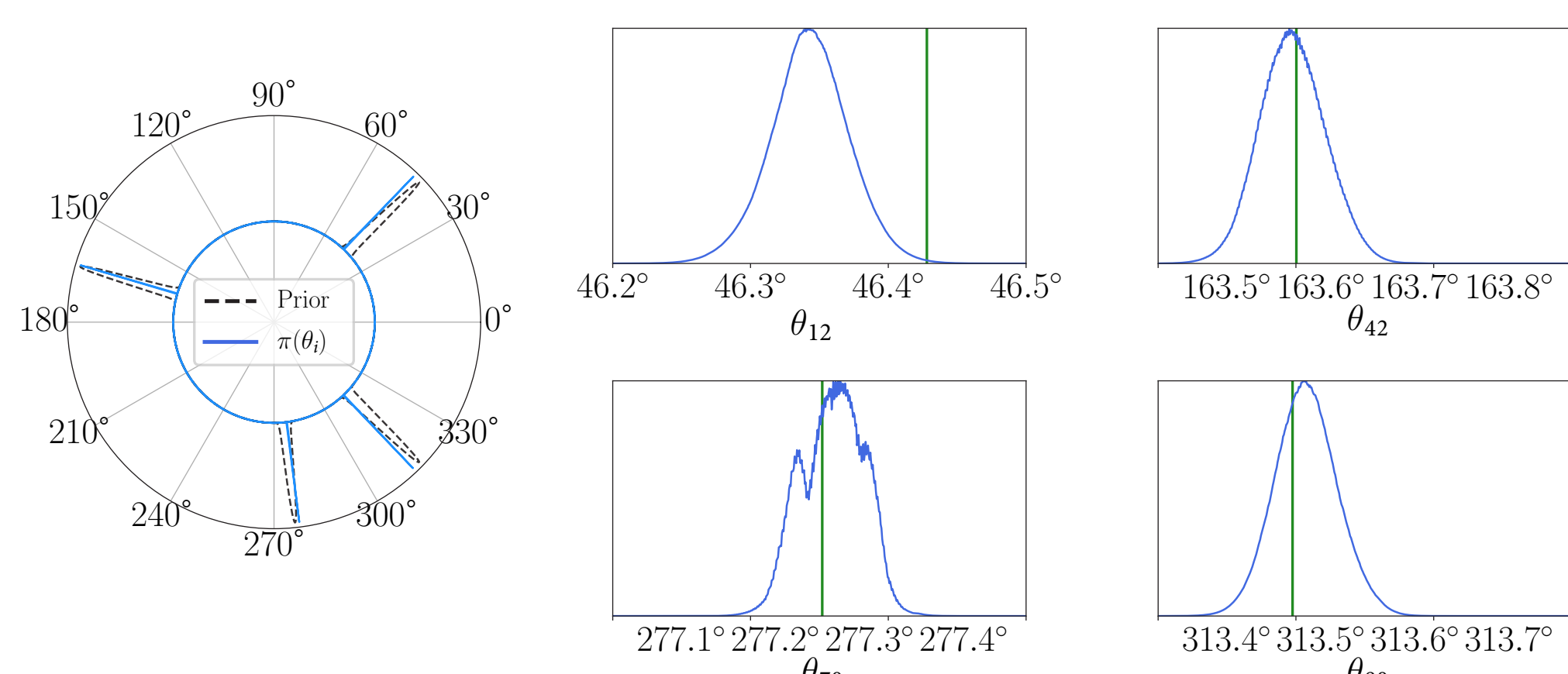
$$\theta^{(j)} \sim \pi_2(\cdot | \mathbf{x}^{(j)}, \lambda^{(j-1)}, \kappa^{(j-1)})$$

Sample hyperparameters

$$\lambda^{(j)} \sim \pi_3(\cdot | \mathbf{x}^{(j)}, \theta^{(j)}), \delta^{(j)} \sim \pi_4(\cdot | \mathbf{x}^{(j)}), \kappa^{(j)} \sim \pi_5(\cdot | \theta^{(j)})$$

End

Simulation Results – View Angles



Left: von Mises prior with the respective densities for selected angles in θ .
Right: some component densities and true angles shown as vertical green lines.

Some Details of the Algorithm

π_1 : non-differentiable due to $\|\cdot\|_1$ and nonlinear in θ . Use *Laplace's approximation*, i.e., a Gaussian $\pi_G = \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \mathbf{H}^{-1})$ with $\mathbf{H}(\mathbf{x}^{(j-1)}) \approx \text{Hessian of } -\log \pi_1$ and

$$\boldsymbol{\mu} = \boldsymbol{\mu}(\mathbf{x}) = \lambda \mathbf{H}^{-1}(\mathbf{x}) \mathbf{A}(\theta)^T \mathbf{b} = \text{MAP estimator of } \pi_1,$$

Much easier to work with a Gaussian but we miss the heavy tails of π_1 . We use 10 CGLS iterations to compute the LS solution that gives the sample $\mathbf{x}^{(j)}$.

π_2 : samples from π_2 are drawn sequentially by componentwise Metropolis:

$$\theta_1^{[k+1]} \sim \pi_2(\theta | \mathbf{x}, \lambda, \kappa, [\theta_2^{[k]}, \theta_3^{[k]}, \dots, \theta_p^{[k]}]),$$

$$\theta_2^{[k+1]} \sim \pi_2(\theta | \mathbf{x}, \lambda, \kappa, [\theta_1^{[k+1]}, \theta_3^{[k]}, \dots, \theta_p^{[k]}]),$$

⋮

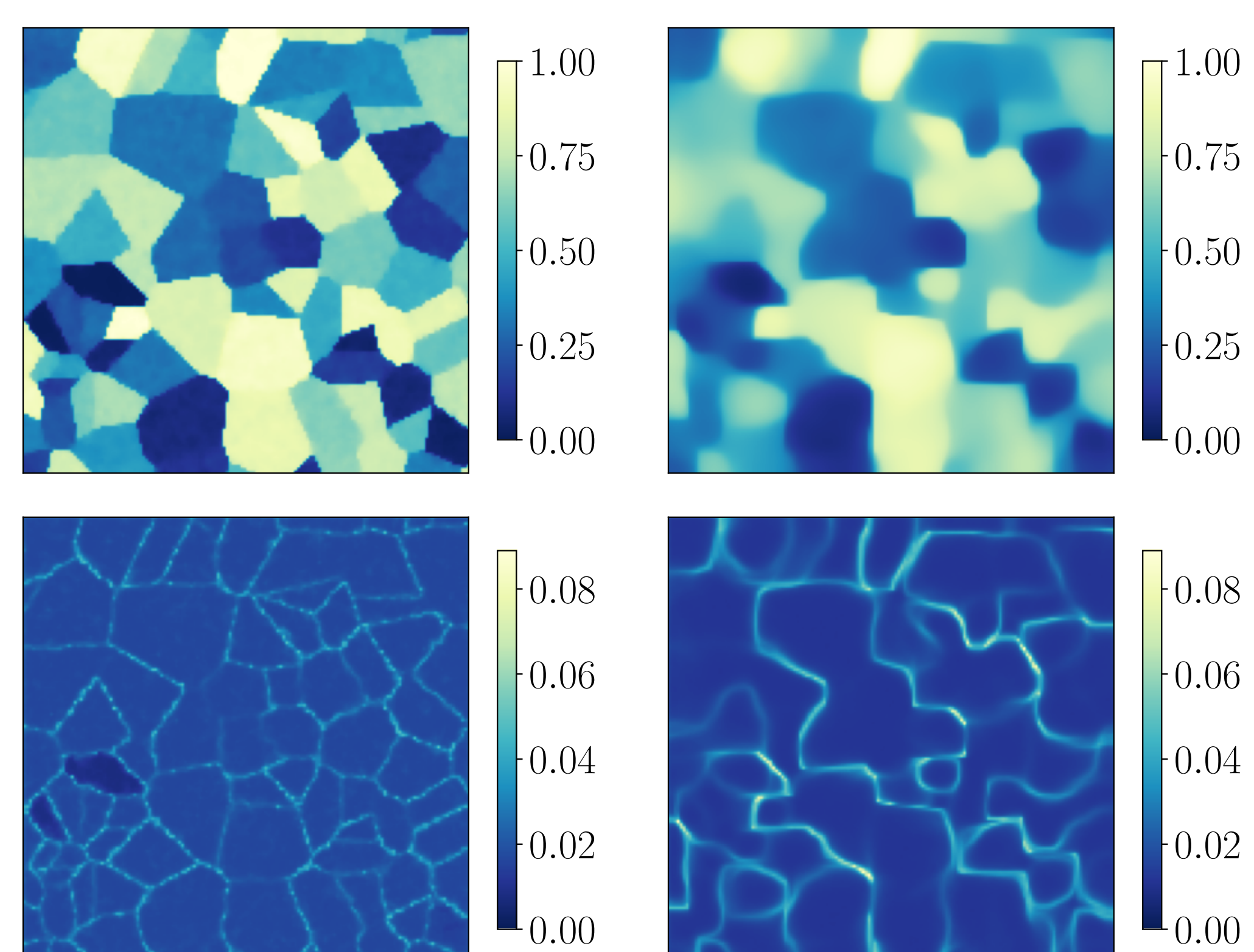
$$\theta_p^{[k+1]} \sim \pi_2(\theta | \mathbf{x}, \lambda, \kappa, [\theta_1^{[k+1]}, \theta_2^{[k+1]}, \dots, \theta_{p-1}^{[k+1]}]).$$

After 20 cycles we obtain $\theta^{(j)} = \theta^{[20]}$.

π_3 and π_4 : can be written and approximated, respectively, in closed form.

π_5 : sampled with standard random-walk Metropolis.

Simulation Results – Metallic Grains Phantom



Left: our method.
Right: using the incorrect nominal angles.
Top: posterior mean
Bottom: st. dev.

Nominal angles give a blurry image with uncertain boundaries.

We compute a sharper image with uncertainty confined to pixels on grain boundaries.

Appendix – Definition of Priors

$$\pi_{\text{pri}}(\mathbf{x} | \delta) = \left(\frac{\delta}{2}\right)^n \exp(-\delta (\|(I \otimes D) \mathbf{x}\|_1 + \|(D \otimes I) \mathbf{x}\|_1))$$

$$\pi_{\text{pri}}(\theta | \kappa) = \left(\frac{1}{2\pi I_0(\kappa)}\right)^p \exp(\kappa \mathbf{1}^T \cos(\theta - \bar{\theta}))$$

Reference and Funding

F. Uribe, J. M. Bardsley, Y. Dong, P. C. Hansen, N. Riis, *A hybrid Gibbs sampler for edge-preserving tomographic reconstruction with uncertain view angles*, SIAM/ASA J. Uncertain. Quantific., 10 (2022), pp. 1293–1320, doi: 10.1137/21M1412268.

A part of the project Computational Uncertainty Quantification for Inverse problems (CUQI), funded by Villum Investigator grant no. 25893 from The Villum Foundation.